

THE APPROXIMATION OF JOINT DISTRIBUTION FUNCTIONS FOR  
APPLICATION IN PROBABILISTIC MECHANICAL DESIGN

by

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## CHAPTER I

### INTRODUCTION

Design engineers have begun to appreciate the limitations of the traditional or classical method of design. For example, the classical criteria for strength-limited design is that the overall strength of the system must be greater than the maximum applied load to the system multiplied by a factor of safety. This safety factor is usually a value between 1.3 and 6, depending on the structure and the materials. Although this approach is very simple in application, it often leads to designs that are overdesigned in terms of weight and bulk. Haugen [1] points out that "Since safety factors are not a performance-related measure there is no way by which an engineer can know whether his designs are near optimum or overconservative."

Designers are now starting to realize that applied loads, dimensions, material properties, etc. are all subject to variation. From this realization has emerged

a new design method called Probabilistic Design. Probabilistic Design treats the design inputs as random variables with each random variable having its own probability density function. After the variables have been defined, the problem can be solved using probability mathematics. The results can then be analyzed to determine if the design satisfies the operational and economic requirements with an acceptable reliability of success.

Although formulating the functional relationships of the random variables into a joint probability integral form is usually possible, evaluating this integral can be a very complex mathematical operation. Springer [2] has written an entire book on the subject of algebraic operations of random variables. Many of the chapters of his book deal with deriving the joint distribution and joint density functions of rather simple functions which contain random variables. Although the mathematical techniques that he describes are certainly important from a theoretical standpoint, there are some disadvantages in practical applications. The designer needs to have a working knowledge of differential and integral calculus, statistical inference, and probability distribution theory. Another drawback is that there are many

functions for which it is impossible to obtain the exact solution to the probability density function. Many times the system is so involved and cumbersome that the average engineer does not have the time or the knowledge to solve the problem.

An alternative to the traditional mathematical approach to Probabilistic Design is a technique known as Monte Carlo Simulation. Law and Kelton [3] define Monte Carlo simulation as "...a scheme employing random numbers which is used for solving certain stochastic or deterministic problems where the passage of time plays no substantive role". The random numbers are generated so that they simulate the physical random process of the probabilistic problem. The system is simulated many times so the frequencies of the events can be studied.

Monte Carlo [4] simulation originated as a tool for the development of the atomic bomb during World War II. The scientists used Monte Carlo simulation to simulate the randomness of neutron diffusion in fissile material. In the last thirty years Monte Carlo techniques have been used for operational research and nuclear physics. Since the motions of neutrons are random, the nuclear physicists have used Monte Carlo to simulate the performance of nuclear reactors. The physicists can

simulate the reactor design without the dangers and expense of an actual physical experiment.

Although the Monte Carlo technique has been around for over 50 years, its application in mechanical engineering design has been limited. The purpose of this thesis is to develop a simulation algorithm that is capable of approximating the joint probability distribution function of complex algebraic expressions encountered in design engineering. This computer algorithm is similar to the Monte Carlo method of simulation. However, this new technique does not use generated random numbers to represent the probability density functions of the individual random variables. Instead, the program divides each density function into a given number of discrete sections with equal probabilities and then takes all possible outcomes into consideration before assembling the joint distribution and density curves.

The literature review that follows is an overall review of probabilistic design and how simulation has been used in different engineering applications.

## CHAPTER II

### LITERATURE REVIEW

#### Probabilistic Machine Design

Bury [5] compared the probabilistic machine design approach to the traditional design method of a simple tension link with a given safety factor, material property, applied load, and cross sectional area. Bury pointed out that many design engineers are dissatisfied [6] with the traditional method because applied loads, material properties, and actual machining dimensions can vary dramatically from one application to another. While the traditional method assumes the input variables to be deterministic, the probabilistic method assumes the variables to be random with unique probability density functions. Once the density functions have been determined, the problem can then be solved as a probabilistic event. However, he points out that this probabilistic approach can only be taken if data is available to properly describe the random variables.

Balkey, Meyer, and Witt [7] discussed the importance of using Probability Structural Mechanics as a tool for evaluating the reliability and structural risk of components and structures. Probability Structural Mechanics, or PMS, combines the traditional methods of structural mechanics with the probabilistic methods to determine the probability of structural failure. They pointed out the necessity of including uncertainties in the mechanical and structural design process. They concluded that "...as components and structures become more sophisticated, failure mechanisms will be probabilistically modeled from the beginning of the design process, and potential design improvements will be evaluated to assess their effects on reducing overall risk."

#### Monte Carlo Simulation

Haugen [1] discussed Monte Carlo simulation throughout his text. One rather simple example was a simulation of strain energy in a statically loaded wire. Haugen solved the problem by Monte Carlo simulation and by an approximate partial derivative method and compared the results. After 200 simulations the strain energy



values were plotted on normal and lognormal probability paper.

Elishakoff [8] used Monte Carlo simulation to predict the buckling time of an elastic bar. The initial imperfections of the bar were assumed to be normally distributed with a given mean and autocorrelation function. The problem was then simulated many times to determine the reliability of the structure. A histogram and distribution of the buckling time was plotted. The results were then compared to the previously published works of Linberg [9]. The Monte Carlo solution was very similar to the results obtained by Linberg.

Elishakoff continued his work of simulating structural imperfections using the Monte Carlo method. His later paper [10] discussed buckling of finite uniform columns resting on a "softening" nonlinear elastic foundation. The initial imperfections were again assumed to be normally distributed. The system was simulated many times and the reliability was plotted vs. nondimensional actual load. The nondimensional buckling loads were then plotted as a histogram and analyzed.

Elishakoff's latest research [11] was concerned with the effect of nonsymmetric random imperfections on the reliability of axially compressed cylindrical shells. He

used the Monte Carlo process to obtain the reliability functions of shells with asymmetric and axisymmetric imperfections. Using the results of his previous works, Elishakoff was able to use measured initial imperfections (variance-covariance matrices and the mean vectors) as a direct input for the Monte Carlo simulation.

Martin [12] developed vectorized and scalar particle transport Monte Carlo algorithms to simulate the transport of photons in a high-density, high-temperature plasma. The behavior of a photon was simulated by drawing random samples from probability distributions that describe the actual physical process. The results were tabulated to chart the overall tallies. Both algorithms were compared to a reference Monte Carlo code from Lawrence Livermore National Laboratory.

Ramsay [13] studied the variability of deflections of ten reinforced concrete fixed-ended T-beams and slabs. The dimensions of the beams, as well as the material properties and type of loadings were assumed to be random variables. A mathematical model was developed to describe the characteristics of the beams and slabs. The model considered the change of stiffness and the redistribution of moment along the length of the structure. Each of the structures was analyzed for four

different loading conditions. For each loading condition, 500 deflection values were generated using Monte Carlo simulation. These values were then converted to deflection ratios and statistically analyzed. The results of the simulation were compared to the deflections of those computed using the ACI 318-63 and ACI 318-71 Code procedures [14,15]. He concluded that the major causes of variability in deflections are the variability of the beam stiffness and the variability of the concrete strength.

Dao-Thien and Massoud [16] discussed the probabilistic distributions of stress and strength relationships. They presented a detailed mathematical background of a method used to obtain the stress and strength distribution functions. The density and distribution functions were written in the form of multiple integrals. A computer program was developed that assists the engineer in selecting a probability distribution (normal, log-normal, or Weibull) using linear regression. The program picks the distribution that minimizes the error function. An example was presented that illustrated a typical structural problem in the aerospace industry. The system was simulated 20,000 times using generated random numbers. The normal,

log-normal, and Weibull distributions were compared to the generated stress and strength data. The program picked the Weibull distribution as the best approximation of the design stress and the normal distribution for the design strength. The two distributions were compared to the histogram frequency distributions generated by the Monte Carlo method.

### Fast Probability Integration

Hasofer-Lind [17], Rackwitz-Fiessler [18], Chen-Lind [19], and Wu [20] have developed a technique called Fast Probability Integration (FPI). The FPI method [21] assumes that all of the random variables are independent with known probability density functions. The random variable's density functions are transformed into equivalent normal random variables using a curve-fitting routine. The design point is then approximated using the scheme of normal tail approximation [22]. Wirsching and Wu [23] have tested various combinations of linear and nonlinear functions containing both normal and non-normal random variables. They report that the FPI method is more accurate and approximately 10 to 100 times faster than the Monte Carlo technique for approximating probabilities

of complicated functions involving several random variables. They point out that while the Monte Carlo method may be useful as a research tool, the large amount of computer processing unit time needed to approximate low values of probability tends to make the Monte Carlo method expensive and impractical.

### Discrete Simulation

Lockwood [24] introduced a new method of computing the probability density functions in turbulent flows. Lockwood pointed out that the traditional method of probabilistic design has several shortcomings. The primary disadvantage is that the "...computational burden increases more or less exponentially with dimensionality." He added that he has developed a more economical approach to the solution of the probability density function of the transport equation. He applied his method to a fully turbulent jet having a given inlet pipe Reynolds number and temperature. He assumed the temperature to be a random variable. To calculate the temperature's probability density function, he divided the range into a finite number of discrete points. The various relations were applied to each discrete point. The total solution

was obtained by summing each discrete point. He compared the predicted probability density functions with the measured probability density functions and determined that the agreement was remarkably good. He concluded that his new method has an advantage over the traditional Monte Carlo simulation technique. Since the discrete method does not use generated random numbers, there is no statistical errors associated with the generation of the random numbers, as is the case with the Monte Carlo technique.

In this study, the probability density functions are divided into discrete points similar to the study conducted by Lockwood. However, this paper goes on to approximate the joint distribution function of algebraic expressions containing several independent random variables with known density functions. The methods and procedures for this simulation are discussed in Chapter 3. The computer programs that are used to approximate the joint probability distribution and density functions are described in Chapter 4. In Chapter 5 the discrete approximation is compared to the traditional Monte Carlo technique for a joint probability distribution that can be solved exactly. Chapter 6 contains an example of how this probabilistic method can

be used as a tool by engineers to aid in the design and development of structures and components. In Chapter 7 a different technique is discussed that determines the upper and lower bounds of a joint probability distribution using the discrete approximation. The conclusions and recommendations for further study are given in Chapters 8 and 9 respectively.

## CHAPTER III

### METHOD AND PROCEDURE

The purpose of this thesis is to develop a simulation algorithm that is capable of approximating the joint probability distribution function of complex algebraic expressions containing several independent random variables with known density functions. If the complex algebraic expression is represented by a variable  $Z$ , then  $Z$  can be written as

$$Z = Z(X_1, X_2, \dots, X_n) \quad (3.1)$$

where  $X_1, X_2, \dots, X_n$  are continuous, independent random variables with individual probability density functions of  $f_1(x_1), f_2(x_2), \dots, f_n(x_n)$  respectively. Since the variables in this study are assumed to be independent, the joint probability density function of  $Z$  can be written as the product of the  $X$  random variables' density functions.



$$f(z) = f_1(x_1)f_2(x_2).....f_n(x_n) \quad (3.2)$$

The joint probability distribution function of the random variable for any n-dimensional set of Z can be written as,

$$F(z) = \int \int \dots \int_Z f_1(x_1)f_2(x_2).....f_n(x_n)dx_n.....dx_2dx_1 \quad (3.3)$$

If the random variables are bounded by the intervals  $(a_1, b_1), (a_2, b_2), \dots, (a_n, b_n)$ , the probability

$$P(a_1 \leq X_1 \leq b_1, a_2 \leq X_2 \leq b_2, \dots, a_n \leq X_n \leq b_n) = \int_{a_1}^{b_1} \int_{a_2}^{b_2} \dots \int_{a_n}^{b_n} f(x_1, x_2, \dots, x_n) dx_n \dots dx_2 dx_1 \quad (3.4)$$

Although formulating the density functions into a multiple integral form similar to equations 3.3 and 3.4 is usually possible, integrating the function over the range of the Z event can be a very complex mathematical operation.

In this chapter, a method that approximates the joint probability distribution function is discussed. The method, called discrete simulation, has four steps:

1. Cell Development
2. Discrete Point Representation
3. Simulation
4. Distribution Approximation

#### CELL DEVELOPMENT

In order to approximate the continuous density function of each random variable, the known density functions are divided into  $n$  sections or cells. Each cell is computed to have equal area, thus equal probability of occurrence.

$$1. \int_{a_0}^{a_1} f(x) dx = \frac{1}{n} \quad (3.5)$$

$$2. \int_{a_1}^{a_2} f(x) dx = \frac{1}{n}$$

$$\cdot \int_{a_2}^{a_3} f(x) dx = \frac{1}{n}$$

$$n. \int_{a_{n-1}}^{a_n} f(x) dx = \frac{1}{n}$$

Figure 3.1 shows an example of a normal random variable density function divided into seven cells with each cell having equal area. For sake of illustration, the density function has been divided into only seven cells. For practical applications the density function can contain anywhere from 30 to 200 cells.

#### DISCRETE POINT REPRESENTATION

After the function has been divided into equal areas, each cell is represented by a single discrete point. This point is located at the cell's center of gravity where:

$$\int_{a_o}^{a_{cg}} f(x) dx = \int_{a_{cg}}^{a_1} f(x) dx = \frac{1}{2n} \quad (3.6)$$

and has a magnitude of  $1/n$ . Figure 3.2 shows a normal density function divided into seven cells with its center of gravity locations. Table 3.1 lists the center of gravity locations of a normal distribution approximation for seven cells.

FIGURE 3.1  
NORMAL DISTRIBUTION  
DIVIDED INTO 7 CELLS

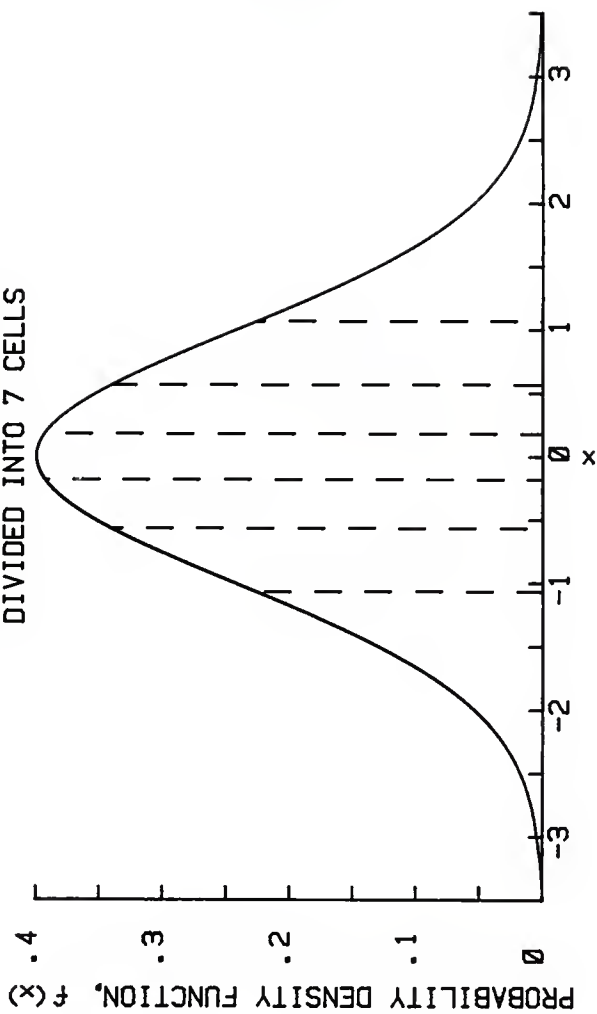


FIGURE 3.2  
7 CELL DISCRETE APPROXIMATION  
OF A NORMAL DISTRIBUTION

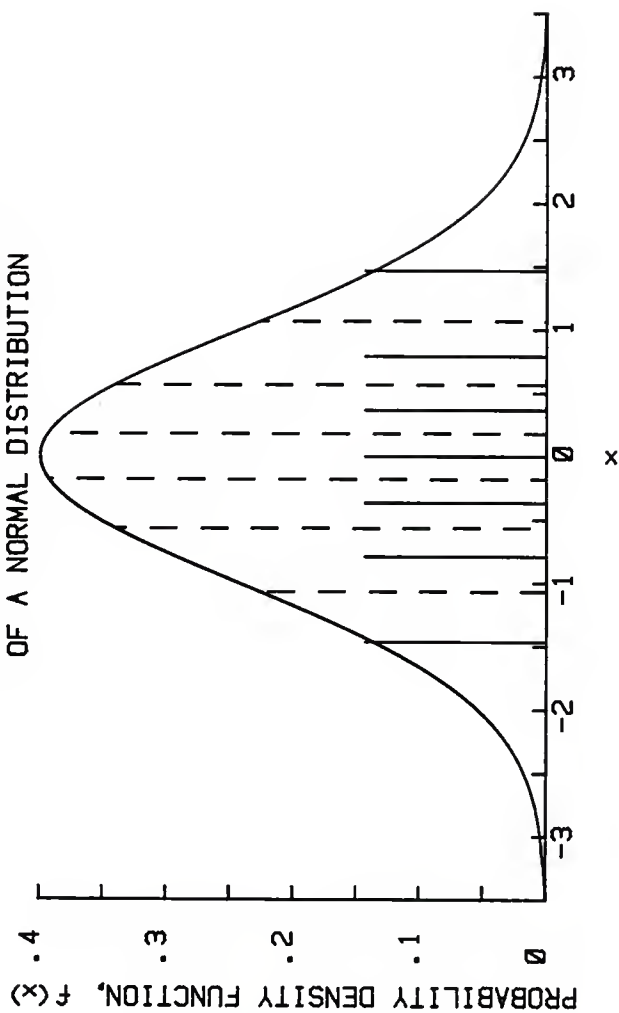


Table 3.1

Normal Distribution Approximation  
Using Seven Discrete Points

Distribution: Normal  
Mean: 0.0  
Variance: 1.0  
Standard Deviation: 1.0  
Number of Cells: 7

Cell Number	Center of Gravity Location
1	-1.4655158
2	-0.7914320
3	-0.3656621
4	0.0000000
5	0.3656621
6	0.7914320
7	1.4655158

## SIMULATION

Once each variable's density function has been chosen, all possible combinations of the expression are simulated. For example, if the joint probability distribution function of the equation

$$Z = A * B$$

is desired, and if the density functions of A and B are simulated using 100 and 150 discrete points respectively. The total possible combinations of Z is  $100 * 150 = 15,000$ .

# DISTRIBUTION FUNCTION

After the outcomes  $Z$ , have been computed, the data is sorted from smallest to largest. Each outcome has a probability of  $1/m$ , where  $m$  is the number of total outcomes. The joint cumulative distribution function can then be assembled by summing the ordered value's probabilities.

$$F(z) = \begin{cases} 0 & i < z_1 \\ F(z_{i-1}) + \frac{1}{m} & z_1 \leq z < z_{i+1} \\ 1 & z_m \leq z \end{cases} \quad \text{for } i=1, 2, \dots, m-1 \quad (3.7)$$

If the values between the discrete points need to be known, a linear empirical distribution function can be written that interpolates between the known points:

$$G(z) = \begin{cases} 0 & \text{if } z < z_0 \\ G(z_{i-1}) + \frac{z - z_{i-1}}{z_i - z_{i-1}} [G(z_i) - G(z_{i-1})] & \text{if } z_{i-1} \leq z < z_i \\ 1 & \text{if } z_m \leq z \end{cases} \quad (3.8)$$

The computer program that performs the distribution approximation is listed in the Appendix.

## DENSITY FUNCTIONS

Although developing a joint density function is not the primary purpose of this paper, two methods are used with reasonable success. These methods are only used to give an approximate shape of the joint probability density function and not an exact solution. The fitting of density functions is a complex problem in itself. The two methods are: histograms and cubic splines.

### Histograms:

A histogram is simply a graphical estimate of the density function corresponding to the known data. To construct a histogram, the range of data is divided into  $k$  equal intervals. The data is then placed in its appropriate interval. Once all of the data has been sorted, the proportion of the values in each interval is determined.

The histogram is constructed using Sturge's [25] rule. The number of cells  $k$  is determined by the equation

$$k = 1 + 3.3 * \log_{10} m \quad (3.9)$$



The range  $r$ , is estimated by the relationship

$$r = x_{\max} - x_{\min} \quad (3.10)$$

And the width  $w$ , of each interval in the histogram is

$$w = r/k \quad (3.11)$$

Although the histogram is easy to construct, the main disadvantage is the loss of information from grouping the data. Because of this, the histogram is used only as a reference to the possible density function's shape.

#### Cubic Splines:

A cubic spline is a set of cubic polynomials which is joined together to form a continuous function. A cubic spline subroutine is used to compute the first derivative of the function. The subroutine also determines a "smoother" approximation of the distribution function. The spline method uses a cubic interpolating function to approximate a "smooth" line through a given set of data points by minimizing the integral of the second derivative squared of the cubic function. A tolerance is imputed that allows the programmer to

tighten or loosen the curve fit. A detailed description of cubic splines can be found in reference [26].

The computer program listed in Appendix B performs the density approximations outlined above.

#### EXAMPLE

Given: Function  $Z=(X*Y**3)/12$ , where X and Y are continuous independent random variables. X is a uniform random variable with given probability density function:

$$f(x) = \begin{cases} 1/3 & 1 \leq x \leq 4 \\ 0 & \text{otherwise} \end{cases}$$

Y is a triangular random variable with a given probability density function:

$$f(y) = \begin{cases} \frac{2(y-4)}{7} & 1 \leq y \leq 2.5 \\ \frac{2(4-y)}{5.25} & 2.5 \leq y \leq 4 \\ 0 & \text{otherwise} \end{cases}$$

Find: The joint probability distribution of random variable  $Z=Z(X,Y)$ .

For simplicity each cell is approximated using only five cells. The area of each cell is equal to  $1/5$ . The

bound for the first cell of variable X is found by using the relationship:

$$\int_{a_0}^{a_1} f(x) dx = \text{Area of the cell}$$

$$\int_1^{a_1} 1/3 dx = 1/5$$

$$x/3 \Big|_1^{a_1} = 1/5$$

$$a_1 = 8/5$$

The center of gravity for cell 1 is computed using equation 3.6

$$\int_1^{a_{cg}} 1/3 dx = 1/10$$

$$a_{cg} = 13/10$$

The remaining locations can be found using the same method. Figures 3.3 and 3.4 show the two density functions and their respective centers of gravity. Table 3.2 lists the center of gravity locations for both random variables.

FIGURE 3.3  
5 CELL DISCRETE APPROXIMATION  
OF A UNIFORM DISTRIBUTION

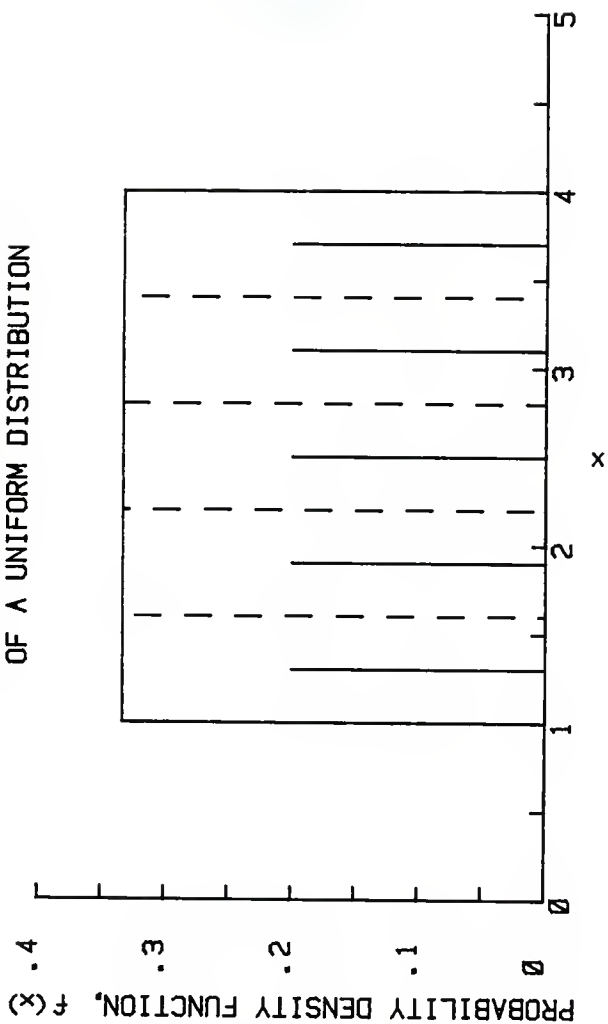


FIGURE 3.4  
5 CELL DISCRETE APPROXIMATION  
OF A TRIANGULAR DISTRIBUTION

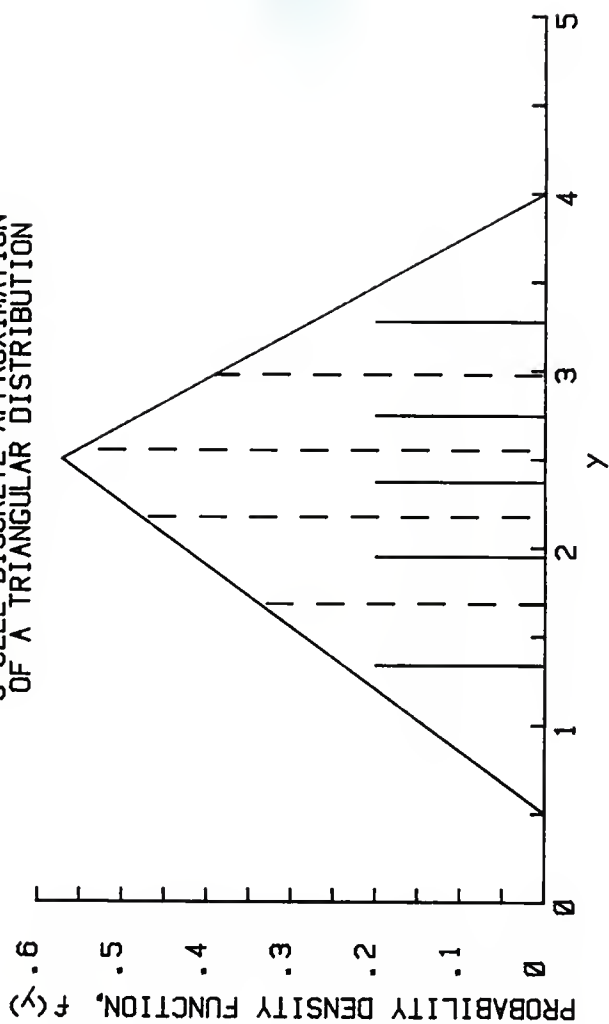


Table 3.2

Discrete Approximation  
of Random Variables X and Y  
Using Five Cells

Random Variable: X	Random Variable: Y
Distribution: Uniform	Distribution: Triangular
Mean: 2.5	Mean: 2.3333
Variance: 0.75	Variance: 0.5139
Number of Cells: 5	Number of Cells: 5

Center of Gravity	Cell	Center of Gravity
<u>Location</u>	<u>Number</u>	<u>Location</u>
1.300000	1	1.336660
1.900000	2	1.949138
2.500000	3	2.370829
3.100000	4	2.745010
3.700000	5	3.275431

After computing all possible values, the 25 unique Z locations are sorted. Each Z location has a probability of 1/25. Figure 3.5 shows the 25 discrete Z locations. Using equations 3.7 and 3.8 a joint probability distribution function is assembled. Table 3.3 shows the sorted Z locations and their distribution values. Figure 3.6 shows the assembled Z distribution function.

Figures 3.7 and 3.8 show the Z distributions of the spline smoothing subroutine for a tolerance of 1.0 and 1.5 respectively.

FIGURE 3.5  
Z DISTRIBUTION OF 25 DISCRETE  
LOCATIONS  $Z=(X*Y^3)/12$

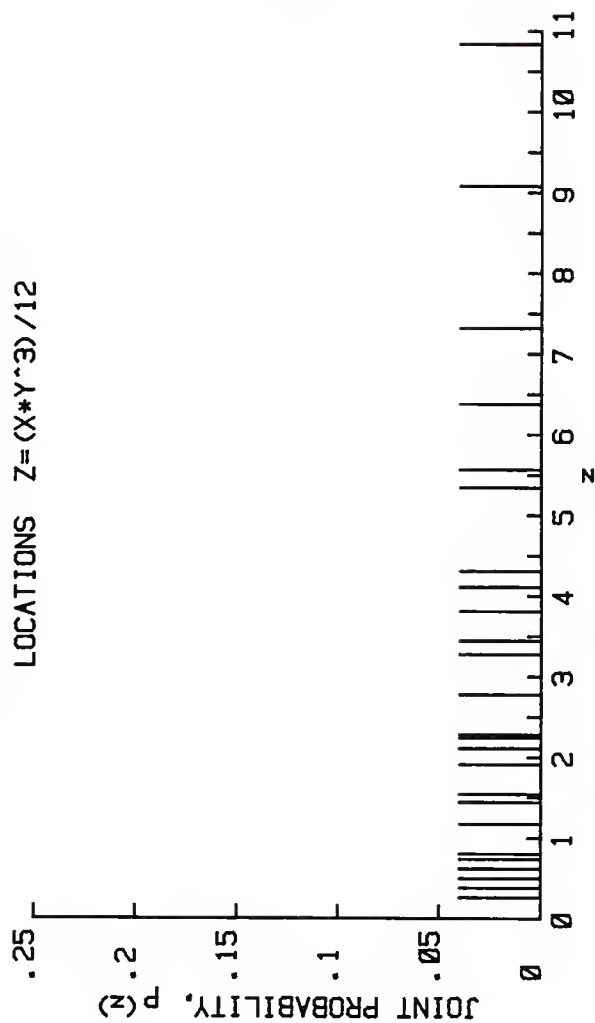


Table 3.3  
Distribution of 25 Discrete  
Z Locations  
 $Z=(X*Y**2)/12$

<u>Number</u>	<u>Z Location</u>	<u>Z Distribution</u>
1	0.2587170	0.0400000
2	0.3781249	0.0800000
3	0.4975327	0.1200000
4	0.6169406	0.1600000
5	0.7363484	0.2000000
6	0.8022129	0.2400000
7	1.1724650	0.2800000
8	1.4436524	0.3200000
9	1.5427172	0.3600000
10	1.9129693	0.4000000
11	2.1099535	0.4400000
12	2.2407525	0.4800000
13	2.2832214	0.5200000
14	2.7762546	0.5600000
15	3.2749459	0.6000000
16	3.4425557	0.6400000
17	3.8068655	0.6800000
18	4.1088568	0.7200000
19	4.3091393	0.7600000
20	5.3433328	0.8000000
21	5.5638804	0.8400000
22	6.3775262	0.8800000
23	7.3208952	0.9200000
24	9.0779101	0.9600000
25	10.8349250	1.0000000



FIGURE 3.6  
Z DISTRIBUTION  
 $Z = (X*Y^3)/12$

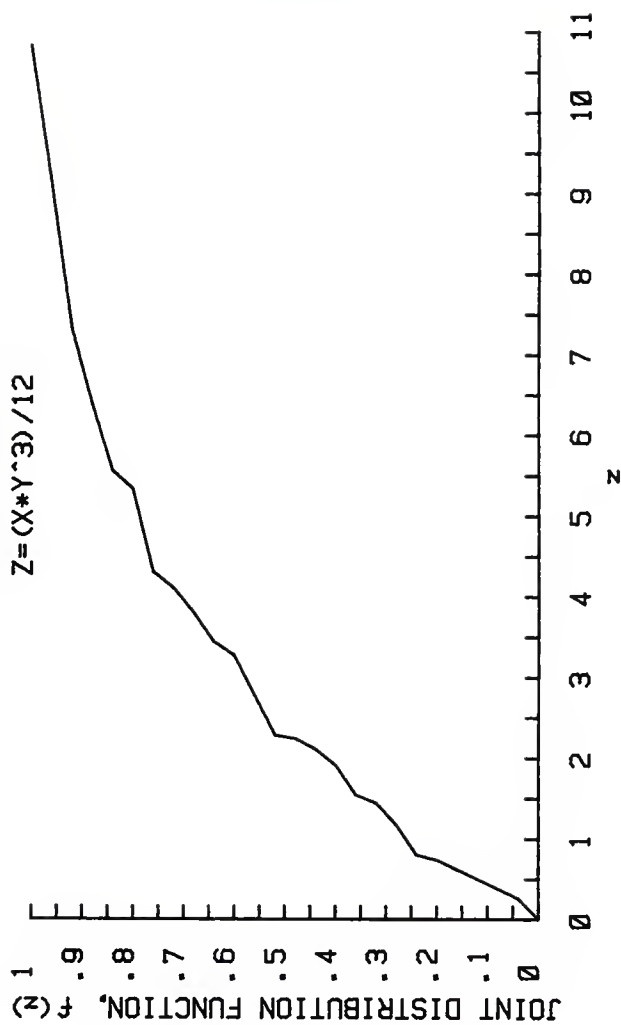


FIGURE 3.7  
 'SMOOTHED' CUBIC SPLINE  
 Z DISTRIBUTION,  $Z=(X*Y^3)/12$   
 TOLERANCE=1.0

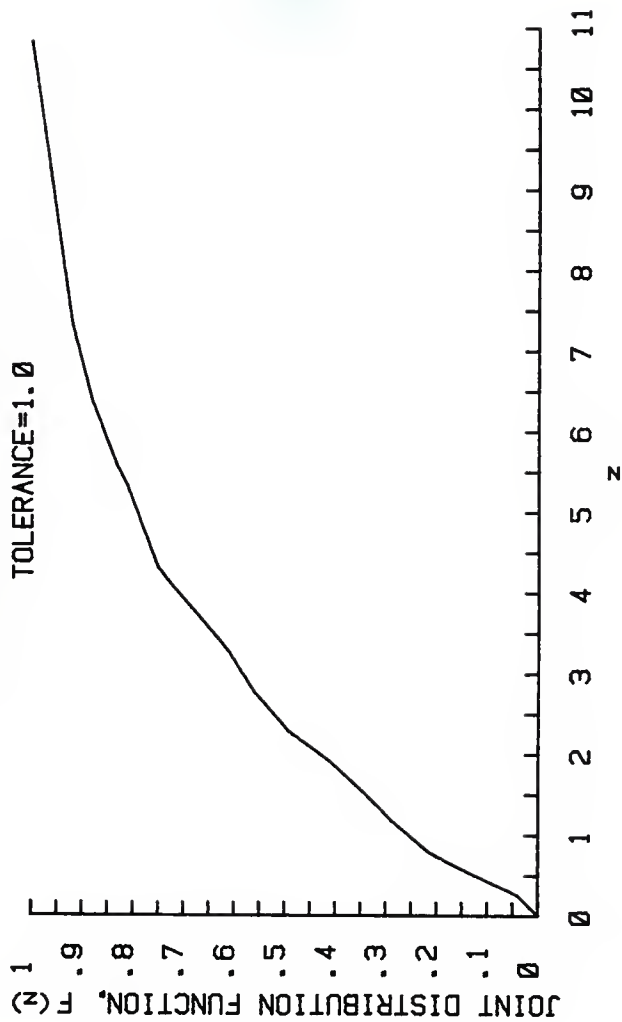
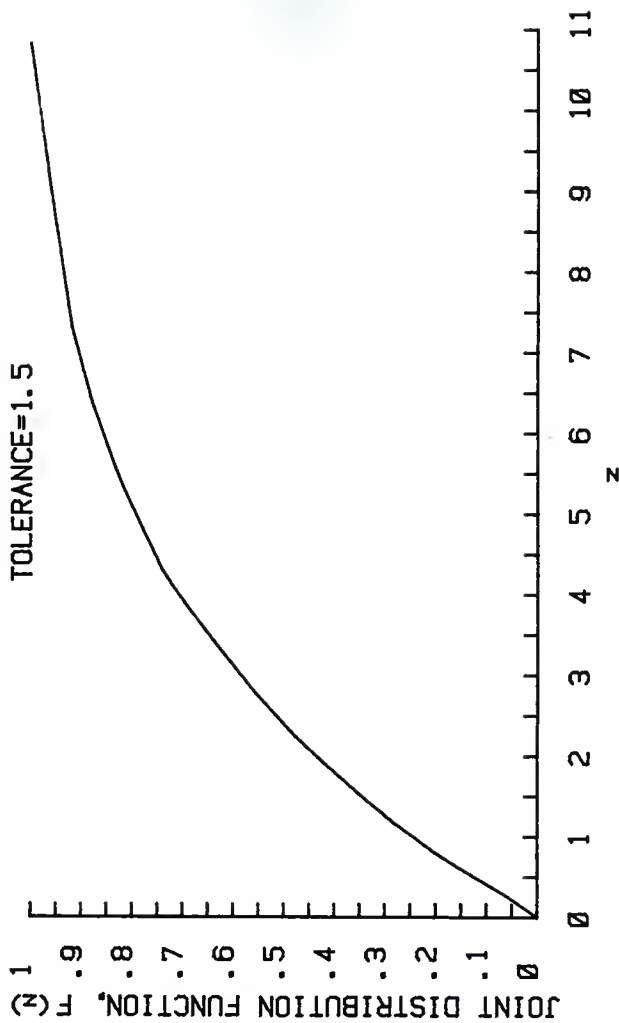


FIGURE 3.8  
 'SMOOTHED' CUBIC SPLINE  
 Z DISTRIBUTION,  $Z=(X*Y^3)/12$   
 TOLERANCE=1.5



The number of cells for the histogram is found using equation 3.9.

$$k = 1 + 3.3 * \log_{10} 25$$

$$k = 6$$

Using equation 3.10, the range is equal to

$$r = 10.8349 - 0.2587$$

$$r = 10.5762$$

And each cell has a width of

$$w = 10.5762/6$$

$$w = 1.7627$$

Figure 3.9 shows the histogram approximation for the function. The cubic spline density approximations are plotted in Figures 3.10 and 3.11. The cubic spline approximations are rough. However, if the number of cells per density function were increased, the cubic spline curves would tend to be smoother with less wiggle

It should be noted that this is a simplified case. In practice there could be an unlimited number of random variables with each random variable having 30 to 200 cells, resulting in literally thousands of calculations.

FIGURE 3.9  
JOINT PROBABILITY  
HISTOGRAM  
 $Z = (X \cdot Y^3) / 12$

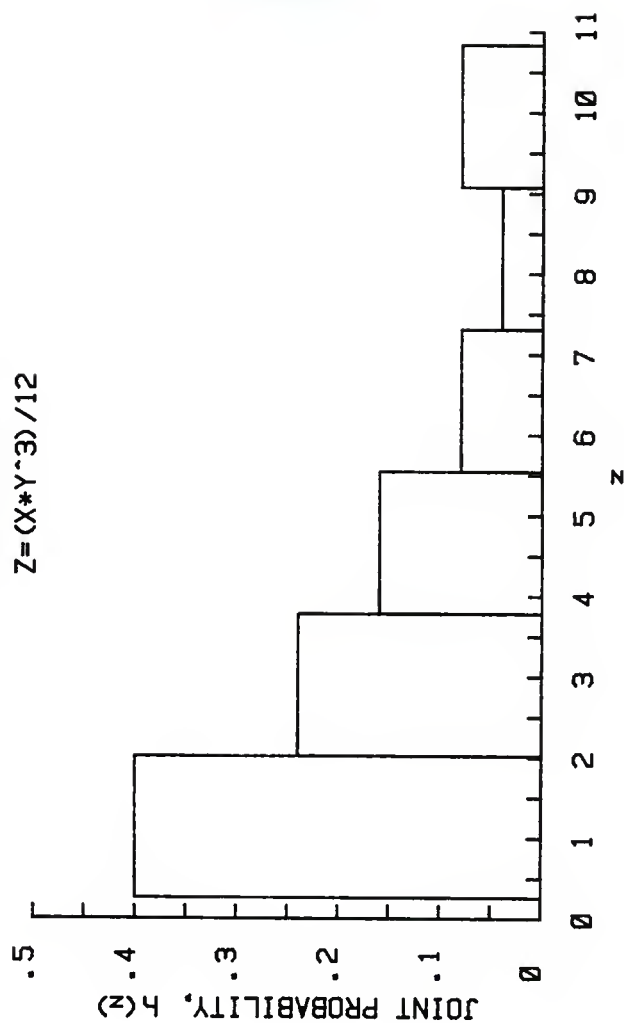


FIGURE 3.10  
CUBIC SPLINE  
DENSITY FUNCTION  
 $Z = (X*Y^3)/12$   
TOLERANCE=1.0

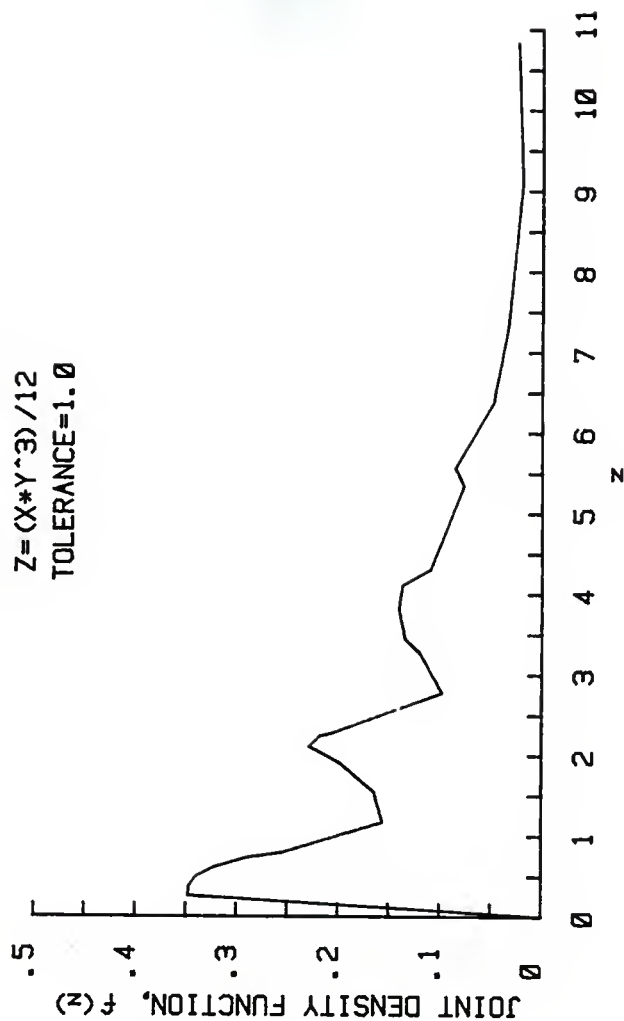
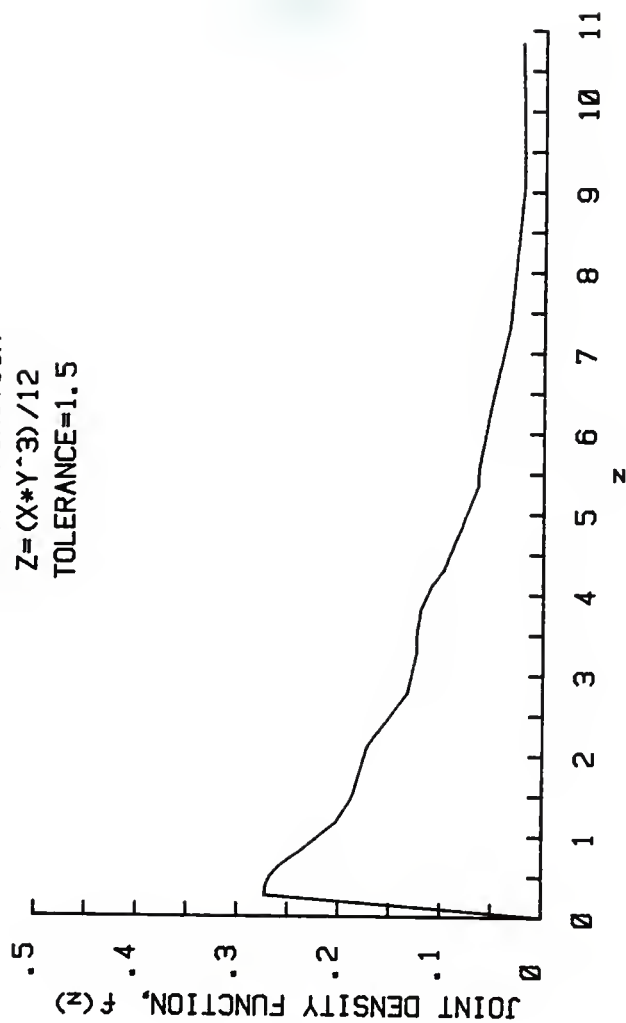


FIGURE 3.11  
CUBIC SPLINE  
DENSITY FUNCTION  
 $Z = (X * Y^3) / 12$   
TOLERANCE = 1.5



## CHAPTER IV

### PROGRAM DESCRIPTION

In this chapter, the computer programs that approximate the joint probability distribution and density functions are discussed. These programs are written in Fortran 77 on a Harris H-800 Super Mini-Computer. Because of storage problems, the simulation package had to be broken down into four separate programs. These programs are:

1. Discrete Simulation
2. Density Function Approximation
3. Data Management
4. Graphics

A complete listing of all the programs can be found in the Appendix.



## DISCRETE SIMULATION

Program Name: DISCSIM

The discrete simulation program approximates the joint probability distribution function of two random variables with known probability density functions. Before executing the program, the desired function must be inserted in the "Function" subroutine. When run, the program prompts the user for the type of distribution of each random variable and the number of discrete points used to approximate each random variable's density function. The density functions available to the user are:

1. Normal: The normal distribution can be used to represent quantities such as measured errors.
2. Triangular: The triangular distribution is used in the absence of data to approximate a rough model of a random variable's density function.
3. Uniform: The uniform distribution is usually used as a first model for a variable that is suspected of being random.

4. Special: The special distribution is used if the user has a distribution function stored on file. e.g. If a distribution function was generated from experimental or measured results.

After the program has approximated the joint probability distribution function, the mean, variance, locations of the discrete points for each of the two random variables, and the joint probability distribution function of the algebraic expression is printed to work file "W2". The joint probability distribution is also printed to a file (designated by the user) that is compatible with the density program, the special distribution subroutine, the data management program, and the plotting programs.

Although this program simulates only two random variables, a joint probability distribution function of an algebraic expression containing more than two random variables can be approximated. To do this, the function must be separated into groups of two random variables. Using the special distribution subroutine, the joint probability distribution function of each group of two random variables can be combined to form the joint probability distribution function of the algebraic

expression. If an expression contains  $n$  random variables, the program must be run  $n-1$  times to simulate the joint probability distribution function of the algebraic expression.

#### DENSITY FUNCTION APPROXIMATION

Program Name: DENSITY

The density function program is a compatible program used to approximate the density function after the joint probability distribution function has been simulated using the discrete simulation program. The density function is approximated using the first derivative of a cubic spline curve fit and a histogram (See Chapter 3). The density program prompts the user for the name of the joint distribution file. The user is able to control the tolerance of the cubic spline curve fit. With careful adjustment of the tolerance, a smooth density curve can be approximated. The histogram data is printed to work file "W1" and the spline data is printed to a file (designated by the user) that is compatible with the plotting programs.

## DATA MANAGEMENT

Program Name: MANAGE

The data management program is used if a problem occurs with the calculation of the cubic spline approximation program. One of the drawbacks of the cubic spline curve fit is that if the data is grouped too close together, the program will halt because of a floating point overflow. The data management program sorts through the data and eliminates any data points that are closer than 0.000001 along the x axis.

## GRAPHICS

Program Names: GRAPH

The plotting program is written using Precision Visuals' DI3000 computer graphics package. Program GRAPH plots the density and distribution functions on a Selanar Hirez 100 graphics terminal. The user is asked for the input data file name, title of the graph, and the X and Y axis labels. The plot is dumped to an Hewlett Packard 7470A plotter.

## CHAPTER V

### COMPARISON

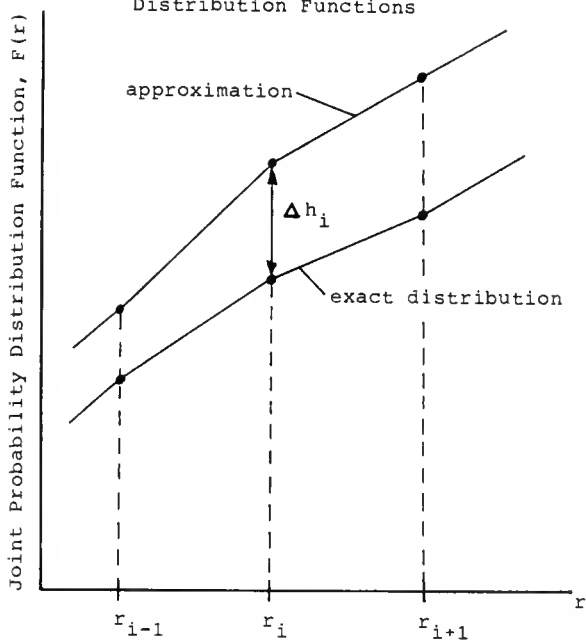
In this chapter, the discrete simulation method discussed in Chapters 3 and 4 is compared to the Monte Carlo method. The joint distribution function of an equation is simulated using both methods and the results are compared to the exact solution of the joint distribution function. The accuracy of each method is measured using two parameters: Mean squared error and maximum deviation.

The mean squared error (MSE) is (See Figure 5.1),

$$MSE = \sum_{i=1}^n \frac{\Delta h_i^2}{n}$$

where  $n$  is the number of  $r$  locations. The maximum deviation is equal to the largest absolute value of  $\Delta h$ .

Figure 5.1  
Approximation and Exact  
Joint Probability  
Distribution Functions



# PROBLEM

Consider the equation of a circle,

$$r = \sqrt{x^2 + y^2} \quad 0 \leq r < \infty$$

where variables  $X$  and  $Y$  are normally distributed independent random variables. Each variable has a mean of zero and a standard deviation equal to one. Find the joint probability distribution function.

## EXACT SOLUTION

The marginal probability density functions of  $X$  and  $Y$  are,

$$f_X(x) = \begin{cases} \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] & \text{for } -\infty \leq x \leq \infty \\ 0 & \text{otherwise} \end{cases}$$

$$f_Y(y) = \begin{cases} \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left[-\frac{(y-\mu)^2}{2\sigma^2}\right] & \text{for } -\infty \leq y \leq \infty \\ 0 & \text{otherwise} \end{cases}$$

When  $\mu_X = 0$ ,  $\sigma_X^2 = 1$  and  $\mu_Y = 0$ ,  $\sigma_Y^2 = 1$ , the marginal

probability density functions become,

$$f_X(x) = \begin{cases} \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{x^2}{2}\right] & \text{for } -\infty \leq x \leq \infty \\ 0 & \text{otherwise} \end{cases}$$

$$f_Y(y) = \begin{cases} \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{y^2}{2}\right] & \text{for } -\infty \leq y \leq \infty \\ 0 & \text{otherwise} \end{cases}$$

Since X and Y are mutually exclusive independent random variables, their joint probability density function is equal to the product of the two marginal probability density functions. Thus,

$$f(x, y) = f_X(x)f_Y(y)$$

So,

$$\begin{aligned} f(x, y) &= \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{x^2}{2}\right] \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{y^2}{2}\right] \\ &= \frac{1}{2\pi} \exp\left[-\frac{(x^2+y^2)}{2}\right] \end{aligned}$$



and the joint probability distribution is,

$$F(x,y) = \iint_D f(x,y) dx dy$$

$$= \frac{1}{2\pi} \int_0^r \int_0^{\sqrt{r^2 - y^2}} \exp\left[-\frac{(x^2 + y^2)}{2}\right] dx dy$$

in polar form this can be written as,

$$F_R(r) = P(R \leq r) = \frac{1}{2\pi} \int_0^{2\pi} \int_0^r \exp\left[-\frac{r^2}{2}\right] r dr d\theta$$

to integrate by parts let,

$$u = r^2/2 \quad \text{and} \quad du = r dr$$

$$P(R \leq r) = \frac{1}{2\pi} \int_0^{2\pi} \int_0^u \exp(-u) du d\theta$$

$$\begin{aligned}
&= \frac{1}{2\pi} \int_0^{2\pi} -\exp(-u) \bigg|_0^u d\theta \\
&= \frac{1}{2\pi} \int_0^{2\pi} \left[ -\exp(-r^2/2) + 1 \right] d\theta \\
&= \frac{1}{2\pi} \left[ -\exp(-r^2/2) + 1 \right] \theta \bigg|_0^{2\pi}
\end{aligned}$$

$$\Pr(R \leq r) = 1 - \exp(-r^2/2)$$

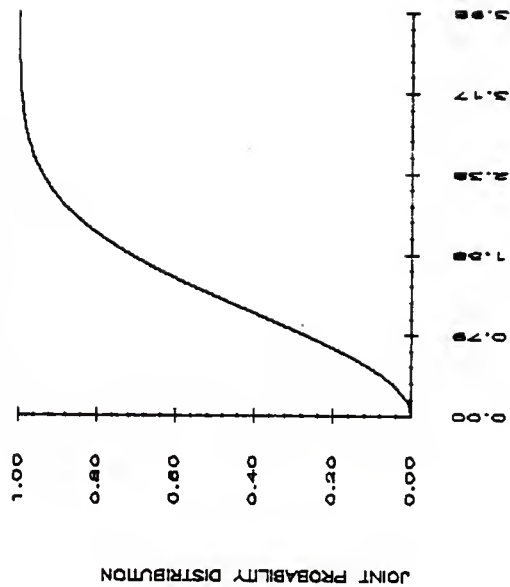
The joint probability distribution function is,

$$F_R(r) = \begin{cases} 0 & r \leq 0 \\ 1 - \exp(-r^2/2) & 0 < r \leq \infty \end{cases}$$

The exact joint probability distribution function is plotted in Figure 5.2

Differentiating the joint probability distribution function yields the joint probability density function.

Figure 5.2  
Exact Joint Probability  
Distribution Function



$$f_R(r) = \begin{cases} 0 & r \leq 0 \\ \text{rexp}(-r^2/2) & 0 < r \leq \infty \end{cases}$$

Figure 5.3 shows a plot of the exact joint probability density function.

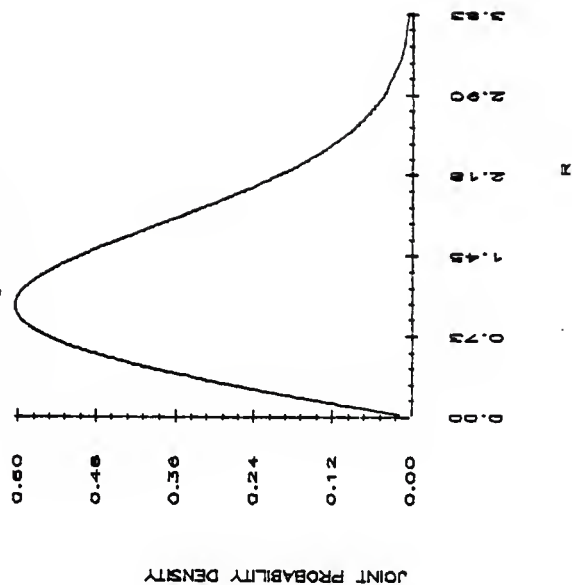
#### MONTE CARLO SIMULATION vs. DISCRETE SIMULATION

A Monte Carlo computer program was developed to compare the accuracies of the two simulations with the exact solution. The Monte Carlo program generates random numbers that are normally distributed to approximate the X and Y random variable's density function. The subroutines used for generating the joint cumulative distribution function are the same subroutines used in the discrete approximation program.

The programs were run three times. The discrete simulation was run first. Each time an increasing number of discrete points was used to represent each random variable's density function. After each run, the number of unique r locations was noted. The Monte Carlo program was then run using the same number of unique r locations. A detailed explanation of the Monte Carlo method is located in Appendix G.

On the first run, 50 discrete points were used to approximate each random variable's density function.

Figure 5.3  
Exact Joint Probability  
Density Function



After all possible combinations were calculated, there were 325 unique  $r$  locations. Figure 5.4 shows the approximated joint probability distribution function of the discrete approximation. In Figure 5.5, the distribution function is plotted along with the exact solution. The first half of the discrete approximation is almost identical to the exact distribution function. However, in the upper 20% of the curve, the discrete approximation deviates quite a bit from the exact solution. The cause of the deviation could be the small number of discrete points used to represent the random variables density function. Because a relatively small number of points were used, the tails of the density functions were not adequately represented. The mean squared error and the maximum deviation for this approximation was 0.000028 and 0.0196446 respectively.

The Monte Carlo Simulation was then run using 325 unique  $r$  locations. The joint distribution function is plotted in Figure 5.6. Figure 5.7 shows the joint probability distribution function for the Monte Carlo approximation and the exact solution. As can be seen, the Monte Carlo approximation is not as good as the discrete approximation. The mean squared error is 0.0014709 and the maximum deviation is 0.0739843.

Figure 5.4  
 Joint Probability Distribution Function  
 for the Discrete Simulation using  
 50 Discrete Points per Random Variable

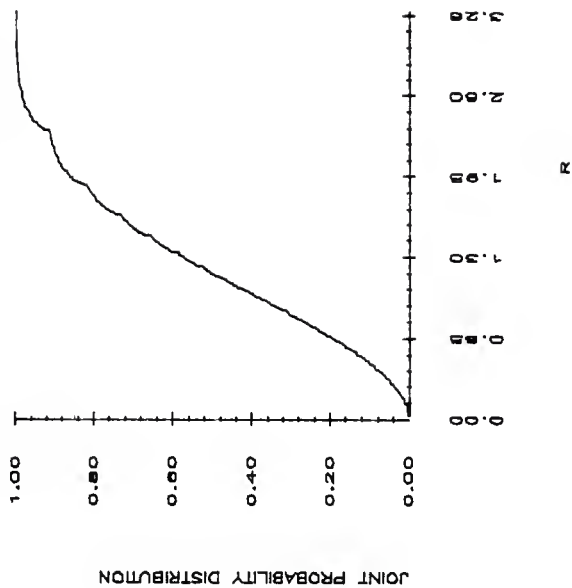


Figure 5.5  
 Joint Probability Distribution Functions  
 for the Discrete Simulation using  
 50 Discrete Points per Random Variable  
 and the Exact Solution

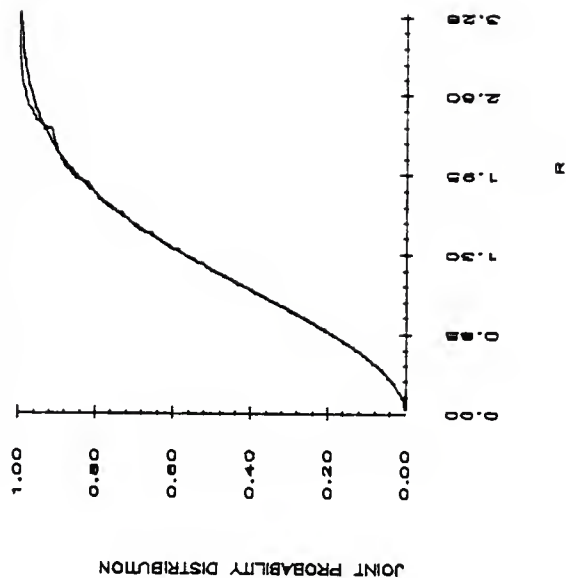




Figure 5.6  
 Joint Probability Distribution Function  
 for the Monte Carlo Simulation using  
 325 Unique R Locations

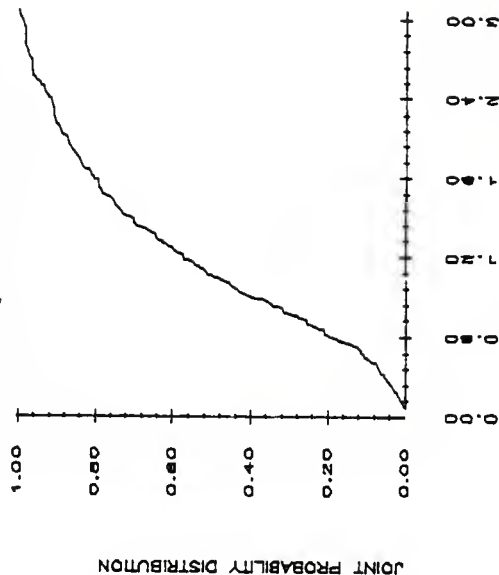
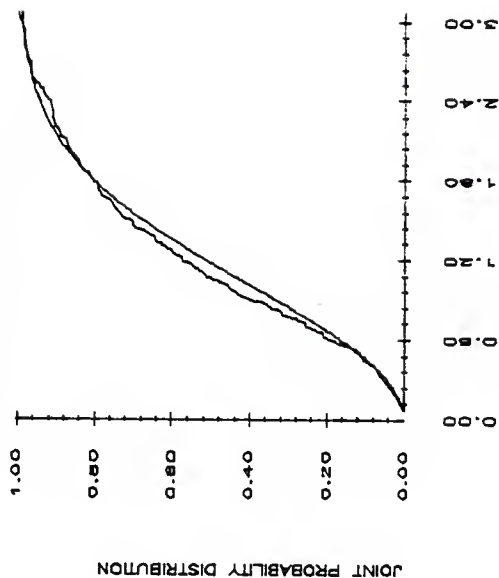


Figure 5.7  
 Joint Probability Distribution Functions  
 for the Monte Carlo Simulation using  
 325 Unique R Locations  
 and the Exact Solution



On the second run, 100 discrete points were used to approximate the X and Y random variable's density functions. This simulation resulted in 1275 unique r locations. Figure 5.8 shows the approximation of the second discrete approximation. In Figure 5.9, the discrete distribution curve is plotted with the exact solution. The curve still has a small dip at the upper portion of the curve. However, the maximum deviation has decreased to 0.0101403 and the mean squared error is 0.0000041. The 100 point approximation is an improvement over the 50 point approximation.

Figure 5.10 shows the Monte Carlo approximation of the joint distribution function for 1275 unique r locations. The distribution curve is plotted with the exact solution in Figure 5.11. Although the accuracy of the Monte Carlo simulation has improved, the approximation is still not as good as the discrete approximation. The maximum deviation is equal to 0.0314508 and the mean squared error is 0.0002180.

On the final run, 200 discrete points were used to approximate each variable's density function. This resulted in 5049 unique r locations. The joint distribution function is plotted in Figure 5.12. Figure 5.13 shows the approximated joint distribution function with the exact solution. The two distribution functions

Figure 5.8  
 Joint Probability Distribution Function  
 for the Discrete Simulation using  
 100 Discrete Points per Random Variable

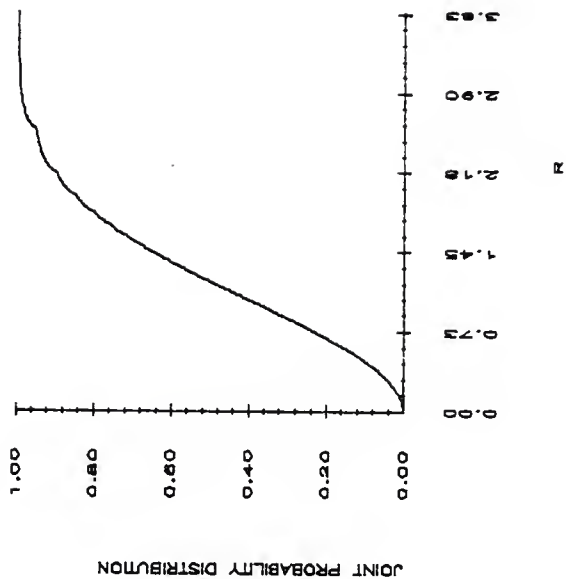


Figure 5.9  
 Joint Probability Distribution Functions  
 for the Discrete Simulation using  
 100 Discrete Points per Random Variable  
 and the Exact Solution

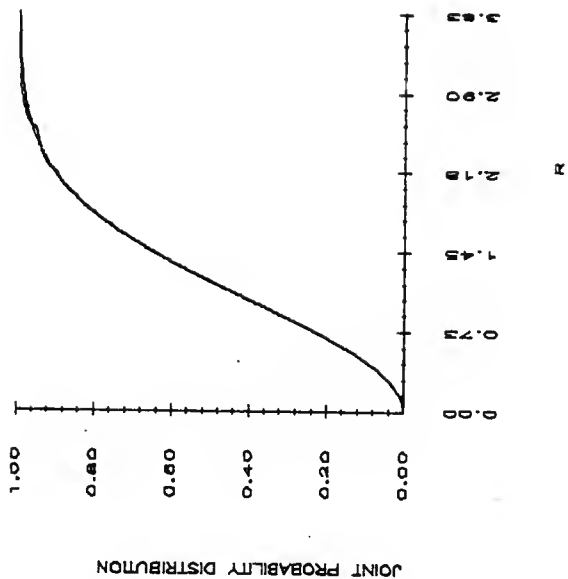


Figure 5.10  
 Joint Probability Distribution Function  
 for the Monte Carlo Simulation using  
 1275 Unique R Locations

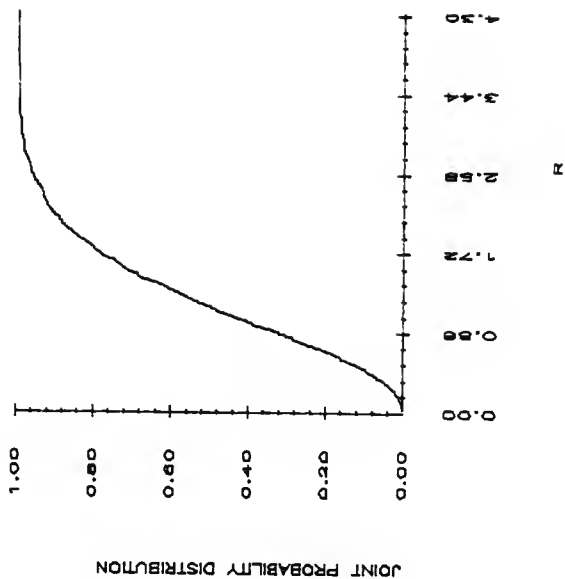


Figure 5.11  
 Joint Probability Distribution Functions  
 for the Monte Carlo Simulation using  
 1275 Unique R Locations  
 and the Exact Solution

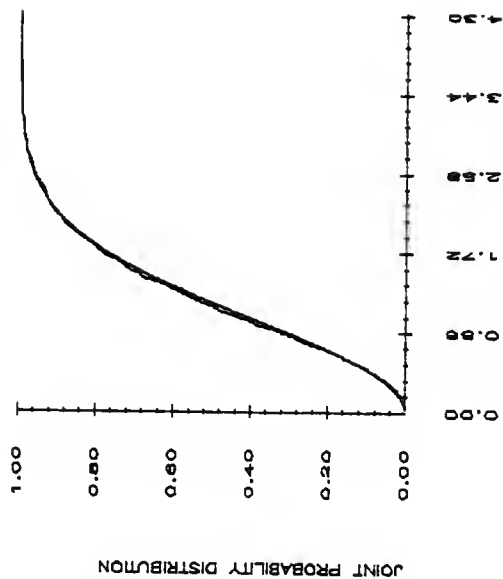


Figure 5.12  
 Joint Probability Distribution Function  
 for the Discrete Simulation Using  
 200 Discrete Points per Random Variable

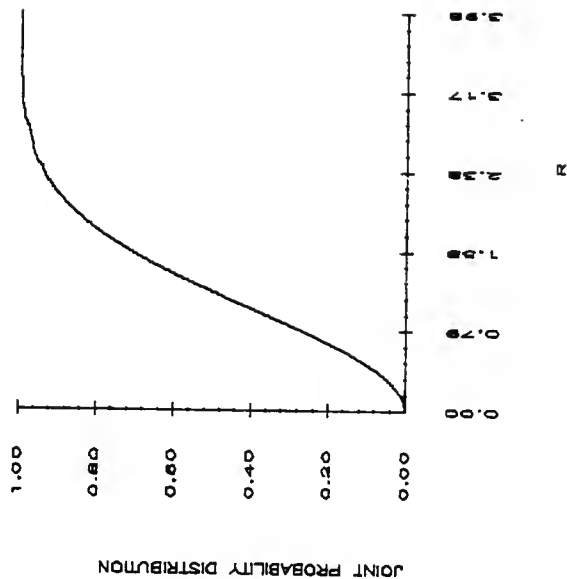
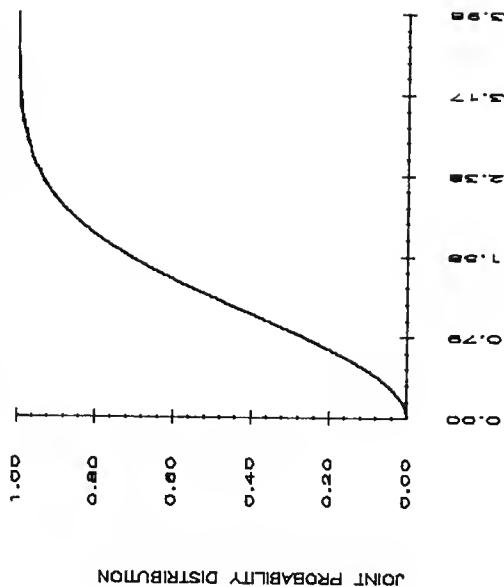




Figure 5.13  
 Joint Probability Distribution Functions  
 for the Discrete Simulation using  
 200 Discrete Points per Random Variable  
 and the Exact Solution



appear to be virtually identical. The dip that was present in the 50 and 100 point approximation is no longer noticeable. The error has been reduced to 0.0000006 and the maximum deviation is 0.0051283.

The Monte Carlo approximation for 5049 unique  $r$  locations is plotted in Figures 5.14 and 5.15. The Monte Carlo approximation has improved with an increase in  $r$  locations. Nevertheless, the discrete approximation is still more accurate than the Monte Carlo approximation. The mean squared error is 0.0000302, which is fifty times larger than the discrete approximation's error. The maximum deviation is three times larger than the discrete approximation at 0.0151759.

In Figure 5.16 and 5.17 the mean squared error and the maximum standard deviation is plotted as a function of unique  $r$  locations for both the discrete simulation program and the Monte Carlo program. In Figure 5.16, the discrete approximation approaches zero with little variation as the number of unique  $r$  locations increases. The Monte Carlo approximation decreased to zero also, however the distribution is scattered and unpredictable. In Figure 5.17, the maximum deviation of the two simulations decrease as the number of unique  $r$  locations increases. Again, the Monte Carlo data is scattered from one value to the next. The general trend of the Monte

Figure 5.14  
 Joint Probability Distribution Function  
 for the Monte Carlo Simulation using  
 5049 Unique R Locations

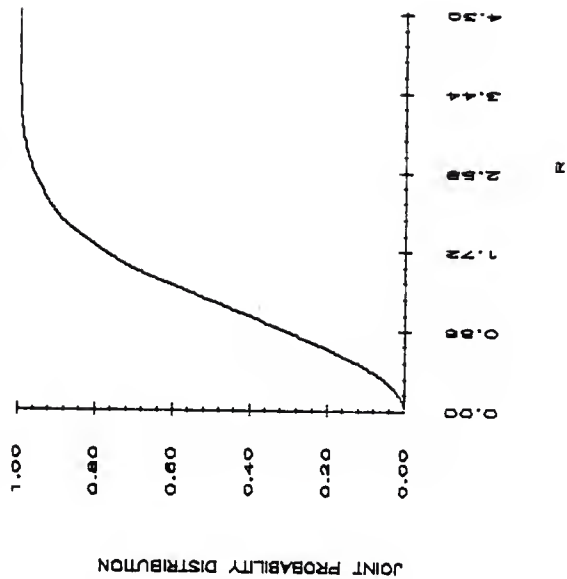


Figure 5.15  
 Joint Probability Distribution Functions  
 for the Monte Carlo Simulation using  
 5049 Unique R Locations  
 and the Exact Solution

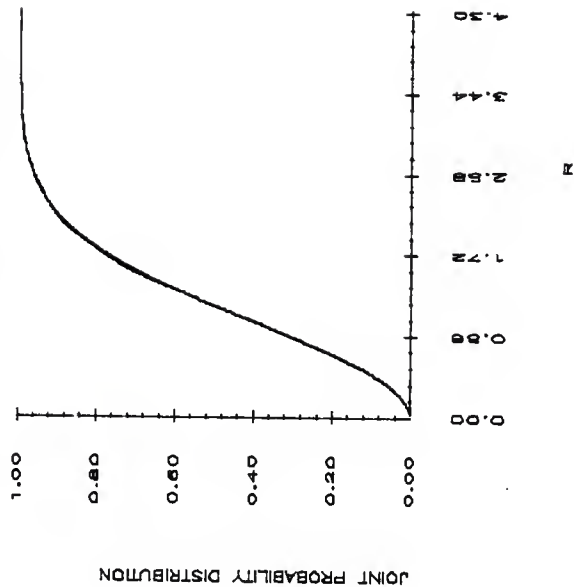


Figure 5.16  
 Mean Squared Error vs.  
 Number of Unique R Locations  
 for the Monte Carlo and Discrete Simulation

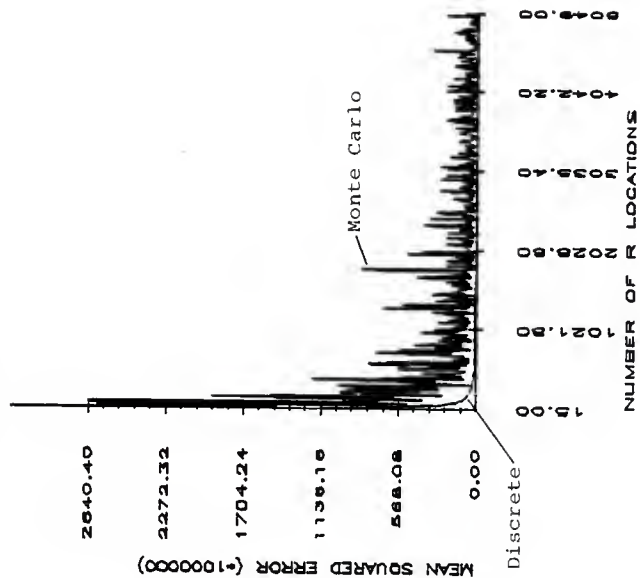
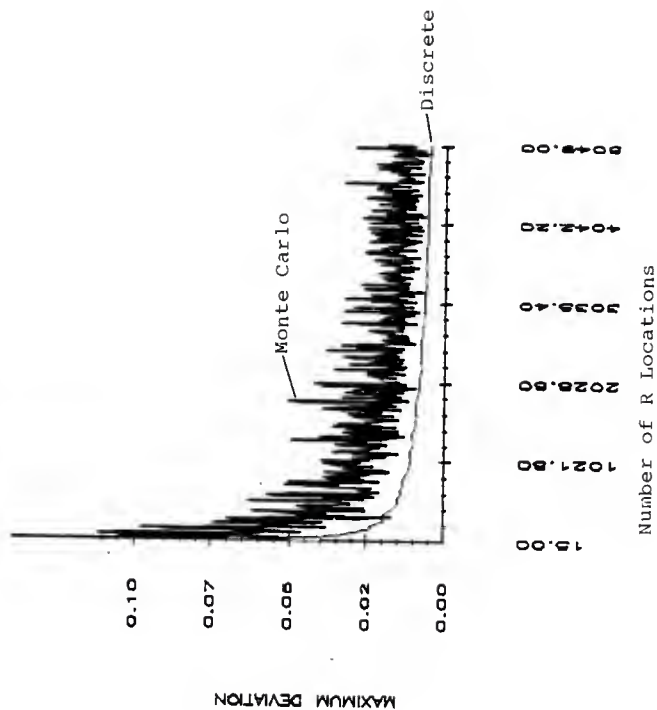


Figure 5.17  
Maximum Deviation vs.  
Number of Unique R Locations  
for the Monte Carlo and Discrete Simulation



Carlo simulation is a decrease in the mean squared error and the maximum deviation as the number of  $r$  locations increase. The large scatter of error and deviation of the Monte Carlo method is caused by the statistical error associated with the generation of the random numbers. Because the generated numbers are random, the variation from one  $r$  location to the next is unpredictable.

In Figures 5.18 and 5.19 the mean squared error and the maximum deviation is plotted again for both the discrete simulation and the maximum deviation. However, in Figures 5.18 and 5.19 these values are plotted as functions of computer processing unit time. The figures are very similar to the previous two plots. The discrete simulation data is constantly decreasing and very predictable, while the Monte Carlo data is quite scattered.

A way to minimize the Monte Carlo scatter is to use a larger number of  $r$  locations. Increasing the number of  $r$  locations would lead to an increase of computing time and money. Therefore, the main advantage of the discrete simulation over the Monte Carlo simulation is that a predictable and more accurate representation of the joint probability distribution function can be approximated using fewer calculations and less computing time.

Figure 5.18  
Mean Squared Error vs.  
Computer Processing Unit Time  
for the Monte Carlo and Discrete Simulation

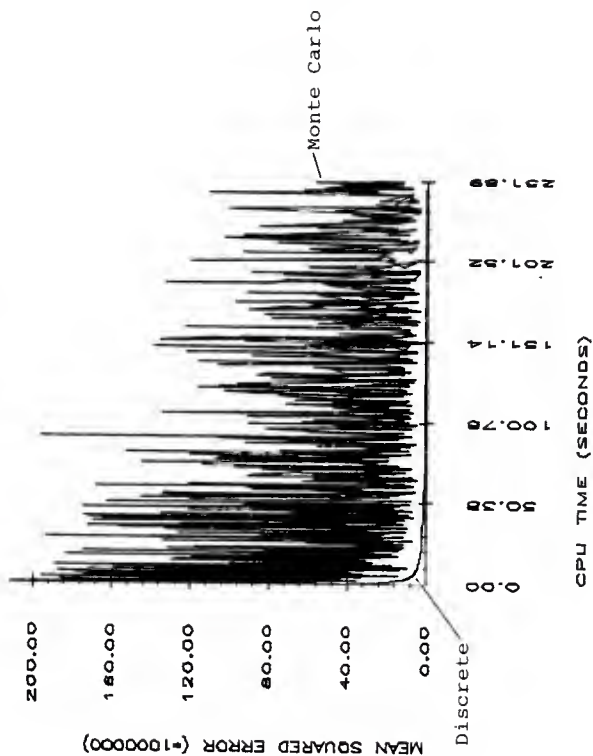
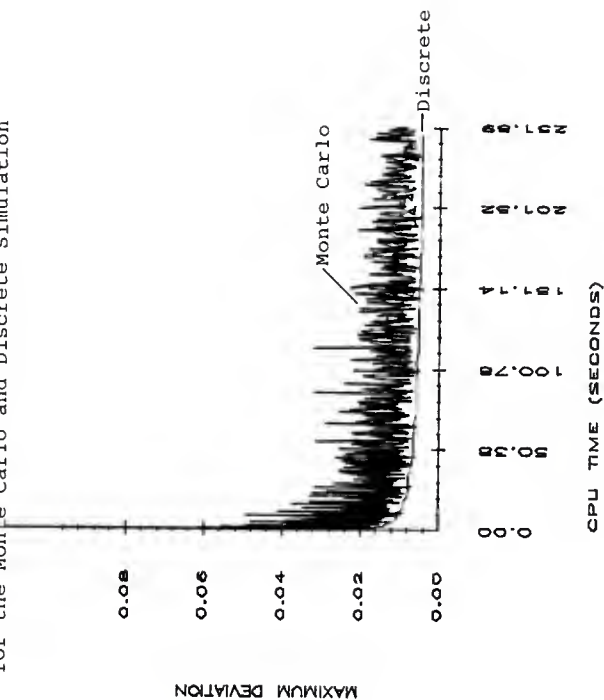




Figure 5.19  
Maximum Deviation vs.  
Computer Processing Unit Time  
for the Monte Carlo and Discrete Simulation



## CHAPTER VI

### CANTILEVER I-BEAM

The beam problem discussed in this section was chosen to illustrate how the discrete simulation program could be used in an engineering application. Each random variable's density function was simulated using 100 discrete points. The "special" distribution subroutine was used to construct the joint probability density functions of algebraic expressions containing more than two random variables.

#### PROBLEM

The cantilever I-beam, shown in Figures 6.1 and 6.2, is made by welding three steel (ASTM A7) plates together. Determine:

a) The joint probability density and distribution functions for:

1. The maximum moment  $M$
2. The deflection at  $B$
3. The slope at  $B$

Figure 6.1a  
Cantilever I-Beam  
Subjected to a Uniform Load,  $W_0$

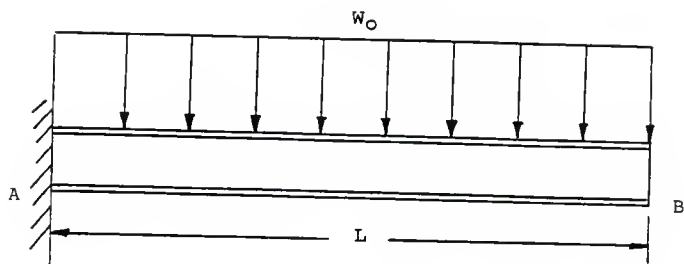


Figure 6.1b  
Free-Body Diagram  
of a Cantilever I-Beam  
Subjected to a Uniform Load,  $W_0$

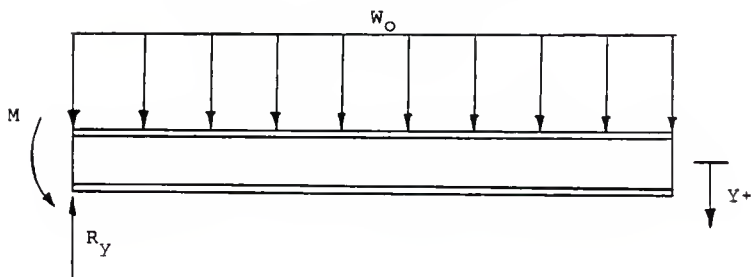
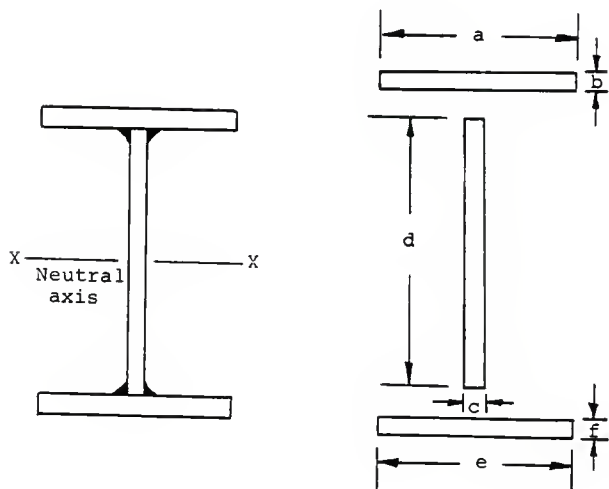


Figure 6.2  
Cross Section of the Cantilever I-Beam



4. The maximum stress
  5. The maximum shear
- b) The reliability of the I-beam.

The dimensions and metal properties of the I-beam are assumed to be random variables with known probability density functions. These values are listed in Table 6.1.

Table 6.1  
Random Variables for a Cantilever I-Beam

Random Variable	Mean	Standard Deviation	Units	Distribution
a	6.00	0.2500	inches	Normal
b	0.50	0.0625	inches	Normal
c	0.50	0.0625	inches	Normal
d	8.00	0.2500	inches	Normal
e	6.00	0.2500	inches	Normal
f	0.50	0.0625	inches	Normal
E	30E6	4.5E5	psi	Normal
L	16.00	1.0000	feet	Normal
W <sub>o</sub>	45.00	33.3333	lb/inch	Uniform
S <sub>u</sub>	65.6E3	2.43E3	psi	Normal
S <sub>y</sub>	42.7E3	4.86E3	psi	Normal

# SOLUTION

a) 1. The maximum moment for a cantilever beam with a uniform load can be expressed as,

$$M_{\max} = \frac{W_0 L^2}{2} \quad (6.1)$$

Figures 6.3 and 6.4 show the joint probability distribution and density functions of the maximum moment. The maximum moment ranges from 59,877.02 to 78,638.19 foot-pounds with a mean value of 69,121.85 foot-pounds.

2. The maximum deflection at point B is equal to

$$Y_{\max} = \frac{W_0 L^4}{8EI} \quad (6.2)$$

Before the deflection can be calculated, the joint probability distribution function of the moment of inertia must be approximated. Using the Parallel-Axis Theorem, the moment of inertia can be expressed as the equation,

$$I = (ab^3)/12 + ab(d/2 + b/2)^2 + (cd^3)/12 + (ef^3)/12 + fe(f/2 + d/2)^2 \quad (6.3)$$

Figure 6.3  
Joint Probability Distribution Function  
of the Maximum Moment  
of a Cantilever I-Beam

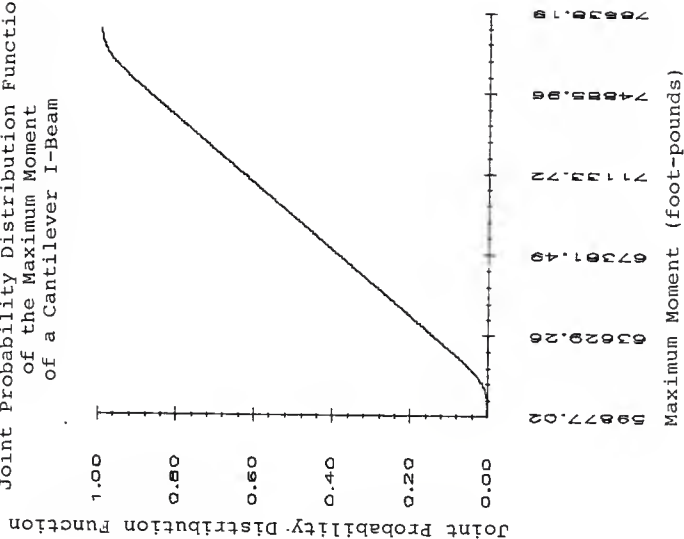
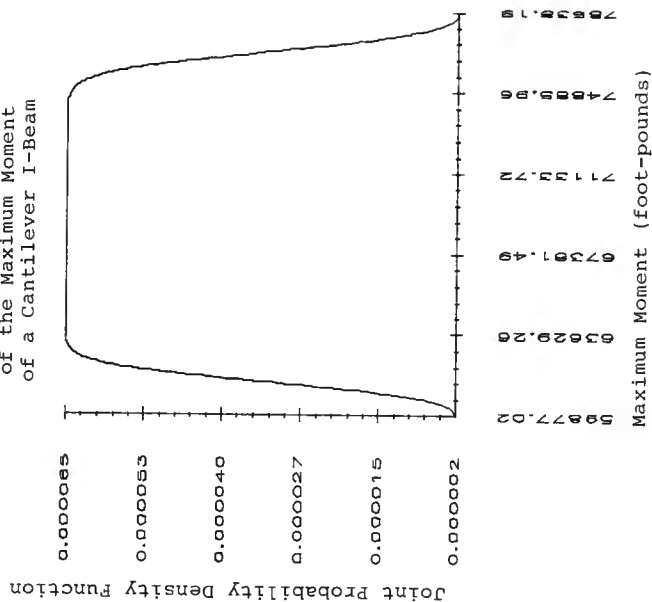


Figure 6.4  
Joint Probability Density Function  
of the Maximum Moment  
of a Cantilever I-Beam





Figures 6.5 and 6.6 show the probability distribution and density functions of the moment of inertia. Once the moment of inertia's distribution has been determined, the joint probability density and distribution functions for the maximum deflection are approximated. Figures 6.7 and 6.8 show the joint probability distribution and density functions of the deflection at point B. The expected value of the deflection is 1.977 inches with a minimum possible value of 1.4 and a maximum value of 2.82 inches.

3. The slope  $\phi$ , at point B can be written as,

$$\phi = \frac{W_0 L^3}{6EI} \quad (6.4)$$

The joint probability distribution and density functions are plotted in Figures 6.9 and 6.10. The discrete simulation predicts that the mean value of the slope at point B is 0.786 degrees, and the range of possible slope values is from 0.56 to 1.11 degrees.

4. The maximum stress of the cantilever I-beam occurs at the furthest distance from the neutral axis. The equation for the maximum stress is

$$S_{\max} = \frac{M(d/2+f)}{I} \quad (6.5)$$

Figure 6.5  
Joint Probability Distribution Function  
of the Moment of Inertia  
of a Cantilever I-Beam

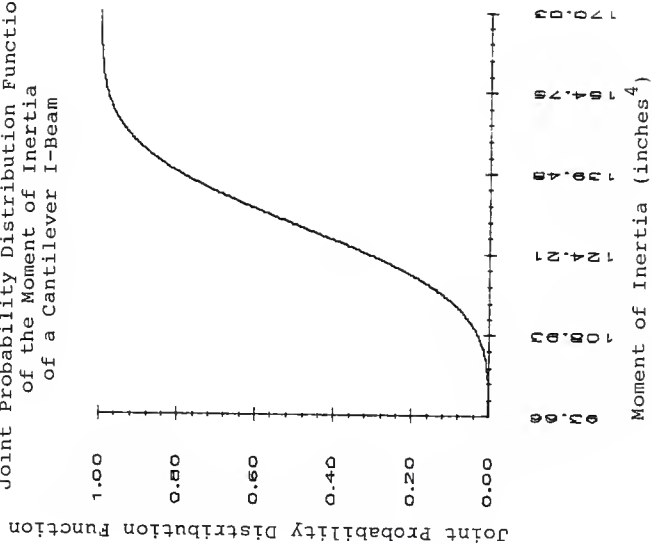


Figure 6.6  
Joint Probability Density Function  
of the Moment of Inertia  
of a Cantilever I-Beam

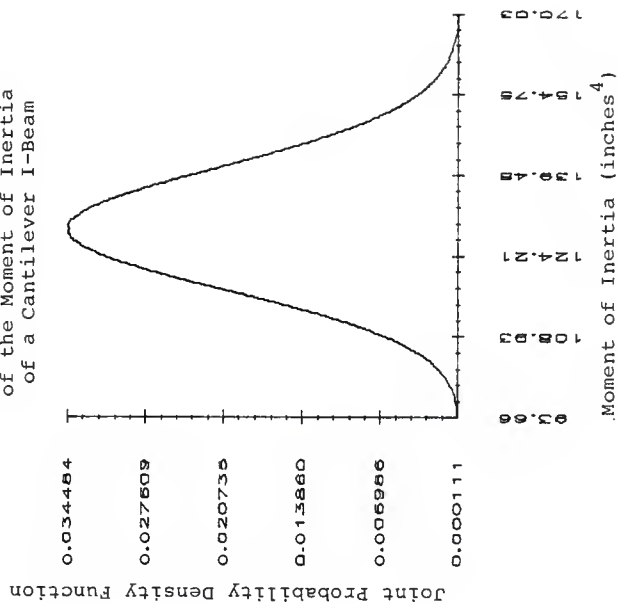


Figure 6.7  
Joint Probability Distribution Function  
of the Maximum Deflection  
of a Cantilever I-Beam

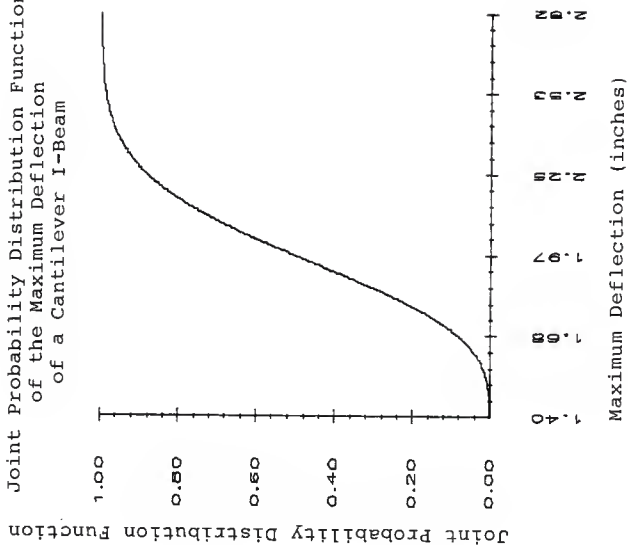
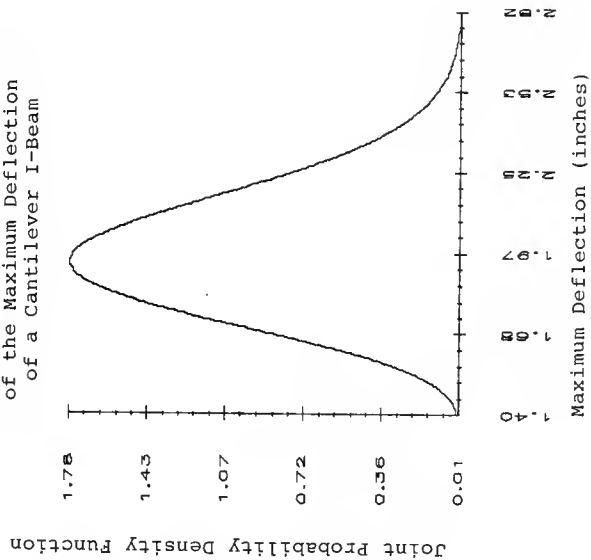


Figure 6.8  
Joint Probability Density Function  
of the Maximum Deflection  
of a Cantilever I-Beam



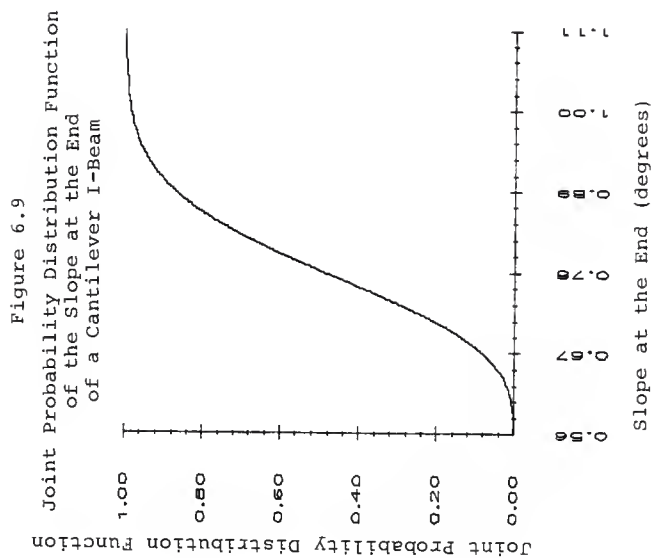
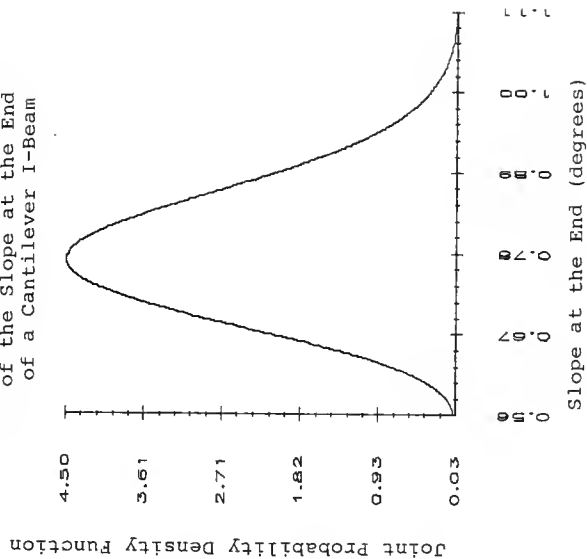


Figure 6.10  
Joint Probability Density Function  
of the Slope at the End  
of a Cantilever I-Beam



The expected value of the maximum stress is 28,944.26 psi. It is interesting to note that the range of possible maximum stress values differ by as much as 22,000 psi. The joint probability density and distribution functions are plotted in Figures 6.11 and 6.12.

5. The maximum value of the first moment of area occurs at the neutral axis. Since the width of the beam is thin at the neutral axis, the maximum shearing stress occurs at this point. The equation for the maximum shear stress can be written as

$$\tau_{\max} = \frac{W_0 L (ab(b/2 + d/2) + cd^2/8)}{I_c} \quad (6.6)$$

The expected value of the maximum shearing stress is 2,280.38 psi, and the range of possible shear values is from 1,171 to 4,441 psi. The joint probability density and distribution functions are plotted in Figures 6.13 and 6.14.

b) The reliability of the beam is equal to unity minus the probability of failure. If the stress in the beam exceeds the yield stress, the beam is presumed to have failed. Using this criteria, if the ratio of the maximum



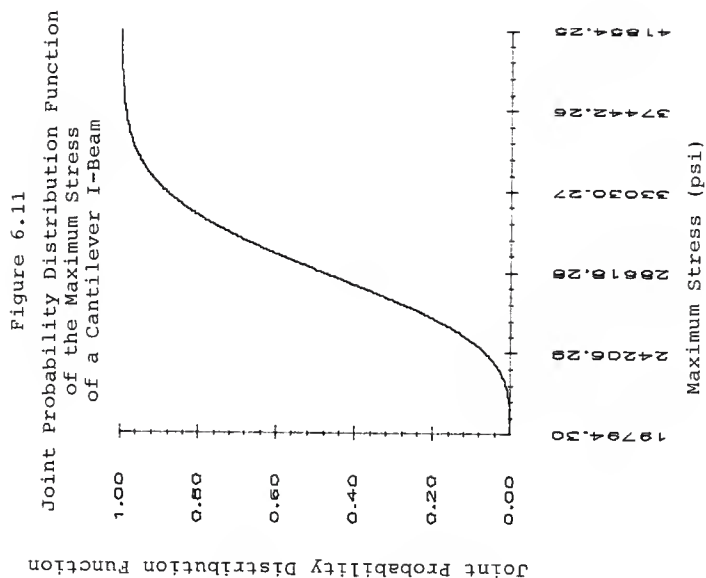


Figure 6.12  
Joint Probability Density Function  
of the Maximum Stress  
of a Cantilever I-Beam

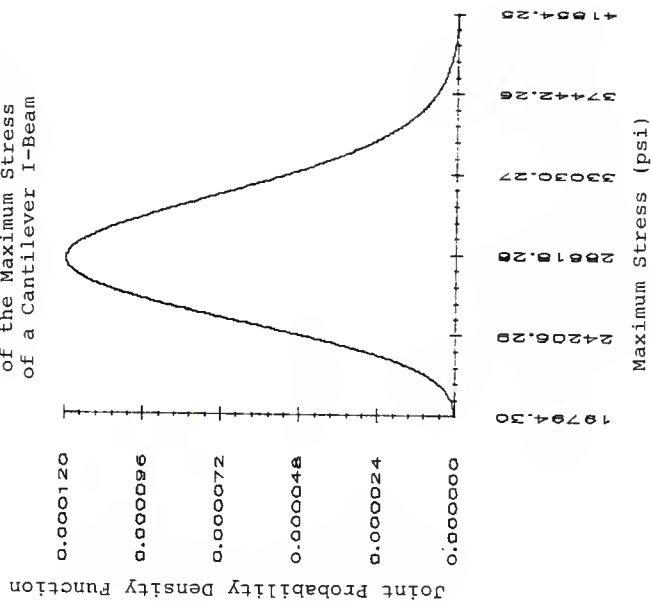


Figure 6.13  
Joint Probability Distribution Function  
of the Maximum Shear Stress  
of a Cantilever I-Beam

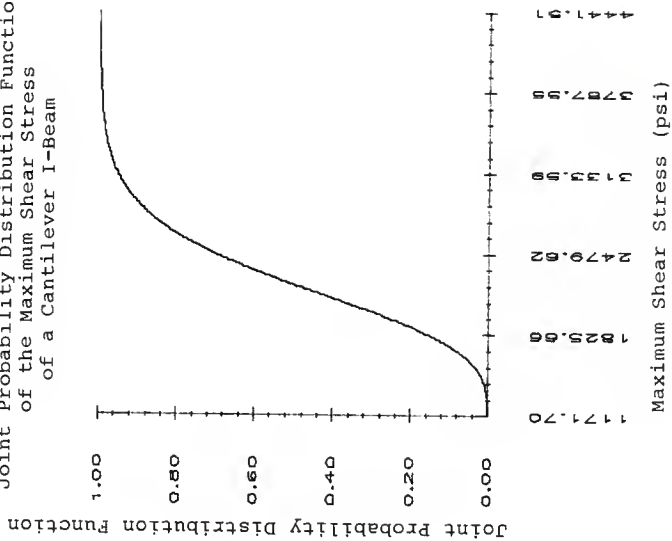
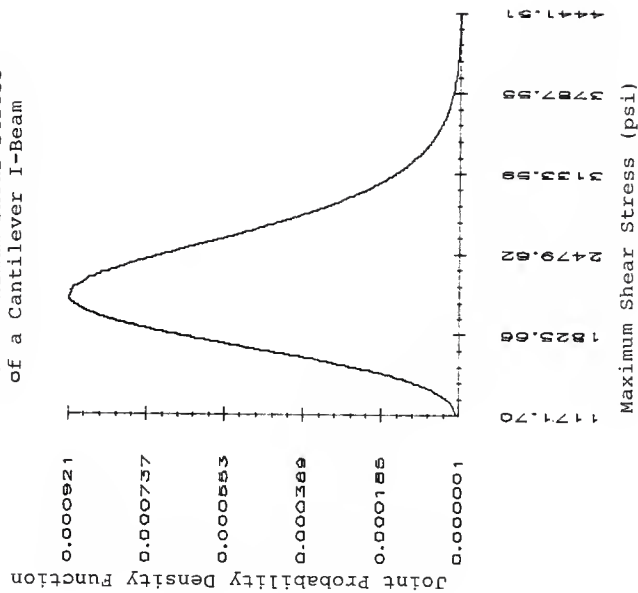


Figure 6.14  
Joint Probability Density Function  
of the Maximum Shear Stress  
of a Cantilever I-Beam



stress in the beam to the yield stress is greater than one, failure occurs. The joint probability and density functions for the stress ratio are plotted in Figures 6.15 and 6.16. Using the joint probability distribution, the reliability of the system can be determined. From Figure 6.15 the probability of the stress ratio being less than or equal to one is equal to approximately 98%. Therefore, the probability of failure is roughly 2% and the probable reliability of the I-beam structure is 98%.

Figure 6.15  
Joint Probability Distribution Function  
of the Ratio  $S_{act}/S_{yield}$   
of a Cantilever I-Beam

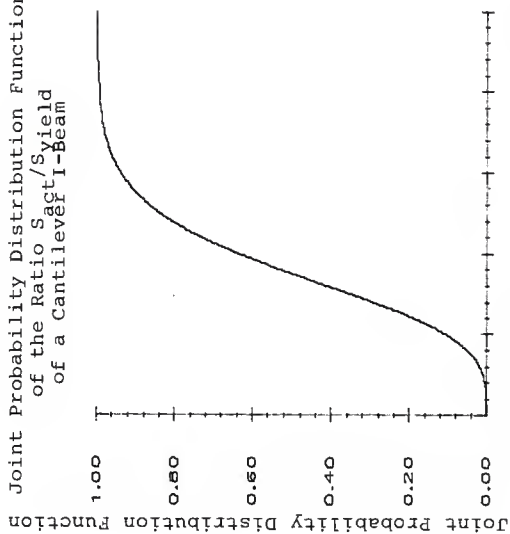
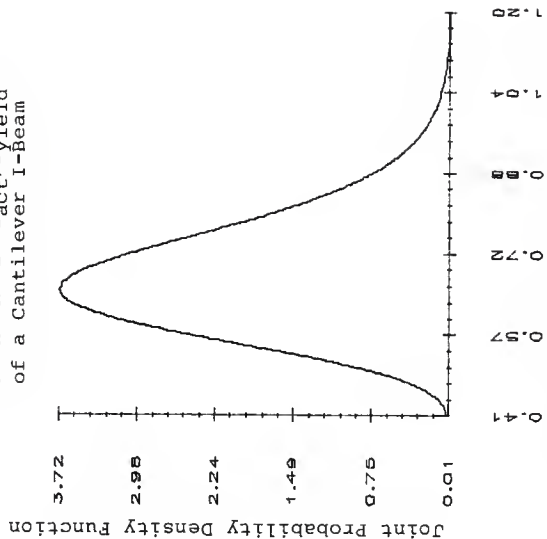


Figure 6.16

Joint Probability Density Function  
of the Ratio  $\text{Sact}/\text{Syield}$   
of a Cantilever I-Beam



Stress in the Beam divided by Yield Stress

## CHAPTER VII

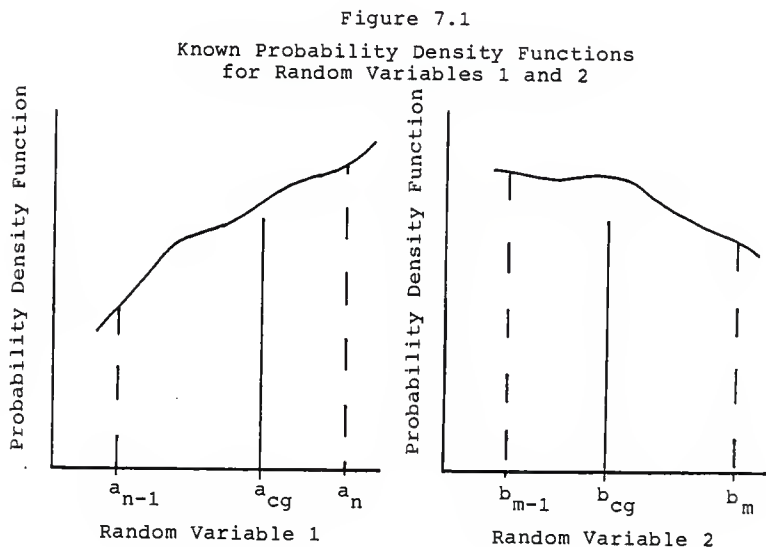
### BOUNDING

During the development of the discrete simulation algorithm, a new technique was investigated that approximates the upper and lower bounds of a joint probability distribution function of an algebraic expression containing several independent random variables. Although a formal proof of the bounding theory has not been written, the results of the program have been compared with several known solutions with good results. In this chapter, the procedure used to obtain the upper and lower bounds of a joint probability distribution function is discussed.

Using the same method of cell development as the discrete simulation technique discussed in Chapter 3, two curves are developed that bound the exact joint probability distribution function of an algebraic expression. The bounding approximation divides the two known random variables' density functions into a given



number of cells. The upper and lower bounds of each cell are noted as well as the cell's center of gravity (See Figure 7.1).



M and n are arbitrary cell divisions within the random variables',  $X_1$  and  $X_2$ , known density functions. Since each cell is represented using three locations,

there are nine possible outcomes that need to be considered.

$$Z_1 = Z(X_1(a_{n-1}), X_2(b_{m-1})) \quad (7.1)$$

$$Z_2 = Z(X_1(a_{n-1}), X_2(b_{cg}))$$

$$Z_3 = Z(X_1(a_{n-1}), X_2(b_m))$$

$$Z_4 = Z(X_1(a_{cg}), X_2(b_{m-1}))$$

$$Z_5 = Z(X_1(a_{cg}), X_2(b_{cg}))$$

$$Z_6 = Z(X_1(a_{cg}), X_2(b_m))$$

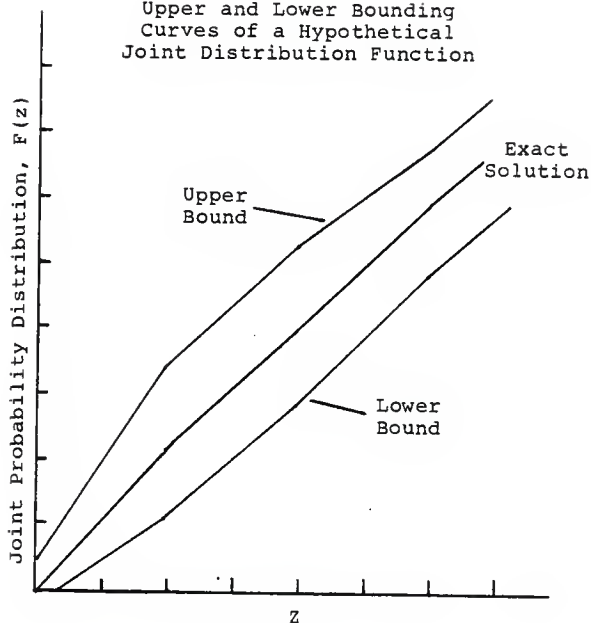
$$Z_7 = Z(X_1(a_n), X_2(b_{m-1}))$$

$$Z_8 = Z(X_1(a_n), X_2(b_{cg}))$$

$$Z_9 = Z(X_1(a_n), X_2(b_m))$$

The nine values are compared with each other. The maximum value is stored in the lower Z bounding array and the minimum value is stored in the upper Z bounding array. Each bounding distribution is then assembled independently using the same distribution and sort subroutines used for the discrete approximation. Figure 7.2 shows the upper and lower bounds of a hypothetical joint distribution function.

Figure 7.2  
Upper and Lower Bounding  
Curves of a Hypothetical  
Joint Distribution Function



The bounding program was compared with the exact solution of the probabilistic pythagorean random variable discussed in Chapter 5. The bounding approximation bounded both the exact solution as well as the approximated joint probabilistic distribution function.

Figure 7.3 shows the exact solution and the upper and lower bounding curves using 100 cells per random variable.

The bounding routine was also compared to Wu's fast probability integration (FPI) approximation 23 of the probability of failure of a bar in tension. The probability of failure,  $P_f$  is symbolized as

$$P_f = P(g \leq 0) \quad (7.2)$$

where  $g$  is equal to the equation

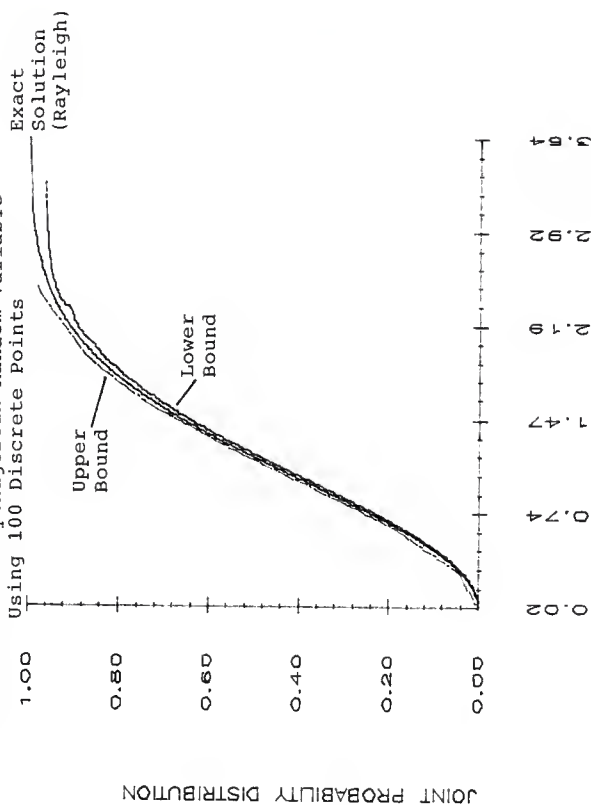
$$g = R - (4T/\pi D^2) \quad (7.3)$$

$R$  and  $D$  are independent normal random variables.  $R$  has a mean of 170.0 N/mm and a variance of 625.0 N/mm.  $D$  has a mean and variance of 29.4 mm and 9.0 mm respectively.  $T$  is equal to 50 kN. Table 7.1 lists Wu's approximation as well as the exact solution and the upper and lower bounding approximations. An example of the output using ten cells per random variable is listed in the Appendix.

Table 7.1 shows that as the number of cell divisions per density function increase for each random variable,

Figure 7.3

The Exact Solution and the Bounding Curves for the Probabilistic Pythagorean Random Variable Using 100 Discrete Points



the upper and lower bounds converge toward the exact solution.

Table 7.1

Bounding Probabilities of Failure of a Bar in Tension				
Number of Discrete Points/ Density Function	Lower Bound	Exact Solution	FPI Approx- imation	Upper Bound
100	0.00030	0.00230	0.00229	0.0114
120	0.00032			0.0098
150	0.00043			0.0083
200	0.00061			0.0068

The subroutines, Bookbound and Outputbound, that approximate the upper and lower bounds are listed in the Appendix.

## CHAPTER VIII

### CONCLUSIONS

The purpose of this study was to develop a computer simulation package that is capable of approximating the joint probability distribution function of a complex algebraic expression containing several independent random variables. A discrete simulation technique was developed and implemented on a Harris H-800 Super Mini-Computer.

In Chapter 3 the discrete simulation technique is discussed. Although the discrete simulation method is similar to the Monte Carlo technique, the discrete method divides each random variable's known density function into a given number of  $n$  discrete sections with equal areas. Each area is then represented by a single discrete point with a probability of  $1/n$ . Using the discrete points, all possible combinations of the algebraic expression are simulated. After all the outcomes have been computed, the joint distribution

function is assembled by sorting the data and summing the ordered values probabilities. The approximate shape of the density function is simulated using cubic splines and a histogram.

The computer programs that are used to approximate the joint probability distribution and density functions are described in Chapter 4. Because of storage problems, the simulation package had to be divided into four separate programs. The programs are listed in the Appendix.

In Chapter 5 the discrete approximation is compared to the traditional Monte Carlo technique for a joint probability distribution that can be solved exactly. The joint probability distribution function of the equation of a circle is approximated using both methods and the results are compared to the exact solution of the joint distribution function. The programs were run three times. Each time an increasing number of points was used to represent each random variable. The joint probability distributions are plotted vs. the exact solution for each case. As expected, the accuracy of each method improves with an increase of number of points used to represent each random variable's density function. However, the accuracy of the discrete simulation is better for each



case than the Monte Carlo simulation. The mean squared error and the maximum deviation is plotted as a function of number of unique outcomes and computer processing unit (CPU) time. The discrete simulation plots are very well behaved and consistently decreasing with an increasing number of outcomes and CPU time, while the Monte Carlo data is scattered and erratic.

In Chapter 6 an example of how the discrete simulation method can be used as a tool by engineers to aid in the design and development of structures and components is presented. A cantilever I-beam is analyzed. Each variable is assumed to be random with a known probability distribution. Using the discrete simulation programs, the joint probability and density functions are approximated for the maximum moment, the deflection at the end of the beam, the slope at the end of the beam, the maximum stress, the maximum shear, and the reliability of the beam.

In Chapter 7 a different technique is discussed that determines the upper and lower bounds of a joint probability distribution function using the discrete approximation. Although a formal proof has not been developed for the bounding theory, the program was run for several known cases with good results.

In summary, a discrete simulation algorithm was developed to approximate the joint probability distribution function of complex algebraic expressions containing several independent random variables. Although there are a few methods available for approximating the joint probability distribution functions of algebraic expressions, they usually involve complex mathematical operations that are too cumbersome for the average design engineer. The discrete simulation method developed in this study does not involve any complex mathematical operations, is easy to use, and is capable of being used in a wide variety of engineering applications.

## CHAPTER IX

### RECOMMENDATIONS FOR FURTHER STUDY

There are many possible extensions of the discrete simulation method that can be investigated following this study. A few suggestions are to:

1. Develop a faster sorting routine which would decrease computing time.
2. Expand the program to allow the use of more discrete points to represent each random variable's density function.
3. Modify the program so that the variables are represented using double precision.
4. Investigate the accuracy of the discrete simulation program for an algebraic expression containing more than two random variables in which an exact solution exists.
5. Develop a formal proof of the bounding theory discussed in Chapter 7.

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APPENDIX A

DISCRETE SIMULATION PROGRAM



```

C *****
C *
C * PROGRAM NAME: DISCSIM (DISCRETE SIMULATION AND BOUNDING)
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 26, 1987
C *
C * PURPOSE: PROGRAM "DISCSIM" APPROXIMATES THE JOINT PROBABILITY
C * DISTRIBUTION FUNCTION OR THE UPPER AND LOWER BOUNDING
C * CURVES OF A DISTRIBUTION FUNCTION FOR AN ALGEBRAIC
C * EXPRESSION CONTAINING TWO INDEPENDENT RANDOM VARIABLES.
C * THE RANDOM VARIABLE'S DENSITY FUNCTIONS ARE APPROXIMATED
C * USING A FINITE NUMBER (2-200) OF DISCRETE LOCATIONS. THE
C * FUNCTION MUST BE INSERTED IN SUBROUTINE "FUNCTION".
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: SUBROUTINES BOOK
C * BOOKBOUND
C * CELLR
C * DIST
C * DISTRIBUTION
C * FUNCTION
C * NORMAL
C * NORMINV
C * OUTPUT
C * OUTPUTBOUND
C * SORT
C * SPECIAL
C * TRIANGLE
C * UNIFORM
C *
C * DESCRIPTION OF PARAMETERS:
C * INPUT ARGUMENTS:
C * ANSB - CHARACTER*1 - IF ANSB IS EQUAL TO 'Y' THEN THE
C * PROGRAM COMPUTES THE UPPER AND LOWER BOUNDS OF
C * THE JOINT DISTRIBUTION FUNCTION AS WELL AS THE
C * DISCRETE APPROXIMATION FUNCTION. IF ANSB IS
C * EQUAL TO 'N' THEN THE PROGRAM COMPUTES ONLY THE
C * DISCRETE APPROXIMATION OF THE JOINT
C * DISTRIBUTION FUNCTION.
C * DISREP(NUM) - CHARACTER*10 - DISREP CONTAINS THE
C * TYPE OF PROBABILITY DISTRIBUTION FOR VARIABLE
C * NUM (NORMAL, UNIFORM, TRIANGULAR, SPECIAL).
C * DISTR(NUM) - INTEGER - THE NUMBER OF THE DISTRIBUTION
C * FOR RANDOM VARIABLE NUM (1=NORMAL, 2=UNIFORM,
C * 3=TRIANGULAR, 4=SPECIAL).
C * NAME - CHARACTER*30 - NAME IS THE INPUT DATA FILE NAME
C * OF THE SPECIAL DISTRIBUTION FUNCTION FOR

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C *          RANDOM VARIABLE NUM. *
C * NAME2 - CHARACTER*30 - NAME2 IS THE OUTPUT DATA FILE *
C * NAME OF THE JOINT DISTRIBUTION FUNCTION. *
C * NCELLS(NUM) - INTECER - THE NUMBER OF DISCRETE *
C * LOCATIONS THAT APPROXIMATE THE PROBABILITY *
C * DENSITY FUNCTION OF RANDOM VARIABLE NUM. *
C * NUM - INTEGER - THE NUMBER OF THE RANDOM VARIABLE. *
C * VARIANCE - REAL - THE VARIANCE VALUE FOR RANDOM *
C * VARIABLE NUM. *
C *
C * OUTPUT ARGUMENTS: *
C * A(N) - REAL - WHERE N=NCELLS*2, ODD VALUES OF N *
C * REPRESENT THE CENTER OF GRAVITY LOCATIONS AND *
C * EVEN VALUES OF N INDICATE THE CELL BOUNDARIES. *
C * B(NCELLS,NUM) - REAL - THE DISCRETE POINT (CENTER OF *
C * GRAVITY) REPRESENTATIONS FOR NUM'S PROBABILITY *
C * DENSITY FUNCTION. *
C * COUNT - INTECER - COUNT IS THE NUMBER OF UNIQUE Z *
C * LOCATIONS OF THE JOINT PROBABILITY DISTRIBUTION *
C * FUNCTION. *
C * LB(NCELLS( ),2) - REAL - CONTAINS THE LOWER BOUND *
C * LOCATION FOR CELL NCELLS( ). *
C * LBCOUNT - INTEGER - LBCOUNT IS THE NUMBER OF UNIQUE *
C * LB( , ) LOCATIONS OF THE LOWER BOUND *
C * DISTRIBUTION FUNCTION. *
C * MEAN(NUM) - REAL - THE MEAN VALUE FOR RANDOM VARIABLE *
C * NUM. *
C * UB(NCELLS( ),2) - REAL - CONTAINS THE UPPER BOUND *
C * LOCATION FOR CELL NCELLS( ). *
C * UBCOUNT - INTEGER - UBCOUNT IS THE NUMBER OF UNIQUE *
C * UB( , ) LOCATIONS OF THE UPPER BOUND *
C * DISTRIBUTION FUNCTION. *
C * Z(COUNT,1) - REAL - THE Z VALUE OF THE JOINT *
C * PROBABILITY DISTRIBUTION FUNCTION. *
C * Z(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS *
C * THAN OR EQUAL TO Z(COUNT,1). *
C * Z(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF *
C * TIMES Z(COUNT,1) WAS REPEATED DURING THE *
C * SIMULATION. *
C * ZLB(COUNT,1) - REAL - THE Z VALUE OF THE LOWER BOUND *
C * OF THE JOINT PROBABILITY DISTRIBUTION FUNCTION. *
C * ZLB(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS *
C * THAN OR EQUAL TO ZLB(COUNT,1). *
C * ZLB(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF *
C * TIMES ZLB(COUNT,1) WAS REPEATED DURING THE *
C * SIMULATION. *
C * ZMEAN - REAL - THE MEAN OF THE JOINT PROBABILITY *
C * DISTRIBUTION FUNCTION. *

```

```

C *          ZSTANDARD_DEV - REAL - THE STANDARD DEVIATION OF THE      *
C *          JOINT PROBABILITY DISTRIBUTION Z.                        *
C *          ZUB(COUNT,1) - REAL - THE Z VALUE OF THE UPPER BOUND     *
C *          OF THE JOINT PROBABILITY DISTRIBUTION FUNCTION.          *
C *          ZUB(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS    *
C *          THAN OR EQUAL TO ZUB(COUNT,1).                            *
C *          ZUB(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF   *
C *          TIMES ZUB(COUNT,1) WAS REPEATED DURING THE               *
C *          SIMULATION.                                              *
C *          ZVARIANCE - REAL - THE VARIANCE OF THE JOINT             *
C *          PROBABILITY DISTRIBUTION FUNCTION.                       *
C *                                                                    *
C *****
C
      INTEGER COUNT,BX,DISTR,NCELLS,NUM,N,UBCOUNT,LBCOUNT
      REAL A,B,Z,AREA,PROB,X,TOLERANCE,F,D,SX,W
      REAL VARIANCE,MEAN,ZMEAN,ZVARIANCE,UB,LB
      CHARACTER DISREP*10,NAME*20,NAME2*20,ANSB*1
      COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT
      COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)
      COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
      COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
      OPEN (UNIT=23,FILE='W2')
      WRITE (3,*)
      WRITE (3,300)
300    FORMAT (2X,'SIMULATION AND BOUNDING (Y/N) ?')
      READ (0,310) ANSB
310    FORMAT (A1)
      IF (ANSB.NE.'Y'.AND.ANSB.NE.'N') THEN
          GOTO 320
      END IF
      DO 10 NUM=1,2
          CALL DIST
10     CONTINUE
      IF (ANSB.EQ.'Y') THEN
          CALL BOOKBOUND
          CALL OUTPUTBOUND
      ELSE
          CALL BOOK
          CALL OUTPUT
      END IF
      CLOSE 23
      STOP
      END

```

```

SUBROUTINE BOOK
C *****
C *
C * PROGRAM NAME: SUBROUTINE BOOK
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 4, 1987
C *
C * PURPOSE: SUBROUTINE BOOK SIMULATES THE JOINT DISTRIBUTION FUNCTION
C *           OF TWO RANDOM VARIABLES FOR THE DISCRETE APPROXIMATION.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: SUBROUTINES DISTRIBUTION
C *                                         FUNCTION
C *                                         SORT
C *
C * USAGE: CALL BOOK
C *
C * DISCUSSION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *     B(NCELLS,NUM) - REAL - THE DISCRETE POINT (CENTER OF
C *                       GRAVITY) REPRESENTATIONS FOR NUM'S PROBABILITY
C *                       PROBABILITY DENSITY FUNCTION.
C *     NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *                       LOCATIONS USED TO APPROXIMATE THE PROBABILITY
C *                       DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *     NUM - INTEGER - VARIABLE NUM IS THE NUMBER OF THE
C *             RANDOM VARIABLE.
C *   OUTPUT ARGUMENTS:
C *     COUNT - INTEGER - COUNT IS THE NUMBER OF UNIQUE Z
C *               LOCATIONS OF THE JOINT PROBABILITY DISTRIBUTION
C *               FUNCTION.
C *     TIME - INTEGER*3 - VARIABLE TIME IS THE TOTAL CPU TIME
C *               OF SIMULATION.
C *     Z(COUNT,1) - REAL - THE Z VALUE OF THE JOINT
C *               PROBABILITY DISTRIBUTION FUNCTION.
C *     Z(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS
C *               THAN OR EQUAL TO Z(COUNT,1).
C *     Z(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF
C *               TIMES Z(COUNT,1) WAS REPEATED DURING THE
C *               SIMULATION.
C *     ZMEAN - REAL - THE MEAN OF THE JOINT PROBABILITY
C *               DISTRIBUTION FUNCTION.
C *     ZVARIANCE - REAL - THE VARIANCE OF THE JOINT
C *               PROBABILITY DISTRIBUTION FUNCTION.
C *
C *****
C

```

```

COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT
COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)
COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
REAL B,Z,SX,F,D,WORK,A,VARIANCE,MEAN,ZMEAN,ZVARIANCE,TOTAL
REAL ZZ,UB,LB
INTEGER COUNT,FLAG,AX,X,Y,NCELLS,NUM,N,DISTR,UBCOUNT,LBCOUNT
INTEGER*3 TIME
CHARACTER DISREP*10,ANSB*1
COUNT=0
AX=0
TOTAL=0.0
CALL STIME
Z(0,1)=-1.0E30

```

```

C -----
C
C SIMULATE ALL POSSIBLE COMBINATIONS OF THE FUNCTION Z
C -----
C

```

```

DO 10 X=1,NCELLS(1)
DO 20 Y=1,NCELLS(2)
AX=AX+1
CALL FUNCTION (B(X,1),B(Y,2),ZZ)
CALL SORT (ZZ,Z,COUNT,AX)
TOTAL=TOTAL+ZZ
20 CONTINUE
10 CONTINUE

```

```

C -----
C
C END OF THE SIMULATION
C -----
C

```

```

CALL WTIME
CALL WTIME2 (TIME)
WRITE (3,*) TIME
ZMEAN=TOTAL/FLOAT(NCELLS(1)*NCELLS(2))

```

```

C -----
C
C CALCULATE THE JOINT PROBABILITY DISTRIBUTION FUNCTION
C -----
C

```

```

CALL DISTRIBUTION (Z,COUNT,ZMEAN,ZVARIANCE,NCELLS(1),

```

\*

RETURN  
END

NCELLS(2))

```

SUBROUTINE BOOKBOUND
C *****
C *
C * PROGRAM NAME: SUBROUTINE BOOKBOUND
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 4, 1987
C *
C * PURPOSE: SUBROUTINE BOOK SIMULATES THE JOINT DISTRIBUTION FUNCTION
C *           OF THE TWO RANDOM VARIABLES.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: SUBROUTINES DISTRIBUTION
C *                                         FUNCTION
C *                                         SORT
C *
C * USAGE: CALL BOOK
C *
C * DISCUSSION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *     BOUND(3,2) - REAL - ARRAY BOUND IS A TEMPORARY VARIABLE
C *                   THAT COMPUTES THE MINIMUM AND MAXIMUM
C *                   COMBINATIONS OF THE TWO CELLS REPRESENTING THE
C *                   TWO RANDOM VARIABLES.
C *     B(NCELLS,NUM) - REAL - THE DISCRETE POINT (CENTER OF
C *                   GRAVITY) REPRESENTATIONS FOR NUM'S PROBABILITY
C *                   PROBABILITY DENSITY FUNCTION.
C *     LB(NCELLS( ),2) - REAL - CONTAINS THE LOWER BOUND
C *                   LOCATION FOR CELL NCELLS( ).
C *     NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *                   LOCATIONS USED TO APPROXIMATE THE PROBABILITY
C *                   DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *     NUM - INTEGER - VARIABLE NUM IS THE NUMBER OF THE
C *                   RANDOM VARIABLE.
C *     UB(NCELLS( ),2) - REAL - CONTAINS THE UPPER BOUND
C *                   LOCATION FOR CELL NCELLS( ).
C *   OUTPUT ARGUMENTS:
C *     COUNT - INTEGER - COUNT IS THE NUMBER OF UNIQUE Z
C *                   LOCATIONS OF THE JOINT PROBABILITY DISTRIBUTION
C *                   FUNCTION.
C *     LBCOUNT - INTEGER - LBCOUNT IS THE NUMBER OF UNIQUE
C *                   LB( , ) LOCATIONS OF THE LOWER BOUND
C *                   DISTRIBUTION FUNCTION.
C *     UBCOUNT - INTEGER - UBCOUNT IS THE NUMBER OF UNIQUE
C *                   UB( , ) LOCATIONS OF THE UPPER BOUND
C *                   DISTRIBUTION FUNCTION.
C *     TIME - INTEGER*3 - VARIABLE TIME IS THE TOTAL CPU TIME
C *                   OF SIMULATION.

```

```

C *      Z(COUNT,1) - REAL - THE Z VALUE OF THE JOINT      *
C *      PROBABILITY DISTRIBUTION FUNCTION.                *
C *      Z(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS *
C *      THAN OR EQUAL TO Z(COUNT,1).                      *
C *      Z(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF *
C *      TIMES Z(COUNT,1) WAS REPEATED DURING THE          *
C *      SIMULATION.                                        *
C *      ZLB(COUNT,1) - REAL - THE Z VALUE OF THE LOWER BOUND *
C *      OF THE JOINT PROBABILITY DISTRIBUTION FUNCTION.   *
C *      ZLB(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS *
C *      THAN OR EQUAL TO ZLB(COUNT,1).                    *
C *      ZLB(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF *
C *      TIMES ZLB(COUNT,1) WAS REPEATED DURING THE        *
C *      SIMULATION.                                        *
C *      ZMEAN - REAL - THE MEAN OF THE JOINT PROBABILTITY *
C *      DISTRIBUTION FUNCTION.                             *
C *      ZUB(COUNT,1) - REAL - THE Z VALUE OF THE UPPER BOUND *
C *      OF THE JOINT PROBABILITY DISTRIBUTION FUNCTION.   *
C *      ZUB(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS *
C *      THAN OR EQUAL TO ZUB(COUNT,1).                    *
C *      ZUB(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF *
C *      TIMES ZUB(COUNT,1) WAS REPEATED DURING THE        *
C *      SIMULATION.                                        *
C *      ZVARIANCE - REAL - THE VARIANCE OF THE JOINT      *
C *      PROBABILITY DISTRIBUTION FUNCTION.                 *
C *                                                         *
C *****
C

```

```

COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT
COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)
COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
DIMENSION BOUND(3,3)
REAL B,Z,SX,F,D,WORK,A,VARIANCE,MEAN,ZMEAN,ZVARIANCE,TOTAL
REAL ZZ,UB,LB,UBTOTAL,LBTOTAL,UBMEAN,LBMEAN,UBVARIANCE
REAL LBVARIANCE
INTEGER COUNT,FLAG,AX,X,Y,NCELLS,NUM,N,DISTR,UBCOUNT,LBCOUNT
INTEGER LBX,UBX
INTECER*3 TIME
CHARACTER DISREP*10,ANSB*1
COUNT=0
UBCOUNT=0
LBCOUNT=0
AX=0
TOTAL=0.0
UBTOTAL=0.0
LBTOTAL=0.0
WRITE (3,123)

```



```

123      FORMAT (2X,'START SIMULATION')
        CALL STIME
        Z(0,1)=-1.0E30

C
C -----
C
C      SIMULATE ALL POSSIBLE COMBINATIONS OF THE FUNCTION Z
C -----
C
      DO 10 X=1,NCELLS(1)
        WRITE (3,*) X
        DO 20 Y=1,NCELLS(2)
          AX=AX+1
          LBX=LBX+1
          UBX=UBX+1
          CALL FUNCTION (B(X,1),B(Y,2),ZZ)
          CALL SORT (ZZ,Z,COUNT,AX)
          ZMIN=1.0E38
          ZMAX=-1.0E38
          BOUND(1,1)=UB(X,1)
          BOUND(2,1)=LB(X,1)
          BOUND(3,1)=B(X,1)
          BOUND(1,2)=UB(Y,2)
          BOUND(2,2)=LB(Y,2)
          BOUND(3,2)=B(Y,2)
          DO 8000 K=1,3
            DO 8010 KK=1,3
              CALL FUNCTION (BOUND(K,1),BOUND(KK,2),ZT)
              IF (ZT.LT.ZMIN) THEN
                ZMIN=ZT
              END IF
              IF (ZT.GT.ZMAX) THEN
                ZMAX=ZT
              END IF
            CONTINUE
          CONTINUE
          CALL SORT(ZMIN,ZUB,UBCOUNT,UBX)
          CALL SORT(ZMAX,ZLB,LBCOUNT,LBX)
          TOTAL=TOTAL+ZZ
          UBTOTAL=UBTOTAL+ZMIN
          LBTOTAL=LBTOTAL+ZMAX
        CONTINUE
      CONTINUE
20
10
C
C -----
C
C      END OF THE SIMULATION

```

```

C
C -----
C
      CALL WTIME
      CALL WTIME2 (TIME)
      WRITE (3,*) TIME
      WRITE (3,124)
124  FORMAT (2X,'END OF SIMULATION')
      ZMEAN=TOTAL/FLOAT(NCELLS(1)*NCELLS(2))
      UBMEAN=UBTOTAL/FLOAT(NCELLS(1)*NCELLS(2))
      LBMEAN=LBTOTAL/FLOAT(NCELLS(1)*NCELLS(2))
C
C -----
C
      CALCULATE THE JOINT PROBABILITY DISTRIBUTION FUNCTIONS
      FOR THE DISCRETE APPROXIMATION AND THE TWO BOUNDING CURVES.
C
C -----
C
      CALL DISTRIBUTION (Z,COUNT,ZMEAN,ZVARIANCE,NCELLS(1),
*                      NCELLS(2))
      CALL DISTRIBUTION (ZLB,LBCOUNT,LBMEAN,LBVARIANCE,NCELLS(1),
*                      NCELLS(2))
      CALL DISTRIBUTION (ZUB,UBCOUNT,UBMEAN,UBVARIANCE,NCELLS(1),
*                      NCELLS(2))
      RETURN
      END

```

```

SUBROUTINE CELLR
C *****
C *
C * PROGRAM NAME: SUBROUTINE CELLR
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 3, 1987
C *
C * PURPOSE: SUBROUTINE CELLR COMPUTES THE CENTER OF GRAVITY LOCATIONS
C *           FOR A GIVEN CELL WITH A KNOWN PROBABILITY DISTRIBUTION.
C *
C * USAGE: CALL CELLR
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: NONE
C *
C * DESCRIPTION OF PRAMETERS:
C *   INPUT ARGUMENTS:
C *       A(N) - REAL - WHERE N=NCELLS*2, ODD VALUES OF N
C *               REPRESENT THE CENTER OF GRAVITY LOCATIONS, AND
C *               THE EVEN VALUES OF N INDICATE THE CELL
C *               BOUNDARIES.
C *       ANSB - CHARACTER*1 - IF ANSB IS EQUAL TO 'Y' THEN THE
C *               PROGRAM COMPUTES THE UPPER AND LOWER BOUNDS OF
C *               THE JOINT DISTRIBUTION FUNCTION AS WELL AS THE
C *               DISCRETE APPROXIMATION FUNCTION. IF ANSB IS
C *               EQUAL TO 'N' THEN THE PROGRAM COMPUTES ONLY THE
C *               DISCRETE APPROXIMATION OF THE JOINT
C *               DISTRIBUTION FUNCTION.
C *       NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *               LOCATIONS THAT APPROXIMATE THE PROBABILITY
C *               DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *       NUM - INTEGER - VARIABLE NUM IS THE NUMBER OF THE
C *               RANDOM VARIABLE.
C *
C *   OUTPUT ARGUMENTS:
C *       B(NCELLS,NUM) - REAL - THE DISCRETE POINT (CENTER OF
C *               GRAVITY) REPRESENTATIONS FOR NUM'S PROBABILITY
C *               DENSITY FUNCTION.
C *       LB(NCELLS( ),2) - REAL - CONTAINS THE LOWER BOUND
C *               LOCATION FOR CELL NCELLS( ).
C *       UB(NCELLS( ),2) - REAL - CONTAINS THE UPPER BOUND
C *               LOCATION FOR CELL NCELLS( ).
C *   OUTPUT ARGUMENTS:
C *       COUNT - INTEGER - COUNT IS THE NUMBER OF UNIQUE Z
C *               LOCATIONS OF THE JOINT PROBABILITY DISTRIBUTION
C *               FUNCTION.
C *

```

```

C *****
C
COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT
COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)
COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
INTEGER BX,NCELLS,NUM,COUNT,DISTR,UBCOUNT,LBCOUNT
REAL A,B,Z,VARIANCE,MEAN,ZMEAN,ZVARIANCE,UB,LB
CHARACTER DISREP*10,ANSB*1
WRITE (3,2)
2  FORMAT (2X,'ENTERED CELLR')
   BX=0
   DO 80 K=1,NCELLS(NUM)*2,2
     BX=BX+1
     B(BX,NUM)=A(K)
80  CONTINUE
C
C -----FOR BOUNDING-----
C
   IF (ANSB.EQ.'Y') THEN
     BX=1
     DO 10 I=0,NCELLS(NUM)*2,2
       LB(BX,NUM)=A(I)
       UB(BX,NUM)=A(I+2)
       BX=BX+1
10  CONTINUE
   END IF
C
C -----END FOR BOUNDING-----
C
RETURN
END

```

```

SUBROUTINE DIST
C *****
C *
C * PROGRAM NAME: SUBROUTINE DIST
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 3, 1987
C *
C * PURPOSE: SUBROUTINE DIST INITIALIZES THE TYPE OF PROBABILITY
C *           DISTRIBUTION FOR EACH RANDOM VARIABLE.
C *
C * USAGE: CALL DIST
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: SUBROUTINES CELLR
C *                                     NORMAL
C *                                     NORMINV
C *                                     TRIANGEL
C *                                     UNIFORM
C *                                     SPECIAL
C *
C * DESCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *     DISTR(NUM) - INTEGER - THE NUMBER OF THE DISTRIBUTION
C *                   FOR RANDOM VARIABLE NUM (1=NORMAL, 2=UNIFORM,
C *                   3=TRIANGULAR, 4=SPECIAL).
C *     NUM - INTEGER - THE NUMBER OF THE RANDOM VARIABLE.
C *
C *   OUTPUT ARGUMENTS:
C *     A(N) - REAL - WHERE N=NCELLS*2, ODD VALUES OF N
C *                   REPRESENT THE CENTER OF GRAVITY LOCATIONS AND
C *                   EVEN VALUES OF N INDICATE THE CELL BOUNDARIES.
C *     DISREP - CHARACTER*10 - DISREP CONTAINS THE TYPE OF
C *                   PROBABILITY DISTRIBUTION FOR VARIABLE NUM
C *                   (NORMAL, UNIFORM, TRIANGULAR, SPECIAL).
C *     NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *                   LOCATIONS THAT APPROXIMATE THE PROBABILITY
C *                   DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *****
C
COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT
COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)
COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
INTEGER DISTR,NUM,NCELLS,COUNT,LBCOUNT,UBCOUNT
REAL VARIANCE,MEAN,A,B,Z,ZMEAN,ZVARIANCE,UB,LB
CHARACTER DISREP*10,ANSB*1

```

```

5      WRITE (3,*)
      WRITE (3,10)
10     FORMAT (2X,'DISTRIBUTIONS')
      WRITE (3,20)
20     FORMAT (2X,'1)  NORMAL')
      WRITE (3,30)
30     FORMAT (2X,'2)  UNIFORM')
      WRITE (3,40)
40     FORMAT (2X,'3)  TRIANGULAR')
      WRITE (3,50)
50     FORMAT (2X,'4)  SPECIAL')
      WRITE (3,80) NUM
80     FORMAT (2X,'ENTER NUMBER OF DISTRIBUTION FOR VARIABLE',I3)
      READ (0,*) DISTR(NUM)
      IF (DISTR(NUM).GT.4) GOTO 5
      IF (DISTR(NUM).EQ.1) THEN
          DISREP(NUM)='NORMAL'
          CALL NORMAL
      ELSE
          IF (DISTR(NUM).EQ.2) THEN
              DISREP(NUM)='UNIFORM'
              CALL UNIFORM
          ELSE
              IF (DISTR(NUM).EQ.3) THEN
                  DISREP(NUM)='TRIANGULAR'
                  CALL TRIANGLE
              ELSE
                  IF (DISTR(NUM).EQ.4) THEN
                      DISREP(NUM)='SPECIAL'
                      CALL SPECIAL
                  ELSE
                      GOTO 5
                  END IF
              END IF
          END IF
      END IF
      RETURN
END

```

```

      SUBROUTINE DISTRIBUTION (Z,COUNT,MEAN,VARIANCE,NC1,NC2)
C *****
C *
C * PROGRAM NAME: SUBROUTINE DISTRIBUTION
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 23, 1987
C *
C * PURPOSE: SUBROUTINE DISTRIBUTION COMPUTES THE JOINT DISTRIBUTION
C *           FUNCTION OF RANDOM VARIABLE Z.
C *
C * SUBROUTINES AND SUBPROCRAMS REQUIRED: NONE
C *
C * USACE: CALL DISTRIBUTION(Z,COUNT,MEAN,VARIANCE,NC1,NC2)
C *
C * DESCRIPTION OF PARAMETERS:
C *   INPUT ARCUMENTS:
C *     COUNT - INTECER - COUNT IS THE NUMBER OF UNIQUE Z
C *              LOCATIONS OF THE JOINT PROBABILITY DISTRIBUTION
C *              FUNCTION.
C *     MEAN - REAL - THE MEAN OF THE JOINT PROBABILITY
C *              DISTRIBUTION FUNCTION.
C *     NC1 - INTECER - THE NUMBER OF DISCRETE LOCATIONS THAT
C *              APPROXIMATE THE PROBABILTIIY DENSITY FUNCTION
C *              OF RANDOM VARIABLE 1.
C *     NC2 - INTECER - THE NUMBER OF DISCRETE LOCATIONS THAT
C *              APPROXIMATE THE PROBABILTIIY DENSITY FUNCTION
C *              OF RANDOM VARIABLE 2.
C *     Z(COUNT,1) - REAL - THE Z VALUE OF THE JOINT
C *              PROBABILITY DISTRIBUTION FUNCTION.
C *     Z(COUNT,2) - REAL - THE NUMBER OF TIMES THE VALUE OF
C *              TIMES Z(COUNT,1) WAS REPEATED DURING THE
C *              SIMULATION.
C *
C *   OUTPUT ARCUMENTS:
C *     VARIANCE - REAL - THE VARIANCE OF THE JOINT PROBABILITY
C *              DISTRIBUTION FUNCTION.
C *     Z(COUNT,3) - REAL - THE PROBABILITY OF Z BEINC LESS
C *              THAN OR EQUAL TO Z(COUNT,1).
C *              TIMES Z(COUNT,1) WAS REPEATED DURING THE
C *              SIMULATION.
C *****
C
      DIMENSION Z(0:40000,3)
      REAL Z,ZT,MEAN,VARIANCE
      INTECER COUNT,NC1,NC2

```

```

      ZT=0.0
      VARIANCE=0.0
      DO 80 I=1,COUNT
        Z(I,3)=Z(I,2)
        ZT=ZT+Z(I,3)
        Z(I,2)=ZT/FLOAT(NC1*NC2)
        VARIANCE=((Z(I,1)-MEAN)**2.0)*(Z(I,3)/FLOAT(NC1*NC2)))+
*          VARIANCE
80    CONTINUE
      RETURN
      END

```



```

SUBROUTINE FUNCTION (X1,X2,ANSZ)
C *****
C *
C * PROGRAM NAME: SUBROUTINE FUNCTION
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 23, 1987
C *
C * PURPOSE: SUBROUTINE FUNCTION COMPUTES THE VALUE OF ANSZ GIVEN
C *           THE VARIABLES X1 AND X2.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: NONE
C *
C * USAGE: CALL FUNCTION (X1,X2,ANSZ)
C *
C * DISCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *       X1 - REAL - THE VALUE OF THE FIRST VARIABLE.
C *       X2 - REAL - THE VALUE OF THE SECOND VARIABLE.
C *   OUTPUT ARGUMENTS:
C *       ANSZ - REAL - THE VALUE OF THE FUNCTION GIVEN THE
C *                VALUES OF THE TWO RANDOM VARIABLES X1 AND X2.
C *****
C
REAL X1,X2,ANSZ
PI=3.141592654

C -----
C
C ENTER FUNCTION TO BE SIMULATED
C       X1 = RANDOM VARIABLE 1
C       X2 = RANDOM VARIABLE 2
C       ANSZ = JOINT RANDOM VARIABLE
C -----
C
C ANSZ=(X1/2.0)+X2
C ANSZ=SQRT(X1**2+X2**2)
C ANSZ=X1*X2*12.0
C ANSZ=2.5-(X1*X2)
C ANSZ=(X1/2.0+X2/2.0)
C ANSZ=(X1*X2**2.0)/2.0
C ANSZ=(X1*(X2**3.0))/6.0
C ANSZ=X1*(X2**2.0/8.0)
C ANSZ=X1*(1/X2)
C ANSZ=X1-((4*50000)/(3.141592654*X2**2))

```

```

C      ANSZ=X1/(1+X2)
C      ANSZ=LOG10((1+X1)/X2)
C      ANSZ=X1-((4.0*50000)/(PI*X2))
C      ANSZ=SQRT(X1**2.+X2**2.)
C
C -----
C
      RETURN
      END

```

```

      SUBROUTINE NORMAL
C *****
C *
C * PROGRAM NAME: SUBROUTINE NORMAL
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 3, 1987
C *
C * PURPOSE: SUBROUTINE NORMAL FINDS THE NCELL LOCATION ON THE X AXIS
C *           GIVEN A PROBABILITY (AREA) FOR A NORMAL RANDOM VARIABLE.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: SUBROUTINES CELLR
C *                                         NORMINV
C *
C * USAGE: CALL NORMAL
C *
C * DISCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *     MEAN(NUM) - REAL - THE MEAN VALUE FOR RANDOM VARIABLE
C *               NUM.
C *     NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *               LOCATIONS THAT APPROXIMATE THE PROBABILITY
C *               DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *     NUM - INTEGER - THE NUMBER OF THE RANDOM VARIABLE.
C *     VARIANCE - REAL - THE VARIANCE VALUE FOR RANDOM
C *               VARIABLE NUM.
C *
C *   OUTPUT ARGUMENTS:
C *     A(N) - REAL - WHERE N=NCELLS*2, ODD VALUES OF N
C *               REPRESENT THE CENTER OF GRAVITY LOCATIONS AND
C *               THE EVEN VALUES OF N INDICATE THE CELL
C *               BOUNDARIES.
C *     B(NCELLS,NUM) - REAL - THE DISCRETE POINT (CENTER OF
C *               GRAVITY) REPRESENTATIONS FOR NUM'S PROBABILITY
C *               DENSITY FUNCTION.
C *
C *****
C
      COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT
      COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)
      COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
      COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
      REAL AREA,A,X,PROB,KAREA,MEAN,VARIANCE,B,Z
      REAL ZMEAN,ZVARIANCE,UB,LB
      INTEGER NCELLS,K,NUM,COUNT,DISTR,UBCOUNT,LBCOUNT
      CHARACTER DISREP*10,ANSB*1
      WRITE (3,*)

```

```

WRITE (3,*)
WRITE (3,20) NUM
20  FORMAT (2X,"NORMAL DISTRIBUTION FOR VARIABLE",I3)
WRITE (3,30)
30  FORMAT (2X,"ENTER MEAN")
    READ (0,*) MEAN(NUM)
    WRITE (3,40)
40  FORMAT (2X,"ENTER STANDARD DEVIATION")
    READ (0,*) VARIANCE(NUM)
    VARIANCE(NUM)=VARIANCE(NUM)**2.0
    WRITE (3,3) NUM
3   FORMAT (2X,"ENTER NUMBER OF CELLS FOR VARIABLE",I3)
    READ (0,*) NCELLS(NUM)
    AREA=1/(FLOAT(NCELLS(NUM))*2.0)
    DO 10 K=1,NCELLS(NUM)*2.0
        KAREA=AREA*(FLOAT(K))
        P=KAREA
        CALL NORMINV(P,X,D,IE)
        A(K)=X*SQRT(VARIANCE(NUM))+MEAN(NUM)
10  CONTINUE
    A(0)=-A(NCELLS(NUM)*2)
    CALL CELLR
    RETURN
    END

```

```

SUBROUTINE NORMINV (P,X,D,IE)
C *****
C *
C * PROGRAM NAME: NORMINV
C *
C * PURPOSE: SUBROUTINE NORMINV WAS SUPPLIED FROM THE SCIENTIFIC
C *           SUBROUTINE PACKAGE (SSP). THE SSP CONTAINS A COLLECTION
C *           OF MATHEMATICAL AND STATISTICAL ROUTINES FOR THE
C *           HARRIS COMPUTER SYSTEM.
C *
C *****
C .....
C
C SUBROUTINE NORMINV
C
C PURPOSE
C COMPUTES  $X = P^{**(-1)}(Y)$ , THE ARGUMENT X SUCH THAT  $Y = P(X) =$ 
C THE PROBABILITY THAT THE RANDOM VARIABLE U, DISTRIBUTED
C NORMALLY(0,1), IS LESS THAN OR EQUAL TO X. F(X), THE
C ORDINATE OF THE NORMAL DENSITY, AT X, IS ALSO COMPUTED.
C
C USAGE
C CALL NORMINV (P,X,D,IER)
C
C DESCRIPTION OF PARAMETERS
C P - INPUT PROBABILITY.
C X - OUTPUT ARGUMENT SUCH THAT  $P = Y =$  THE PROBABILITY THAT
C U, THE RANDOM VARIABLE, IS LESS THAN OR EQUAL TO X.
C D - OUTPUT DENSITY, F(X).
C IER - OUTPUT ERROR CODE
C = -1 IF P IS NOT IN THE INTERVAL (0,1), INCLUSIVE.
C X=D=.99999E+38 IN THIS CASE
C = 0 IF THERE IS NO ERROR. SEE REMARKS, BELOW.
C
C REMARKS
C MAXIMUM ERROR IS 0.00045.
C IF P = 0, X IS SET TO  $-(10)**38$ . D IS SET TO 0.
C IF P = 1, X IS SET TO  $(10)**38$ . D IS SET TO 0.
C
C SUBROUTINES AND SUBPROGRAMS REQUIRED
C NONE
C
C METHOD
C BASED ON APPROXIMATIONS IN C. HASTINGS, APPROXIMATIONS FOR
C DIGITAL COMPUTERS, PRINCETON UNIV. PRESS, PRINCETON, N.J.,
C 1955. SEE EQUATION 26.2.23, HANDBOOK OF MATHEMATICAL
C FUNCTIONS, ABRAMOWITZ AND STEGUN, DOVER PUBLICATIONS, INC.,

```

```

C          NEW YORK.
C
C.....
C
C      REAL X,D,P,T2,T
      INTEGER IE
      IE=0
      X=1.7E+38
      D=X
      IF(P)1,4,2
1  IE=-1
      GO TO 12
2  IF (P-1.0)7,5,1
4  X=-1.7E+38
5  D=0.0
      GO TO 12
C
C
7  D=P
      IF(D-0.5)9,9,8
8  D=1.0-D
9  T2=ALOG(1.0/(D*D))
      T=SQRT(T2)

      X=T-(2.515517+0.802853*T+0.010328*T2)/(1.0+1.432788*T+0.189269*T2
1  +0.001308*T*T2)
      IF(P-0.5)10,10,11
10 X=-X
11 D=0.3989423*EXP(-X*X/2.0)
12 RETURN
      END

```

# SUBROUTINE OUTPUT

```

C *****
C *
C * PROGRAM NAME: SUBROUTINE OUTPUT
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 3, 1987
C *
C * PURPOSE: SUBROUTINE OUTPUT PRINTS THE CENTER OF GRAVITY LOCATIONS
C *          FOR RANDOM VARIABLES A AND B, AND THE JOINT PROBABILITY
C *          DISTRIBUTION OF VARIABLE Z TO FILE W2. OUTPUT ALSO PRINTS
C *          THE JOINT DISTRIBUTION FUNCTION TO A GIVEN FILE SO THE
C *          JOINT DISTRIBUTION CAN BE INPUT AGAIN AS A 'SPECIAL'
C *          DISTRIBUTION FUNCTION.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: NONE
C *
C * USAGE: CALL OUTPUT
C *
C * DESCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *       B(NCELLS,NUM) - REAL - THE DISCRETE POINT (CENTER OF
C *          GRAVITY) REPRESENTATIONS FOR NUM'S PROBABILITY
C *          DENSITY FUNCTION.
C *       COUNT - INTEGER - COUNT IS THE NUMBER OF UNIQUE Z
C *          LOCATIONS OF THE JOINT PROBABILITY DISTRIBUTION
C *          FUNCTION.
C *       DISREP(NUM) - CHARACTER*10 - DISREP CONTAINS THE
C *          TYPE OF PROBABILITY DISTRIBUTION FOR VARIABLE
C *          NUM (NORMAL, UNIFORM, TRIANGULAR, SPECIAL).
C *       MEAN(NUM) - REAL - THE MEAN VALUE FOR RANDOM VARIABLE
C *          NUM.
C *       NAME2 - CHARACTER*30 - NAME2 IS THE OUTPUT DATA FILE
C *          NAME OF THE JOINT DISTRIBUTION FUNCTION.
C *       NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *          LOCATIONS THAT APPROXIMATE THE PROBABILITY
C *          DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *       NUM - INTEGER - RANGE OF 1 TO 2, VARIABLE 'NUM' IS THE
C *          NUMBER OF THE RANDOM VARIABLE.
C *       VARIANCE(NUM) - REAL - THE VARIANCE VALUE FOR RANDOM
C *          VARIABLE NUM.
C *       Z(COUNT,1) - REAL - THE Z VALUE OF THE JOINT
C *          PROBABILITY DISTRIBUTION FUNCTION.
C *       Z(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS
C *          THAN OR EQUAL TO Z(COUNT,1).
C *       Z(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF
C *          TIMES Z(COUNT,1) WAS REPEATED DURING THE

```

```

C *          SIMULATION.
C *          ZMEAN - REAL - THE MEAN OF THE JOINT PROBABILITY
C *          DISTRIBUTION FUNCTION.
C *          ZVARIANCE - REAL - THE VARIANCE OF THE JOINT
C *          PROBABILITY DISTRIBUTION FUNCTION.
C *
C *          OUTPUT ARGUMENTS:
C *          ZSTANDARD_DEV - REAL - THE STANDARD DEVIATION OF THE
C *          JOINT PROBABILITY DISTRIBUTION Z.
C *
C *****
C
COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT
COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)
COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
REAL Z,MAXDEV,B,MEAN,VARIANCE,A,ZMEAN,ZVARIANCE,UB,LB
INTEGER NCELLS,COUNT,DISTR,NUM,UBCOUNT,LBCOUNT
CHARACTER DISREP*10,NAME*20,NAME2*20,ANSB*1

C
C -----
C
C          PRINT DATA TO FILE "W2"
C
C -----
C
WRITE (23,*)
WRITE (23,*)
WRITE (23,70)
70  FORMAT (2X,"RANDOM VARIABLE: 1",18X,"RANDOM VARIABLE: 2")
WRITE (23,80) DISREP(1),DISREP(2)
80  FORMAT (2X,"DISTRIBUTION: ",A11,11X,"DISTRIBUTION: ",A)
WRITE (23,90) MEAN(1),MEAN(2)
90  FORMAT (2X,"MEAN: ",F12.6,18X,"MEAN: ",F12.6)
WRITE (23,100) VARIANCE(1),VARIANCE(2)
100 FORMAT (2X,"VARIANCE: ",F12.6,14X,"VARIANCE: ",F12.6)
STANDARD_DEV1=SQRT(VARIANCE(1))
STANDARD_DEV2=SQRT(VARIANCE(2))
WRITE (23,101) STANDARD_DEV1,STANDARD_DEV2
101 FORMAT (2X,"STANDARD DEVIATION: ",F12.6,4X,
*STANDARD DEVIATION: ",F12.6)
WRITE (23,170) NCELLS(1),NCELLS(2)
170 FORMAT (2X,"NUMBER OF CELLS: ",I4,15X,"NUMBER OF CELLS: ",I4)
WRITE (23,*)
WRITE (23,110)
110 FORMAT (6X,"X LOCATION",26X,"X LOCATION")
WRITE (23,120)
120 FORMAT (4X,"FOR VARIABLE 1",22X,"FOR VARIABLE 2")

```



```

113      WRITE (23,113)
      FORMAT (4X,'-----',22X,'-----')
      IF (NCELLS(1).GT.NCELLS(2)) THEN
        INCELLS=NCELLS(1)
      ELSE
        INCELLS=NCELLS(2)
      END IF
      DO 160 I=1,INCELLS
        IF (I.LE.NCELLS(1).AND.I.LE.NCELLS(2)) THEN
          WRITE (23,130) B(I,1),B(I,2)
130        FORMAT (4X,F12.6,23X,F12.6)
          ELSE
            IF (I.LE.NCELLS(1)) THEN
              WRITE (23,140) B(I,1)
140            FORMAT (4X,F12.6)
            ELSE
              WRITE (23,150) B(I,2)
150            FORMAT (37X,F12.6)
            END IF
          END IF
160      CONTINUE
      WRITE (23,*)
      WRITE (23,161)
161      FORMAT (19X,'***** Z FUNCTION *****')
      WRITE (23,11) COUNT
11      FORMAT (15X,'NUMBER OF UNIQUE Z LOCATIONS: ',I5)
      WRITE (23,34) ZMEAN
34      FORMAT(15X,'MEAN: ',F12.6)
      WRITE (23,162) ZVARIANCE
162      FORMAT (15X,'VARIANCE: ',F12.6)
      ZSTANDARD_DEV=SQRT(ZVARIANCE)
      WRITE (23,163) ZSTANDARD_DEV
163      FORMAT (15X,'STANDARD DEVIATION: ',F12.6)
      WRITE (23,*)
      WRITE (23,13)
13      FORMAT (15X,'CELL')
      WRITE (23,12)
12      FORMAT (16X,'NO.',3X,'Z LOCATION',7X,'Z DISTRIBUTION')
      WRITE (23,132)
132      FORMAT (15X,'---',3X,'-----',7X,'-----')
      DO 10 I=1,COUNT
        WRITE (23,20) I,Z(I,1),Z(I,2)
20        FORMAT (14X,I4,2X,F12.6,6X,F12.6)
10      CONTINUE
C
C -----
C
C
C
      PRINT DISTRIBUTION DATA TO NAME2 FILE

```

```

C
C -----
C
      WRITE (3,116)
116   FORMAT (2X,'ENTER DISTRIBUTION DATA FILE NAME')
      READ (0,117) NAME2
117   FORMAT (A20)
      OPEN (UNIT=25,FILE=NAME2)
      WRITE (25,538) NCELLS(1),NCELLS(2)
538   FORMAT (I5,I5)
      WRITE (25,189) COUNT
189   FORMAT (I5)
      DO 89 I=1,COUNT
          WRITE (25,112) Z(I,1),Z(I,2),Z(I,3)
112   FORMAT (F20.7,F12.7,F12.7)
89   CONTINUE
      CLOSE 25
30   RETURN
      END

```

```

SUBROUTINE OUTPUTBOUND
C *****
C *
C * PROGRAM NAME: SUBROUTINE OUTPUTBOUND
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 3, 1987
C *
C * PURPOSE: SUBROUTINE OUTPUT PRINTS THE CENTER OF GRAVITY LOCATIONS
C *           FOR RANDOM VARIABLES A AND B, AND THE JOINT PROBABILITY
C *           DISTRIBUTION OF VARIABLE Z TO FILE W2. OUTPUT ALSO PRINTS
C *           THE JOINT DISTRIBUTION FUNCTION TO A GIVEN FILE SO THE
C *           JOINT DISTRIBUTION CAN BE INPUT AGAIN AS A 'SPECIAL'
C *           DISTRIBUTION FUNCTION.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: NONE
C *
C * USAGE: CALL OUTPUT
C *
C * DESCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *       B(NCELLS,NUM) - REAL - THE DISCRETE POINT (CENTER OF
C *           GRAVITY) REPRESENTATIONS FOR NUM'S PROBABILITY
C *           DENSITY FUNCTION.
C *       COUNT - INTEGER - COUNT IS THE NUMBER OF UNIQUE Z
C *           LOCATIONS OF THE JOINT PROBABILITY DISTRIBUTION
C *           FUNCTION.
C *       DISREP(NUM) - CHARACTER*10 - DISREP CONTAINS THE
C *           TYPE OF PROBABILITY DISTRIBUTION FOR VARIABLE
C *           NUM (NORMAL, UNIFORM, TRIANGULAR, SPECIAL).
C *       MEAN(NUM) - REAL - THE MEAN VALUE FOR RANDOM VARIABLE
C *           NUM.
C *       NAME2 - CHARACTER*30 - NAME2 IS THE OUTPUT DATA FILE
C *           NAME OF THE JOINT DISTRIBUTION FUNCTION.
C *       NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *           LOCATIONS THAT APPROXIMATE THE PROBABILITY
C *           DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *       NUM - INTEGER - RANGE OF 1 TO 2, VARIABLE 'NUM' IS THE
C *           NUMBER OF THE RANDOM VARIABLE.
C *       VARIANCE(NUM) - REAL - THE VARIANCE VALUE FOR RANDOM
C *           VARIABLE NUM.
C *       Z(COUNT,1) - REAL - THE Z VALUE OF THE JOINT
C *           PROBABILITY DISTRIBUTION FUNCTION.
C *       Z(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS
C *           THAN OR EQUAL TO Z(COUNT,1).
C *       Z(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF
C *           TIMES Z(COUNT,1) WAS REPEATED DURING THE

```

```

C *           SIMULATION. *
C *           ZLB(COUNT,1) - REAL - THE Z VALUE OF THE LOWER BOUND *
C *           OF THE JOINT PROBABILITY DISTRIBUTION FUNCTION. *
C *           ZLB(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS *
C *           THAN OR EQUAL TO ZLB(COUNT,1). *
C *           ZLB(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF *
C *           TIMES ZLB(COUNT,1) WAS REPEATED DURING THE *
C *           SIMULATION. *
C *           ZMEAN - REAL - THE MEAN OF THE JOINT PROBABILITY *
C *           DISTRIBUTION FUNCTION. *
C *           ZUB(COUNT,1) - REAL - THE Z VALUE OF THE UPPER BOUND *
C *           OF THE JOINT PROBABILITY DISTRIBUTION FUNCTION. *
C *           ZUB(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS *
C *           THAN OR EQUAL TO ZUB(COUNT,1). *
C *           ZUB(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF *
C *           TIMES ZUB(COUNT,1) WAS REPEATED DURING THE *
C *           SIMULATION. *
C *           ZVARIANCE - REAL - THE VARIANCE OF THE JOINT *
C *           PROBABILITY DISTRIBUTION FUNCTION. *
C *
C *           OUTPUT ARGUMENTS: *
C *           ZSTANDARD_DEV - REAL - THE STANDARD DEVIATION OF THE *
C *           JOINT PROBABILITY DISTRIBUTION Z. *
C *
C *****
C
COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT
COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)
COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
REAL Z,MAXDEV,B,MEAN,VARIANCE,A,ZMEAN,ZVARIANCE,UB,LB
INTEGER NCELLS,COUNT,DISTR,NUM,UBCOUNT,LBCOUNT
CHARACTER DISREP*10,NAME*20,NAME2*20,ANSB*1

C
C -----
C
C           PRINT DATA TO FILE "W2"
C
C -----
C
WRITE (23,70)
70  FORMAT (14X,"RANDOM VARIABLE: 1",34X,"RANDOM VARIABLE: 2")
WRITE (23,80) DISREP(1),DISREP(2)
80  FORMAT (14X,"DISTRIBUTION: ",A11,27X,"DISTRIBUTION: ",A)
WRITE (23,90) MEAN(1),MEAN(2)
90  FORMAT (14X,"MEAN: ",F12.6,34X,"MEAN: ",F12.6)
WRITE (23,100) VARIANCE(1),VARIANCE(2)
100 FORMAT (14X,"VARIANCE: ",F12.6,30X,"VARIANCE: ",F12.6)

```

```

        STANDARD_DEV1=SQRT(VARIANCE(1))
        STANDARD_DEV2=SQRT(VARIANCE(2))
        WRITE (23,101) STANDARD_DEV1,STANDARD_DEV2
101      FORMAT (14X,'STANDARD DEVIATION: ',F12.6,20X,
* 'STANDARD DEVIATION: ',F12.6)
        WRITE (23,170) NCELLS(1),NCELLS(2)
170      FORMAT (14X,'NUMBER OF CELLS :',I4,31X,'NUMBER OF CELLS: ',I4)
        WRITE (23,*)
        WRITE (23,110)
110      FORMAT (8X,'X1 MINIMUM',7X,'CENTER OF',6X,'X1 MAXIMUM',
* 10X,'X2 MINIMUM',7X,'CENTER OF',6X,'X2 MAXIMUM')
        WRITE (23,120)
120      FORMAT (11X,'VALUE',10X,'GRAVITY',10X,'VALUE',15X,'VALUE',
* 10X,'GRAVITY',10X,'VALUE')
        WRITE (23,123)
123      FORMAT (6X,'-----',4X,'-----',4X,'-----'
* ,8X,'-----',4X,'-----',4X,'-----')
        IF (NCELLS(1).GT.NCELLS(2)) THEN
            INCELLS=NCELLS(1)
        ELSE
            INCELLS=NCELLS(2)
        END IF
        DO 160 I=1,INCELLS
            IF (I.LE.NCELLS(1).AND.I.LE.NCELLS(2)) THEN
                WRITE (23,130) LB(I,1),B(I,1),UB(I,1),LB(I,2),B(I,2),
*          UB(I,2)
130      FORMAT (6X,F12.7,4X,F12.7,4X,F12.7,8X,F12.7,4X,F12.7,
*          4X,F12.7)
            ELSE
                IF (I.LE.NCELLS(1)) THEN
                    WRITE (23,140) LB(I,1),B(I,1),UB(I,1)
140      FORMAT (6X,F12.7,4X,F12.7,4X,F12.7)
                ELSE
                    WRITE (23,150) LB(I,2),B(I,2),UB(I,2)
150      FORMAT (39X,F12.7,4X,F12.7,4X,F12.7)
                END IF
            END IF
        CONTINUE
        WRITE (23,*)
        WRITE (23,*)
        WRITE (23,161)
161      FORMAT (38X,'***** Z FUNCTION *****')
        WRITE (23,11) COUNT
11      FORMAT (38X,'NUMBER OF UNIQUE Z LOCATIONS: ',I5)
        WRITE (23,34) ZMEAN
        WRITE (23,162) ZVARIANCE
162      FORMAT (38X,'VARIANCE: ',F12.5)
        ZSTANDARD_DEV=SQRT(ZVARIANCE)

```

```

WRITE (23,163) ZSTANDARD_DEV
163  FORMAT (38X,'STANDARD DEVIATION: ',F12.5)
34   FORMAT(38X,'MEAN: ',F12.5)
      WRITE (23,*)
      WRITE (23,13)
13   FORMAT (2X,'CELL',6X,'DISCRETE APPROXIMATION',15X,
*     'LOWER BOUNDS',20X,'UPPER BOUNDS')
      WRITE (23,12)
12   FORMAT (3X,'NO.',9X,'Z',9X'DISTRIBUTION',10X,'Z',9X,
*     'DISTRIBUTION',10X,'Z',9X,'DISTRIBUTION')
      WRITE (23,2000)
2000  FORMAT (2X,'-----',3X,'-----',4X,'-----',4X,
*     '-----',4X,'-----',4X,'-----',4X,
*     '-----')
      ICOUNT=-9999
      IF (UBCOUNT.GT.ICOUNT) THEN
        ICOUNT=UBCOUNT
      ELSE
        IF (LBCOUNT.GT.ICOUNT) THEN
          ICOUNT=LBCOUNT
        ELSE
          ICOUNT=COUNT
        END IF
      END IF
      DO 10 I=1,ICOUNT
        IF (I.LE.UBCOUNT.AND.I.LE.LBCOUNT.AND.I.LE.COUNT) THEN
          WRITE (23,20) I,Z(I,1),Z(I,2),ZLB(I,1),ZLB(I,2),
*             ZUB(I,1),ZUB(I,2)
20    FORMAT (1X,I4,3X,F12.7,4X,F12.7,4X,F12.7,4X,F12.7,
*             4X,F12.7,4X,F12.7)
        ELSE
          IF (I.LE.LBCOUNT.AND.I.LE.COUNT.AND.I.GT.UBCOUNT) THEN
            WRITE (23,720) I,Z(I,1),Z(I,2),ZLB(I,1),ZLB(I,2)
720    FORMAT (1X,I4,3X,F12.7,4X,F12.7,4X,F12.7,4X,F12.7)
          ELSE
            IF (I.LE.LBCOUNT.AND.I.GT.COUNT.AND.I.LE.UBCOUNT)
              THEN
                WRITE (23,730) I,ZLB(I,1),ZLB(I,2),ZUB(I,1),
*             ZUB(I,2)
730    FORMAT (1X,I4,35X,F12.7,4X,F12.7,4X,F12.7
*             ,4X,F12.7)
            ELSE
              IF (I.GT.LBCOUNT.AND.I.LE.COUNT.AND.I.LE.UBCOUNT)
                THEN
                  WRITE (23,740) I,Z(I,1),Z(I,2),
*             ZUB(I,1),ZUB(I,2)
740    FORMAT (1X,I4,3X,F12.7,4X,F12.7,36X,F12.7,4X,
*             F12.7)

```

```

ELSE
  IF (I.LE.LBCOUNT.AND.I.GT.COUNT.
    AND.I.GT.UBCOUNT) THEN
    *      WRITE (23,750) I,ZLB(I,1),ZLB(I,2)
750      FORMAT (1X,I4,35X,F12.7,4X,F12.7)
    ELSE
      IF (I.LE.COUNT.AND.I.GT.LBCOUNT.
        AND.I.GT.UBCOUNT) THEN
        *      WRITE (23,760) I,Z(I,1),Z(I,2)
760      FORMAT (1X,I4,3X,F12.7,4X,F12.7)
      ELSE
        WRITE (23,770) I,ZUB(I,1),ZUB(I,2)
770      FORMAT (1X,I4,67X,F12.7,4X,F12.7)
      END IF
    END IF
  END IF
END IF
END IF
END IF
CONTINUE
10
C
C -----
C
C      PRINT DISTRIBUTION DATA TO NAME2 FILE
C
C -----
C
  WRITE (3,116)
  116  FORMAT (2X,"ENTER DISTRIBUTION DATA FILE NAME")
  READ (0,117) NAME2
  117  FORMAT (A20)
  OPEN (UNIT=25,FILE=NAME2)
  WRITE (25,538) NCELLS(1),NCELLS(2)
  538  FORMAT (I5,I5)
  WRITE (25,189) UBCOUNT,LBCOUNT
  189  FORMAT (I5,I5)
  IF (UBCOUNT.GT.LBCOUNT) THEN
    ICOUNT=UBCOUNT
  ELSE
    ICOUNT=LBCOUNT
  END IF
  DO 89 I=1,ICOUNT
    IF (I.LE.LBCOUNT.AND.I.LE.UBCOUNT) THEN
      112  WRITE (25,112) ZUB(I,1),ZUB(I,2),ZLB(I,1),ZLB(I,2)
      FORMAT (F12.6,1X,F12.6,1X,F12.6,1X,F12.6)
    ELSE
      IF (I.LE.LBCOUNT) THEN
        WRITE (25,113) ZLB(I,1),ZLB(I,2)

```

```

113             FORMAT (12X,1X,12X,1X,F12.6,1X,F12.6)
                ELSE
                WRITE (25,114) ZUB(I,1),ZUB(I,2)
114             FORMAT (F12.6,1X,F12.6)
                END IF
            END IF
89    CONTINUE
        CLOSE 25
30    RETURN
    END

```





```

IXMIN=1
IXMAX=AX-1
40  JUMP=NINT((IXMAX-IXMIN)/2.0)+IXMIN
DIFFER=ABS(ZZ-Z(JUMP,1))
IF (ZZ.EQ.Z(JUMP,1).OR.DIFFER.LE.1.0E-7) THEN
    Z(JUMP,2)=Z(JUMP,2)+1
    AX=AX-1
    RETURN
END IF
IF (ZZ.LT.Z(JUMP,1).AND.ZZ.GT.Z(JUMP-1,1)) THEN
    DO 30 K=AX,JUMP+1,-1
        Z(K,1)=Z(K-1,1)
        Z(K,2)=Z(K-1,2)
30  CONTINUE
    Z(JUMP,1)=ZZ
    Z(JUMP,2)=1
    COUNT=COUNT+1
ELSE
    IF (ZZ.LT.Z(JUMP,1)) THEN
        IXMAX=JUMP-1
    ELSE
        IXMIN=JUMP
    END IF
    GOTO 40
END IF
RETURN
END

```

```

SUBROUTINE SPECIAL
C *****
C *
C * PROGRAM NAME: SUBROUTINE SPECIAL
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 4, 1987
C *
C * PURPOSE: SUBROUTINE SPECIAL READS A DISTRIBUTION FUNCTION STORED IN *
C *           A DATA FILE. SPECIAL THEN FINDS THE NCELL LOCATION ON THE *
C *           X AXIS GIVEN A PROBABILITY (AREA) FOR THE DISTRIBUTION *
C *           FUNCTION.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: SUBROUTINE CELLR
C *
C * USAGE: CALL SPECIAL
C *
C * DISCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *       NAME - CHARACTER*30 - NAME IS THE INPUT DATA FILE NAME
C *             OF THE SPECIAL DISTRIBUTION FUNCTION FOR
C *             RANDOM VARIABLE NUM.
C *       NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *             LOCATIONS THAT APPROXIMATE THE PROBABILITY
C *             DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *       NUM - INTEGER - VARIABLE NUM IS THE NUMBER OF THE
C *             RANDOM VARIABLE.
C *   OUTPUT ARGUMENTS:
C *       A(N) - REAL - WHERE N=NCELLS*2, ODD VALUES OF N
C *             REPRESENT THE CENTER OF GRAVITY LOCATIONS AND
C *             THE EVEN VALUES OF N INDICATE THE CELL
C *             BOUNDARIES.
C *       B(NCELLS,NUM) - REAL - THE DISCRETE POINT (CENTER OF
C *             GRAVITY) REPRESENTATIONS FOR NUM'S PROBABILITY
C *             DENSITY FUNCTION.
C *       MEAN(NUM) - REAL - MEAN(NUM) IS THE MEAN VALUE OF
C *             RANDOM VARIABLE NUM.
C *       VARIANCE(NUM) - REAL - VARIANCE(NUM) IS THE VARIANCE OF
C *             RANDOM VARIABLE NUM.
C *
C *****
C
REAL A,TOTAL,DS,AREA,KAREA,MEAN,VARIANCE,B,Z,UB,LB,P
INTEGER NUM,INC,NCELLS,COUNT,DISTR,NUMBER,UBCOUNT,LBCOUNT
INTEGER NC1,NC2
CHARACTER NAME*30,DISREP*10,ANSB*1
COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT

```

```

COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)
COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
COMMON/CSPECIAL/DS(40000,3)
COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
WRITE (3,*)
WRITE (3,10) NUM
10  FORMAT (2X,'SPECIAL DISTRIBUTION FOR VARIABLE',I3)
    WRITE (3,20)
20  FORMAT (2X,'ENTER DATA DISTRIBUTION FILE NAME')
    READ (0,30) NAME
30  FORMAT (A30)
    WRITE (3,110) NUM
110  FORMAT (2X,'ENTER NUMBER OF CELLS FOR VARIABLE',I3)
    READ (0,*) NCELLS(NUM)
    OPEN (UNIT=26,FILE=NAME)
    READ (26,120) NC1,NC2
120  FORMAT (I5,I5)
    READ (26,40) NUMBER
40  FORMAT (I5)
    TOTAL=0.0
    DO 50 I=1,NUMBER
        READ (26,60) DS(I,1),DS(I,2),DS(I,3)
60    FORMAT (F20.7,F12.7,F12.7)
        TOTAL=TOTAL+DS(I,1)
50    CONTINUE
    MEAN(NUM)=TOTAL/FLOAT(NUMBER)
    VARIANCE(NUM)=0.0
    DO 90 I=1,NUMBER
        P=DS(I,3)/FLOAT(NC1*NC2)
        VARIANCE(NUM)=(DS(I,1)-MEAN(NUM))**2)*P+VARIANCE(NUM)
90    CONTINUE
    CLOSE 26
    INC=0
    AREA=1/(FLOAT(NCELLS(NUM))*2.0)
    DO 70 K=1,NCELLS(NUM)*2
        KAREA=AREA*(FLOAT(K))
100    INC=INC+1
        IF (DS(INC,2).GE.KAREA) THEN
            A(K)=(KAREA-DS(INC-1,2))*(DS(INC,1)-DS(INC-1,1))/
*          (DS(INC,2)-DS(INC-1,2))+DS(INC-1,1)
            INC=INC-1
        ELSE
            GOTO 100
        END IF
70    CONTINUE
    CALL CELLR
    RETURN
END

```

```

SUBROUTINE TRIANGLE
C *****
C *
C * PROGRAM NAME: SUBROUTINE TRIANGLE
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 4, 1987
C *
C * PURPOSE: SUBROUTINE TRIANGLE FINDS THE NCELL LOCATION ON THE X AXIS
C *          GIVEN A PROBABILITY (AREA) FOR A TRIANGULAR RANDOM
C *          VARIABLE.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: SUBROUTINE CELLR
C *
C * USAGE: CALL TRIANGLE
C *
C * DISCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *       AA - REAL - VARIABLE AA IS THE MINIMUM VALUE OF THE
C *             TRIANGULAR DENSITY FUNCTION.
C *       BB - REAL - VARIABLE BB IS THE MAXIMUM VALUE OF THE
C *             TRIANGULAR DENSITY FUNCTION.
C *       CC - REAL - VARIABLE CC IS THE MEDIAN VALUE OF THE
C *             TRIANGULAR DENSITY FUNCTION.
C *       NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *             LOCATIONS THAT APPROXIMATE THE PROBABILITY
C *             THE PROBABILITY DENSITY FUNCTION OF RANDOM
C *             VARIABLE NUM.
C *       NUM - INTEGER - VARIABLE NUM IS THE NUMBER OF THE
C *             RANDOM VARIABLE.
C *   OUPUT ARGUMENTS:
C *       A(N) - REAL - WHERE N=NCELLS*2, ODD VALUES OF N
C *             REPRESENT THE CENTER OF GRAVITY LOCATIONS AND
C *             THE EVEN VALUES OF N INDICATE THE CELL
C *             BOUNDARIES.
C *       B(NCELLS,NUM) - REAL - THE DISCRETE POINT (CENTER OF
C *             GRAVITY) REPRESENTATIONS FOR NUM'S PROBABILITY
C *             DENSITY FUNCTION.
C *       MEAN(NUM) - REAL - MEAN(NUM) IS THE MEAN VALUE OF
C *             RANDOM VARIABLE NUM.
C *       VARIANCE(NUM) - REAL - VARIANCE(NUM) IS THE VARIANCE OF
C *             RANDOM VARIABLE NUM.
C *****
C
REAL AREA,KAREA,A,B,MEAN,VARIANCE,AA,BB,CC,TAREA,Z
REAL ZMEAN,ZVARIANCE,UB,LB

```

```

INTEGER NCELLS,NUM,K,COUNT,DISTR,UBCOUNT,LBCOUNT
CHARACTER DISREP*10,ANSB
COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT
COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)
COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
5  WRITE (3,*)
    WRITE (3,*)
    WRITE (3,10) NUM
10  FORMAT (2X,'TRIANGULAR DISTRIBUTION FOR VARIABLE',I3)
    WRITE (3,20)
20  FORMAT (2X,'PARAMETERS: A<C<B')
    WRITE (3,30)
30  FORMAT (2X,'RANGE: (A,B)')
    WRITE (3,40)
40  FORMAT (2X,'ENTER "A" (MINIMUM) VALUE')
    READ (0,*) AA
    WRITE (3,50)
50  FORMAT (2X,'ENTER "C" (MIDDLE) VALUE')
    READ (0,*) CC
    IF (CC-AA.LT.0) THEN
        GOTO 5
    END IF
    WRITE (3,60)
60  FORMAT (2X,'ENTER "B" (MAXIMUM) VALUE')
    READ (0,*) BB
    IF (BB-CC.LT.0) THEN
        GOTO 5
    END IF
    WRITE (3,70) NUM
70  FORMAT (2X,'ENTER NUMBER OF CELLS FOR VARIABLE',I3)
    READ (0,*) NCELLS(NUM)
    AREA=1/(FLOAT(NCELLS(NUM))*2.0)
    TAREA=(CC-AA)*(2/(BB-AA))/2
    DO 80 K=1,NCELLS(NUM)*2.0
        KAREA=AREA*(FLOAT(K))
        IF (KAREA.GT.TAREA) THEN
            A(K)=BB-SQRT((1-KAREA)*(BB-AA)*(BB-CC))
        ELSE
            A(K)=SQRT(KAREA*(BB-AA)*(CC-AA))+AA
        END IF
80  CONTINUE
    MEAN(NUM)=(AA+BB+CC)/3
    VARIANCE(NUM)=(AA**2+BB**2+CC**2-AA*BB-AA*CC-BB*CC)/18
    CALL CELLR
    RETURN
END

```

```

      SUBROUTINE UNIFORM
C *****
C *
C * PROGRAM NAME: SUBROUTINE UNIFORM
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 4, 1987
C *
C * PURPOSE: SUBROUTINE UNIFORM FINDS THE NCELL LOCATION ON THE X AXIS
C *           GIVEN A PROBABILITY (AREA) FOR A UNIFORM RANDOM VARIABLE.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: SUBROUTINE CELLR
C *
C * USAGE: CALL UNIFORM
C *
C * DISCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *       MAXX - REAL - MAXX IS THE MAXIMUM POSSIBLE VALUE FOR
C *               THE UNIFORM RANDOM VARIABLE.
C *       MINX - REAL - MINX IS THE MINIMUM POSSIBLE VALUE FOR
C *               THE UNIFORM RANDOM VARIABLE.
C *       NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *               LOCATIONS THAT APPROXIMATE THE PROBABILITY
C *               DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *       NUM - INTEGER - VARIABLE 'NUM' IS THE NUMBER OF THE
C *               RANDOM VARIABLE.
C *   OUTPUT ARGUMENTS:
C *       A(N) - REAL - WHERE N=NCELLS*2, ODD VALUES OF N
C *               REPRESENT THE CENTER OF GRAVITY LOCATIONS AND
C *               THE EVEN VALUES OF N INDICATE THE CELL
C *               BOUNDARIES.
C *       B(NCELLS,NUM) - REAL - THE DISCRETE POINT (CENTER OF
C *               GRAVITY) REPRESENTATIONS FOR NUM'S PROBABILITY
C *               DENSITY FUNCTION.
C *       MEAN(NUM) - REAL - MEAN(NUM) IS THE MEAN VALUE OF
C *               RANDOM VARIABLE NUM.
C *       VARIANCE(NUM) - REAL - VARIANCE(NUM) IS THE VARIANCE OF
C *               RANDOM VARIABLE NUM.
C *****
C
      REAL AREA,KAREA,MINX,MAXX,A,B,MEAN,VARIANCE,Z,ZMEAN,ZVARIANCE
      REAL UB,LB
      INTEGER NCELLS,NUM,K,COUNT,DISTR,UBCOUNT,LBCOUNT
      CHARACTER DISREP*10,ANSB*1
      COMMON/FIRST/A(0:400),B(0:400,2),Z(0:40000,3),COUNT
      COMMON/SECOND/DISTR(2),NCELLS(2),VARIANCE(2),MEAN(2),DISREP(2)

```

```

COMMON/THIRD/NUM,ZMEAN,ZVARIANCE,UB(0:400,2),LB(0:400,2),ANSB
COMMON/CBOUND/ZLB(0:40000,3),ZUB(0:40000,3),UBCOUNT,LBCOUNT
WRITE (3,*)
WRITE (3,*)
5  WRITE (3,10) NUM
10  FORMAT (2X,'UNIFORM DISTRIBUTION FOR VARIABLE',I3)
    WRITE (3,30)
30  FORMAT (2X,'RANGE (MINX,MAXX)')
    WRITE (3,40)
40  FORMAT (2X,'ENTER MINX VALUE')
    READ (0,*) MINX
    WRITE (3,50)
50  FORMAT (2X,'ENTER MAXX VALUE')
    READ (0,*) MAXX
    IF (MAXX-MINX.LT.0) THEN
        GOTO 5
    END IF
    WRITE (3,60) NUM
60  FORMAT (2X,'ENTER NUMBER OF CELLS FOR VARIABLE',I3)
    READ (0,*) NCELLS(NUM)
    AREA=1/(FLOAT(NCELLS(NUM))*2.0)
    DO 20 K=1,NCELLS(NUM)*2.0
        KAREA=AREA*(FLOAT(K))
        A(K)=MINX+KAREA*(MAXX-MINX)
20  CONTINUE
    MEAN(NUM)=(MINX+MAXX)/2.0
    VARIANCE(NUM)=((MAXX-MINX)**2)/12.0
    CALL CELLR
    RETURN
END

```



APPENDIX B

DENSITY FUNCTION PROGRAM

```

C *****
C *
C * PROGRAM NAME: DENSITY
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 26, 1987
C *
C * PURPOSE: PROGRAM "DENSITY" APPROXIMATES THE JOINT PROBABILITY
C *           DENSITY FUNCTION USING A CUBIC SPLINE AND A HISTOGRAM.
C *           THE DISTRIBUTION DATA IS READ IN FROM A FILE DETERMINED
C *           BY THE USER. IF THE PROGRAM FAILS DURING THE
C *           CUBIC SPLINE APPROXIMATION, THE "MANAGE" PROGRAM MUST BE
C *           USED TO ADJUST THE DATA.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: SUBROUTINES HISTOGRAM
C *                                         OUTPUTD
C *                                         SMOOTH
C *
C * DESCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *     COUNT - INTEGER - COUNT IS THE NUMBER OF UNIQUE Z
C *               LOCATIONS OF THE JOINT PROBABILITY DISTRIBUTION
C *               FUNCTION.
C *     NAME - CHARACTER*30 - NAME IS THE INPUT/OUTPUT DATA
C *               FILE NAME OF THE DISTRIBUTION/DENSITY FUNCTION.
C *     NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *               LOCATIONS THAT APPROXIMATE THE PROBABILIITY
C *               DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *     Z(COUNT,1) - REAL - THE Z VALUE OF THE JOINT
C *               PROBABILITY DISTRIBUTION FUNCTION.
C *     Z(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS
C *               THAN OR EQUAL TO Z(COUNT,1).
C *     Z(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF
C *               TIMES Z(COUNT,1) WAS REPEATED DURING THE
C *               SIMULATION.
C *
C *   OUTPUT ARGUMENTS:
C *     H(N,1) - REAL - THE Z LOCATION OF THE NTH CELL OF THE
C *               HISTOGRAM.
C *     H(N,2) - REAL - THE FRACTION OF Z LOCATIONS WITHIN THE
C *               RANGE OF H(N,1).
C *     NUMBER - REAL - THE NUMBER OF TOTAL DIVISIONS WITHIN
C *               THE HISTOGRAM.
C *     RANGE - REAL - THE RANGE OF VALUES OF THE HISTOGRAM.
C *     R(N) - REAL - THE "SMOOTHED" VALUES OF THE Z
C *               PROBABILITY DISTRIBUTION FOR VARIABLE N AFTER
C *               THE CUBIC SPLINE APPROXIMATION.

```

```

C *          R1(N) - REAL - THE DENSITY FUNCTION AFTER THE CUBIC      *
C *          SPINE APPROXIMATION.                                     *
C *          SIZE - REAL - THE SIZE OF EACH CELL WITHIN THE         *
C *          HISTOGRAM.                                              *
C *                                                                  *
C *****
C
C
C          INTEGER COUNT,NCELLS
C          REAL Z
C          CHARACTER NAME*30
C          COMMON/MAIN/Z(20000,4),COUNT,NCELLS(2),H(200,2),R(20004),
*          R1(20004),NUMBER
C          WRITE (3,10)
10          FORMAT (2X,'ENTER DISTRIBUTION DATA FILE NAME')
C          READ (0,20) NAME
20          FORMAT (A30)
C          OPEN (UNIT=23,FILE=NAME)
C          READ (23,60) NCELLS(1),NCELLS(2)
60          FORMAT (I5,I5)
C          READ (23,30) COUNT
30          FORMAT (I5)
C          DO 40 I=1,COUNT
C             READ (23,50) Z(I,1),Z(I,2),Z(I,3)
50          FORMAT (F20.7,F12.7,F12.7)
40          CONTINUE
C          CLOSE 23
C          CALL HISTOGRAM
C          CALL SMOOTH
C          CALL OUTPUTD
C          STOP
C          END

```

# SUBROUTINE HISTOGRAM

```

C *****
C *
C * PROGRAM NAME: SUBROUTINE HISTOGRAM
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 26, 1987
C *
C * PURPOSE: SUBROUTINE "HISTOGRAM" BUILDS A HISTOGRAM FROM A
C *           DISTRIBUTION FUNCTION STORED IN A DATA FILE.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: NONE
C *
C * USAGE: CALL HISTOGRAM
C *
C * DESCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *     COUNT - INTEGER - COUNT IS THE NUMBER OF UNIQUE Z
C *              LOCATIONS OF THE JOINT PROBABILITY DISTRIBUTION
C *              FUNCTION.
C *     NAME - CHARACTER*30 - NAME IS THE OUTPUT DATA FILE NAME
C *              OF THE HISTOGRAM.
C *     NCELLS(NUM) - INTEGER - THE NUMBER OF DISCRETE
C *              LOCATIONS THAT APPROXIMATE THE PROBABILITY
C *              DENSITY FUNCTION OF RANDOM VARIABLE NUM.
C *     Z(COUNT,1) - REAL - THE Z VALUE OF THE JOINT
C *              PROBABILITY DISTRIBUTION FUNCTION.
C *     Z(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS
C *              THAN OR EQUAL TO Z(COUNT,1).
C *     Z(COUNT,3) - REAL - THE NUMBER OF TIMES THE VALUE OF
C *              TIMES Z(COUNT,1) WAS REPEATED DURING THE
C *              SIMULATION.
C *
C *   OUTPUT ARGUMENTS:
C *     H(N,1) - REAL - THE Z LOCATION OF THE NTH CELL OF THE
C *              HISTOGRAM.
C *     H(N,2) - REAL - THE FRACTION OF Z LOCATIONS WITHIN THE
C *              RANGE OF H(N,1).
C *     NUMBER - REAL - THE NUMBER OF TOTAL DIVISIONS WITHIN
C *              THE HISTOGRAM.
C *     RANGE - REAL - THE RANGE OF VALUES OF THE HISTOGRAM.
C *     SIZE - REAL - THE SIZE OF EACH CELL WITHIN THE
C *              HISTOGRAM.
C *****
C
REAL RANGE,SIZE,H,Z

```

```

      INTEGER NUMBER,COUNT,NCELLS
      CHARACTER NAME*30
      COMMON/MAIN/Z(20000,4),COUNT,NCELLS(2),H(200,2),R(20004),
*          R1(20004),NUMBER
C      COMMON/
      NUMBER=NINT(1+3.3*LOG10(FLOAT(NCELLS(1)*NCELLS(2))))
      RANGE=Z(COUNT,1)-Z(1,1)
      SIZE=RANGE/FLOAT(NUMBER)
      DO 10 I=1,NUMBER
          H(I,1)=SIZE*FLOAT(I)+Z(1,1)
          H(I,2)=0.0
10      CONTINUE
          J=1
          DO 20 I=1,COUNT
30          DIFFER=ABS(Z(I,1)-H(J,1))
              IF (Z(I,1).LE.H(J,1).OR.DIFFER.LE.1.0E-6) THEN
                  H(J,2)=H(J,2)+Z(I,3)
              ELSE
                  J=J+1
                  GOTO 30
              END IF
20      CONTINUE
          DO 40 I=1,NUMBER
              H(I,2)=H(I,2)/(NCELLS(1)*NCELLS(2))
40      CONTINUE
          WRITE (3,50)
50      FORMAT (2X,'ENTER HISTOGRAM DATA FILE NAME')
          READ (0,60) NAME
60      FORMAT (A30)
          OPEN (UNIT=23,FILE=NAME)
          WRITE (23,70) NUMBER
70      FORMAT (I5)
          DO 90 I=1,NUMBER
              WRITE (23,80) H(I,1),H(I,2)
80      FORMAT (F12.7,F12.7)
90      CONTINUE
          CLOSE 23
          RETURN
      END

```

```

      SUBROUTINE OUTPUTD
C *****
C *
C * PROGRAM NAME: SUBROUTINE OUTPUTD
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 26, 1987
C *
C * PURPOSE: SUBROUTINE OUTPUTD PRINTS THE JOINT PROBABILITY DENSITY
C *           FUNCTION TO A FILE DESIGNATED BY THE USER.
C *
C * SUBROUTINES AND SUBPROCRAMS REQUIRED: NONE
C *
C * USAGE: CALL OUTPUTD
C *
C * DESCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *       COUNT - INTEGER - COUNT IS THE NUMBER OF UNIQUE Z
C *               LOCATIONS OF THE JOINT PROBABILITY DISTRIBUTION
C *               FUNCTION.
C *       NAME - CHARACTER*30 - NAME IS THE OUTPUT DATA FILE NAME
C *               OF THE APPROXIMATED JOINT PROBABILITY DENSITY
C *               FUNCTION.
C *       Z(COUNT,1) - REAL - THE Z VALUE OF THE JOINT
C *               PROBABILITY DISTRIBUTION FUNCTION.
C *       R(N) - REAL - THE "SMOOTHED" VALUES OF THE Z
C *               PROBABILITY DISTRIBUTION FOR VARIABLE N AFTER
C *               THE CUBIC SPLINE APPROXIMATION.
C *       RI(N) - REAL - THE DENSITY FUNCTION AFTER THE CUBIC
C *               SPLINE APPROXIMATION.
C *****
C
      INTEGER NCELLS,COUNT,NUMBER
      REAL Z,RANGE,SIZE,H,R,RI
      CHARACTER NAME*30
      COMMON/MAIN/Z(20000,4),COUNT,NCELLS(2),H(200,2),R(20004),
*          RI(20004),NUMBER
      WRITE (3,60)
      FORMAT (2X,"ENTER SPLINE DATA FILE NAME")
      READ (0,70) NAME
      FORMAT (A30)
      OPEN (UNIT=23,FILE=NAME)
      WRITE (23,80) COUNT
      FORMAT (I5)
      DO 90 I=1,COUNT
          WRITE (23,100) Z(I,1),R(I),RI(I)
          FORMAT (F20.7,F12.7,F12.7)
100      CONTINUE
      CLOSE 23
      RETURN
      END

```

## SUBROUTINE SMOOTH

```

C *****
C *
C * PROGRAM NAME: SUBROUTINE SMOOTH
C *
C * PURPOSE: SUBROUTINE SMOOTH APPROXIMATES THE DENSITY FUNCTION FROM
C *           A GIVEN SET OF DISTRIBUTION DATA.
C *
C *****
C
C
C SANDIA MATHEMATICAL PROGRAM LIBRARY
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C * * OF ANY * *
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C * * APPARATUS, * *
C * ***** PRODUCT ***** *
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C * * USE WOULD NOT ** *
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C ** OWNED **
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```

CONVERSION FROM THE ALGOL BY RONDALL E JONES  
 REFERENCE -- NUMERISHE MATHEMATIK 10,177-183 (1967) C H REINSCH

# ABSTRACT

SMOO FITS A SMOOTH SPLINE THROUGH A GIVEN SET OF DATA POINTS  
 BY MINIMIZING THE INTEGRAL OF THE SECOND DERIVATIVE SQUARED,  
 SUBJECT TO THE CONSTRAINT THAT

$$\sum_{I=1}^N (R(I)-Y(I))/DY(I) **2 \leq S$$

(WHERE R(I) IS THE ORDINATE OF THE SMOOTH SPLINE AT X(I).)  
 SMOO RETURNS THE VALUES OF THE SPLINE FUNCTION, R,  
 ITS FIRST DERIVATIVE, R1, AND ITS SECOND DERIVATIVE, R2,  
 EVALUATED AT THE ABCISSAS OF THE GIVEN DATA POINTS.  
 THE RESULTING SPLINE, DEFINED BY THE ARRAYS X, R, AND R2,  
 MAY THEN BE INTERPOLATED (IF DESIRED) USING SPLINT.  
 FOR AN EXACT SPLINE FIT SEE SUBROUTINE SPLIFT.

## DESCRIPTION OF ARGUMENTS

INPUT ARGUMENTS --

N - NUMBER OF DATA VALUES (AT LEAST 3)

X - ABCISSA ARRAY (INCREASING ORDER)

Y - ORDINATE ARRAY

DY - ARRAY OF ERROR ESTIMATES. DY(I) SHOULD BE AN ESTIMATE  
 OF THE ERROR (ACTUALLY, THE STANDARD DEVIATION) IN Y(I).  
 THUS, THE UNITS OF DY ARE THE SAME AS THE UNITS OF Y.  
 LARGER VALUES OF DY(I) ALLOW A LOOSER, SMOOTHER FIT.  
 SMALLER VALUES OF DY(I) CAUSE A TIGHTER FIT. SETTING  
 DY(I)=0 AT ALL POINTS RESULTS IN AN EXACT FIT.(SEE SPLIFT)  
 BY APPROPRIATELY ADJUSTING DY(I) AT EACH POINT, THE SPLINE  
 CAN BE MADE TIGHT AT CRITICAL POINTS AND LOOSE AT OTHERS.  
 S - SHOULD NORMALLY = N. (NOTE-- S IS FLOATING POINT - DONT  
 USE N DIRECTLY FOR S.) IF YOU WISH TO TIGHTEN OR LOOSEN  
 THE SPLINE FIT BY MULTIPLYING EACH ELEMENT OF DY BY  
 SOME FACTOR F, YOU MAY ALTERNATIVELY SIMPLY MULTIPLY  
 S BY F\*\*2.

OUTPUT ARGUMENTS --

A,B,C,D - CUBIC BETWEEN X(I) AND X(I+1) IS  
 $A(I) + B(I)*H + C(I)*H**2 + D(I)*H**3$   
 WHERE H IS DESIRED ABCISSA MINUS X(I).

R - ARRAY OF SMOOTH SPLINE VALUES

R1 - ARRAY OF SMOOTH SPLINE DERIVATIVES

R2 - ARRAY OF SMOOTH SPLINE SECOND DERIVATIVES

T,T1,U,V - WORK ARRAYS

IERR- A STATUS CODE

--NORMAL CODE

=1 MEANS THE REQUESTED SPLINE WAS COMPUTED.

--ABNORMAL CODE

=2 MEANS EITHER N IS LESS THAN 3, OR S IS NEGATIVE,  
 OR THE X-AXIS VALUES ARE MISORDERED.

X,Y,DY,A,B,C,D MUST BE DIMENSIONED AT LEAST N

R,R1,R2,T,T1,U,V MUST BE DIMENSION AT LEAST N+2



```

C      THE ORIGINAL N1 WAS FIXED AT 1 TO AVOID WASTED WORK ARRAY SPACE
C      THE ORIGINAL N2 IS CALLED N HERE
C      ALL WORK ARRAY INDICES ARE 1 LARGER THAN IN THE ALGOL,
C      TO AVOID A ZERO SUBSCRIPT.
C
      COMMON/MAIN/Z(20000,4),COUNT,NCCELLS(2),HHH(200,2),R(20004),
*      R1(20004),NUMBER
      COMMON/SMOO/T(20004),T1(20004),U(20004),V(20004),R2(20004),
*X(20000),Y(20000),DY(20000),A(20000),B(20000),C(20000),D(20000)
      INTECER COUNT
      REAL AAA,BBB,Z,R2,X,Y,DY,A,B,C,D
      WRITE (3,66)
66     FORMAT(2X,'ENTER TOLERANCE FOR SPLINE FIT')
      READ (0,*) F
      S=.05**2
C      S=FLOAT(COUNT)
      N=COUNT
      DO 15 I=1,COUNT
          X(I)=Z(I,1)
          Y(I)=Z(I,2)
          DY(I)=F
15     CONTINUE
C
C      CHECK INPUT DATA
C
      IERR = 1
      IF (N.LT.3) GO TO 10
      IF (S.LT.0.0) GO TO 10
      DO 5 I=2,N
          IF (X(I)-X(I-1)) 10,10,5
5      CONTINUE
      CO TO 20
10     WRITE (3,76)
76     FORMAT (2X,'PROGRAM BOMBED BECAUSE OF MISSORDER')
      IERR = 2
      RETURN
20     CONTINUE
C
C      INITIALIZE
C
      R(1) = 0.0
      R(2) = 0.0
      R1(N+1) = 0.0
      R2(N+1) = 0.0
      R2(N+2) = 0.0
      U(1) = 0.0
      U(2) = 0.0
      U(N+1) = 0.0
      U(N+2) = 0.0
      P = 0.0
      M2 = N-1
C
      H = X(2)-X(1)
      IF (H.EQ.0.0) THEN

```

```

      WRITE (3,439) X(2),X(1)
439  FORMAT (2X,'X(2)=',F12.7,'X(1)=',F12.7)
      END IF
      F = (Y(2)-Y(1))/H
      DO 100 I=2,M2
      G = H
      H = X(I+1)-X(I)
      E = F
      F = (Y(I+1)-Y(I))/H
      A(I) = F-E
      T(I+1) = 2.0*(G+H)/3.0
      T1(I+1) = H/3.0
      R2(I+1) = DY(I-1)/G
      R(I+1) = DY(I+1)/H
100  R1(I+1) = -DY(I)/G - DY(I)/H
C
      DO 200 I=2,M2
      B(I) = R(I+1)*R(I+1) + R1(I+1)*R1(I+1) + R2(I+1)*R2(I+1)
      C(I) = R(I+1)*R1(I+2) + R1(I+1)*R2(I+2)
200  D(I) = R(I+1)*R2(I+3)
      F2 = -S
C
C      NEXT ITERATION
C
250  DO 300 I=2,M2
      R1(I) = F*R(I)
      R2(I-1) = G*R(I-1)
      R(I+1) = 1.0/(P*B(I)+T(I+1)-F*R1(I)-G*R2(I-1))
      U(I+1) = A(I)-R1(I)*U(I)-R2(I-1)*U(I-1)
      F = P*C(I)+T1(I+1)-H*R1(I)
      G = H
300  H = D(I)*P
C
      DO 400 J=2,M2
      I=M2-(J-2)
400  U(I+1) = R(I+1)*U(I+1)-R1(I+1)*U(I+2)-R2(I+1)*U(I+3)
      E=0.0
      H=0.0
C
      DO 500 I=1,M2
      G = H
      H = (U(I+2)-U(I+1))/(X(I+1)-X(I))
      V(I+1) = (H-G)*DY(I)*DY(I)
500  E = E+V(I+1)*(H-G)
      V(N+1) = -H*DY(N)*DY(N)
      G = V(N+1)
      E = E-G*H
      G = F2
      F2 = E*P*P
      IF ((F2.GE.S).OR.(F2.LE.G)) GO TO 650
      F = 0.0
      H = (V(3)-V(2))/(X(2)-X(1))
C
      DO 600 I=2,M2

```

```

      G = H
      H = (V(I+2)-V(I+1))/(X(I+1)-X(I))
      G=H-G-R1(I)*R(I)-R2(I-1)*R(I-1)
      F = F+G*R(I+1)*G
600  R(I+1) = G
      H = E-P*F
      IF (H.LE.0.0) GO TO 650
      P = P+(S-F2)/((SQRT(S/E)+P)*H)
      GO TO 250

C
C   FINISH
C
650  DO 700 I=1,N
      A(I)=Y(I)-P*V(I+1)
700  C(I)=U(I+1)
C
      DO 800 I=1,M2
      H=X(I+1)-X(I)
      D(I)=(C(I+1)-C(I))/(3.0*H)
800  B(I)=(A(I+1)-A(I))/H-(H*D(I)+C(I))*H
      B(N)=0.0
      D(N)=0.0

C
C   OUTPUT DERIVATIVES, ETC.
C
      DO 900 I=1,M2
      R(I) = A(I)
      R1(I) = B(I)
900  R2(I) = 2.0*C(I)
      R(N) = A(N)
      H = X(N)-X(M2)
      R1(N)=(3.0*D(M2)*H+2.0*C(M2))*H+B(M2)
      R2(N) = 0.0
      DO 16 I=1,COUNT
          Z(I,3)=R(I)
          Z(I,4)=R1(I)
16   CONTINUE
      RETURN
      END

```

## APPENDIX C

### DATA MANAGEMENT PROGRAM

```

C *****
C *
C * PROGRAM NAME: MANAGE
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 6, 1987
C *
C * PURPOSE: PROGRAM "MANAGE" IS USED IF A PROBLEM OCCURS WITH THE
C * CALCULATION OF THE CUBIC SPLINE DENSITY APPROXIMATION
C * PROGRAM. THIS PROGRAM READS IN A DISTRIBUTION DATA FILE
C * AND FINDS THE SMALLEST ABSCISSA AND PRINTS OUT THE VALUE
C * TO THE SCREEN. IT THEN ELIMINATES ANY DATA POINTS THAT
C * ARE CLOSER THAN 0.000001 ALONG THE X AXIS.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: NONE
C *
C * DESCRIPTION OF PARAMETERS:
C * INPUT ARGUMENTS:
C * ANSWER - CHARACTER*1 - (Y OR N) AN INPUT VARIABLE USED
C * TO INDICATED IF THE USER WANTS TO PERFORM
C * DATA MANAGEMENT TO THE DISTRIBUTION FILE.
C * COUNT - INTEGER - NUMBER OF UNIQUE Z LOCATIONS.
C * NAME - CHARACTER*30 - NAME IS THE INPUT DATA FILE NAME
C * OF THE JOINT DISTRIBUTION FUNCTION.
C * NAME2 - CHARACTER*30 - NAME2 IS THE OUTPUT DATA FILE
C * NAME OF THE JOINT DISTRIBUTION FUNCTION.
C * NCELLS(1 OR 2) - INTEGER - THE NUMBER OF DISCRETE
C * LOCATIONS THAT APPROXIMATE THE PROBABILITY
C * DENSITY FUNCTION OF RANDOM VARIABLE 1 OR 2.
C * SMALL - REAL - THE SMALLEST X INCREMENT VALUE OF THE
C * JOINT DISTRIBUTION FUNCTION.
C * Z(COUNT,1) - REAL - THE ABSISCA VALUE OF THE JOINT
C * PROBABILITY DISTRIBUTION FUNCTION.
C * Z(COUNT,2) - REAL - THE PROBABILITY OF Z BEING LESS
C * THAN OR EQUAL TO Z(COUNT,1).
C *
C * OUTPUT ARGUMENTS:
C * INC - REAL - THE NUMBER OF UNIQUE Z LOCATIONS AFTER
C * THE DATA MANAGEMENT HAS BEEN PERFORMED.
C * ZD(INC,1) - REAL - THE ABSISCA VALUE OF THE JOINT
C * PROBABILITY DISTRIBUTION FUNCTION AFTER DATA
C * MANAGEMENT HAS BEEN PERFORMED.
C * ZD(INC,2) - REAL - THE PROBABILITY OF Z BEING LESS THAN
C * OR EQUAL TO Z(INC,1).
C *
C *****
C

```

```

REAL Z,DIFF1,DIFF2,SMALL,ZD,TOLERANCE
INTEGER COUNT,NCELLS1,NCELLS2
CHARACTER NAME*20,ANSWER*1,NAME2*20
COMMON/MAIN/Z(2,10000),ZD(2,10000)
TOLERANCE=.00001
SMALL=99999
WRITE (3,10)
10  FORMAT (2X,'ENTER INPUT FILE NAME')
    READ (0,20) NAME
20  FORMAT (A20)
    OPEN (UNIT=23,FILE=NAME)
    READ (23,30) NCELLS1,NCELLS2
30  FORMAT (I5,I5)
    READ (23,40) COUNT
40  FORMAT (I5)
    DO 50 I=1,COUNT
        READ (23,60) Z(1,I),Z(2,I),DUMMY
60  FORMAT (F20.7,F12.7,F12.7)
50  CONTINUE
    DO 70 I=2,COUNT
        DIFF1=ABS(Z(1,I)-Z(1,I-1))
        IF (DIFF1.LT.SMALL) THEN
            SMALL=DIFF1
        END IF
70  CONTINUE
    WRITE (3,80) SMALL
80  FORMAT (2X,'SMALLEST X INCREMENT=',F12.7)
    CLOSE 23
    WRITE (3,90)
90  FORMAT (2X,'PERFORM DATA MANAGEMENT? (Y/N)')
    READ (0,100) ANSWER
100 FORMAT (A1)
    IF (ANSWER.EQ.'N') THEN
        GOTO 9999
    END IF
    WRITE (3,200)
200 FORMAT (2X,'ENTER OUTPUT FILE NAME')
    READ (0,210) NAME2
210 FORMAT(A20)
    OPEN (UNIT=24,FILE=NAME2)
    INC=1
    ZD(1,1)=Z(1,1)
    ZD(2,1)=Z(2,1)
    DO 110 I=2,COUNT
        DIFF1=ABS(Z(1,I)-Z(1,I-1))
        IF (DIFF1.LT.TOLERANCE) THEN
            GOTO 110
        ELSE

```

```

        ZD(1,INC)=Z(1,I)
        ZD(2,INC)=Z(2,I)
        INC=INC+1
      END IF
110    CONTINUE
      INC=INC-1
      WRITE (24,220) NCELLS1,NCELLS2
220    FORMAT (I5,I5)
      WRITE (24,130) INC
130    FORMAT (I5)
      DO 120 I=1,INC
        Z(I,1)=ZD(I,1)
        Z(I,2)=ZD(I,2)
        WRITE (24,140) ZD(1,I),ZD(2,I)
140    FORMAT (F20.7,F12.7)
120    CONTINUE
      CLOSE 24
9999   STOP
      END

```

APPENDIX D

GRAPHICS PROGRAM



```

C *****
C *
C * PROGRAM NAME: GRAPH
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 27, 1987
C *
C * PURPOSE: PROGRAM "GRAPH" PLOTS THE JOINT DISTRIBUTION AND DENSITY
C *           FUNCTIONS ON A HIREZ SELANAR GRAPHICS VT100 TERMINAL.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: SUBROUTINES AUTOSCL
C *                                       GVDENSITY
C *                                       GVDISTRIBUTION
C *                                       LABEL
C *                                       PLLABEL
C *                                       PLOTT
C *                                       TICKS
C *
C * NOTE: SUBROUTINES AUTOSCL, LABEL, PLLABEL, PLOTT, AND TICKS WERE
C *        WRITTEN BY DONALD A. SMITH AND ARE USED BY PERMISSION.
C *
C * DESCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *       COUNT - INTEGER - THE NUMBER OF POINTS TO BE PLOTTED.
C *       GTITLE - CHARACTER*35 - THE TITLE OF THE GRAPH.
C *       XLABEL - CHARACTER*30 - THE LABEL ON THE X AXIS OF THE
C *                GRAPH.
C *       XX(N,M) - REAL - THE X VALUE FOR VARIABLE M, CURVE
C *                NUMBER N.
C *       YLABEL - CHARACTER*30 - THE LABEL ON THE Y AXIS OF THE
C *                GRAPH.
C *       YY(N,M) - REAL - THE Y VALUE FOR VARIABLE M, CURVE
C *                NUMBER N.
C *
C *****
C
      REAL MATL,MITL,XX,YY,ZZ,ERROR,MAXDEV
      INTEGER COUNT,INCELLS,ICOUNT,ITIME,ANS
      CHARACTER XLABEL*30,YLABEL*30,GTITLE*35,NAME*20
      LOGICAL AUTOSC
      COMMON/ONE/XX(6,10000),YY(6,10000),COUNT
      COMMON/MISC/INCELLS(10000),ERROR(10000),MAXDEV(10000),
      *           ICOUNT(10000),ITIME(10000)
1      WRITE (3,10)
10     FORMAT (2X,'      PLOTT: ')
      WRITE (3,20)

```

```

20     FORMAT (2X,'1. DISTRIBUTION FUNCTION')
      WRITE (3,30)
30     FORMAT (2X,'2. DENSITY FUNCTION')
      WRITE (3,40)
40     FORMAT (2X,'ENTER NUMBER FOR DESIRED CURVE')
      READ (0,*) ANS
      IF (ANS.EQ.1) THEN
        CALL GVTDISTRIBUTION
      ELSE
        IF (ANS.EQ.2) THEN
          CALL GVTDENSITY
        ELSE
          GOTO 1
        END IF
      END IF
      ICURVES=1
      WRITE (3,*)
      DO 6000 I=1,COUNT
        WRITE (3,*) XX(1,I),YY(1,I)
6000  CONTINUE
      WRITE (3,90)
90     FORMAT (2X,'ENTER TITLE')
      READ (0,100) GTITLE
100    FORMAT (A35)
      WRITE (3,92)
92     FORMAT (2X,'ENTER X LABEL')
      READ (0,93) XLABEL
93     FORMAT (A30)
      WRITE (3,94)
94     FORMAT (2X,'ENTER Y LABEL')
      READ (0,96) YLABEL
96     FORMAT (A30)
      MATL=0.03
      MITL=0.01
      IDEV=2
      AUTOSC=.TRUE.
      CALL PLOTT (D1,D2,D3,D4,MATL,MITL,XLABEL,YLABEL,
        *GTITLE,XX,YY,COUNT,ICURVES,IDEV,AUTOSC)
      END

```

```

      SUBROUTINE AUTOSCL (XX,YY,K,XMAX,XMIN,YMAX,YMIN,MAT SX,MITSX,
      @ MATSY,MITSY,AUTOSC,XSTART,YSTART,
      @ DX,DY)
C *****
C *
C * PROGRAM NAME: SUBROUTINE AUTOSCL
C *
C * WRITTEN BY: DONALD A. SMITH
C *
C *
C *
C *****
C
      REAL XX(6,10000),YY(6,10000),MAT SX,MATSY,MITSX,MITSY,DX(4),
      @ DY(4)
      LOGICAL AUTOSC
      IF (AUTOSC) THEN
        XMAX=XX(1,1)
        XMIN=XX(1,1)
        YMAX=YY(1,1)
        YMIN=YY(1,1)
        WRITE (3,2222)
2222 FORMAT (2X'ENTERED AUTOSCL')
        DO 10 I=1,K
          IF (XX(1,I).GT.XMAX) XMAX=XX(1,I)
          IF (XX(1,I).LT.XMIN) XMIN=XX(1,I)
          IF (YY(1,I).GT.YMAX) YMAX=YY(1,I)
          IF (YY(1,I).LT.YMIN) YMIN=YY(1,I)
10 CONTINUE
        ENDIF
        A=ABS((XMAX-XMIN)/5.)
        MAT SX=A
        C=ABS((YMAX-YMIN)/5.)
        MATSY=C
        IF ((XMAX.GT.0.) .AND. (ABS(XMAX-XMIN).LT.ABS(XMAX))) THEN
          XSTART=XMIN
          GOTO 14
        ENDIF
        IF ((XMAX.LE.0.) .AND. (ABS(XMAX-XMIN).LT.ABS(XMIN))) THEN
          XSTART=XMAX
        ELSE
          XSTART=0.
        ENDIF
14 CONTINUE
        IF ((YMAX.GT.0.) .AND. (ABS(YMAX-YMIN).LT.ABS(YMAX))) THEN
          YSTART=YMIN
          GOTO 30
        ENDIF

```

```

      IF ((YMAX.LE.0.) .AND. (ABS(XMAX-XMIN).LT.ABS(XMIN))) THEN
        YSTART=YMAX
      ELSE
        YSTART=0.
      ENDIF
30    CONTINUE
D    WRITE (3,*) 'TICKS CALC DX'
210  CONTINUE
      DX(1)=(XMAX-XSTART)/MAT SX
      DX(2)=(XMIN-XSTART)/MAT SX
      MITSX=MAT SX/5.
      DX(3)=(XMAX-XSTART)/MITSX
      DX(4)=(XMIN-XSTART)/MITSX
      DO 200 I=1,4
        DX(I)=INT(DX(I))
        DX(I)=REAL(DX(I))
        DX(I)=ABS(DX(I))
        IF ((I.LT.3) .AND. (DX(I).GT.20.)) THEN
          MAT SX=MAT SX*1.2
          GOTO 210
        ENDIF
200  CONTINUE
D    WRITE (3,*) 'TICKS CALC DY'
311  CONTINUE
      DY(1)=(YMAX-YSTART)/MAT SY
      DY(2)=(YMIN-YSTART)/MAT SY
      MITSY=MAT SY/5.
      DY(3)=(YMAX-YSTART)/MITSY
      DY(4)=(YMIN-YSTART)/MITSY
      DO 300 I=1,4
        DY(I)=INT(DY(I))
        DY(I)=REAL(DY(I))
        DY(I)=ABS(DY(I))
        IF ((I.LT.3) .AND. (DY(I).GT.20.)) THEN
          MAT SY=1.2*MAT SY
          GOTO 311
        ENDIF
300  CONTINUE
      RETURN
      END

```

```

SUBROUTINE GVTENSITY
C *****
C *
C * PROGRAM NAME: SUBROUTINE GVTENSITY
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 27, 1987
C *
C * PURPOSE: SUBROUTINE "GVTENSITY" READS IN THE DATA FORMATED FOR
C *          THE JOINT DENSITY FUNCTION.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: NONE
C *
C * DESCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *       COUNT - INTEGER - THE NUMBER OF POINTS TO BE PLOTTED.
C *       GRAPH.
C *       XX(N,M) - REAL - THE X VALUE FOR VARIABLE M, CURVE
C *                NUMBER N, FOR THE DENSITY FUNCTION.
C *       YY(N,M) - REAL - THE Y VALUE FOR VARIABLE M, CURVE
C *                NUMBER N, FOR THE DENSITY FUNCTIRON.
C *
C *****
C
      REAL MATL,MITL,XX,YY,ZZ,ERROR,MAXDEV
      INTEGER COUNT,INCELLS,ICOUNT,ITIME
      CHARACTER XLABEL*30,YLABEL*30,GTITLE*35,NAME*20
      COMMON/ONE/XX(6,10000),YY(6,10000),COUNT
      COMMON/MISC/INCELLS(10000),ERROR(10000),MAXDEV(10000),
*          ICOUNT(10000),ITIME(10000)
      WRITE (3,51)
51      FORMAT (2X,'PLOTING ROUTINE FOR DENSITY FUNCTION')
      WRITE (3,50)
50      FORMAT (2X,'ENTER DATA FILE NAME')
      READ (0,60) NAME
60      FORMAT (A20)
      OPEN (UNIT=23,FILE=NAME)
      READ (23,987) COUNT
987      FORMAT (I5)
      WRITE (3,*) COUNT
      DO 30 I=1,COUNT
          READ (23,628) XX(1,I),DUMMY,YY(1,I)
628          FORMAT (F20.7,F12.7,F12.7)
30      CONTINUE
      RETURN
      END

```

```

SUBROUTINE GVTDISTRIBUTION
C *****
C *
C * PROGRAM NAME: SUBROUTINE GVTDISTRIBUTION
C *
C * WRITTEN BY: BRUCE SWANSON
C *
C * DATE: MARCH 27, 1987
C *
C * PURPOSE: SUBROUTINE "GVTDISTRIBUTION" READS IN THE DATA FORMATED
C *           FOR THE JOINT DISTRIBUTION FUNCTION.
C *
C * SUBROUTINES AND SUBPROGRAMS REQUIRED: NONE
C *
C * DESCRIPTION OF PARAMETERS:
C *   INPUT ARGUMENTS:
C *       COUNT - INTEGER - THE NUMBER OF POINTS TO BE PLOTTED.
C *       GRAPH.
C *       XX(N,M) - REAL - THE X VALUE FOR VARIABLE M, CURVE
C *               NUMBER N, FOR THE DISTRIBUTION FUNCTION.
C *       YY(N,M) - REAL - THE Y VALUE FOR VARIABLE M, CURVE
C *               NUMBER N, FOR THE DISTRIBUTION FUNCTION.
C *
C *****
C
      REAL MATL,MITL,XX,YY,ZZ,ERROR,MAXDEV
      INTEGER COUNT,INCELLS,ICOUNT,ITIME
      CHARACTER XLABEL*30,YLABEL*30,CTITLE*35,NAME*20
      COMMON/ONE/XX(6,10000),YY(6,10000),COUNT
      COMMON/MISC/INCELLS(10000),ERROR(10000),MAXDEV(10000),
      *           ICOUNT(10000),ITIME(10000)
      WRITE (3,5000)
5000  FORMAT (2X,'PLOTING ROUTINE FOR THE DISTRIBUTION FUNCTION')
      WRITE (3,50)
50    FORMAT (2X,'ENTER DATA FILE NAME')
      READ (0,60) NAME
60    FORMAT (A20)
      OPEN (UNIT=23,FILE=NAME)
      READ (23,4001) IDUMMY1,IDUMMY2
4001  FORMAT (I5,I5)
      READ (23,987) COUNT
987   FORMAT (I5)
      WRITE (3,*) COUNT
      DO 30 I=1,COUNT
          READ (23,628) XX(1,I),YY(1,I),DUMMY
628   FORMAT (F20.7,F12.7,F12.7)
          WRITE (3,204) XX(1,I),YY(1,I)
204   FORMAT (F12.7,F12.7)

```

30

CONTINUE  
RETURN  
END

```

      SUBROUTINE LABEL (XMAX,XMIN,YMAX,YMIN,DX,DY,MATXS,MATSY,
&                      XSTART,YSTART)
C *****
C *
C * PROGRAM NAME: SUBROUTINE LABEL
C *
C * WRITTEN BY: DONALD A. SMITH
C *
C *
C *
C *****
C
      CHARACTER CNT*10
      REAL DX(20),DY(20),MATXS,MATSY
      XZ=ABS(XMAX-XMIN)
      YZ=ABS(YMAX-YMIN)
      CALL JCOLOR(0)
      CALL JMOVE (XMIN,YSTART)
      CALL JRMOVE(-XZ*.1,0.)
      CALL JJUST (3,2)
      CALL JFONT (1)
      A= (4./100.)*XZ
      B= (3./100.)*YZ
      CALL JSIZE (A,B)
      DO 50 I=1,2
        IF (DY(I).LT.0.9) GOTO 60
        DO 70 CTT=0.,DY(I)
          WRITE (CNT,'(F10.2)') CTT*MATSY+YSTART
          CALL JHSTRG (CNT)
          CALL JRMOVE (0.0,MATSY)
70      CONTINUE
60      CONTINUE
          CALL JMOVE (XMIN,YSTART)
          CALL JRMOVE (-XZ*.1,0.)
          MATSY=-MATSY
50      CONTINUE
          CALL JBASE (0.,1.,0.)
          CALL JMOVE (XSTART,YMIN)
          CALL JRMOVE (0.,-YZ*.1)
          CALL JSIZE (B*3./2.,A*2./3.)
D      WRITE (3,*) DX(1),'DX1',DX(2),'DX2'
D      WRITE (3,*) MATSX,'MATSX',YMIN,'YMIN',YMAX,'YMAX'
      DO 10 I=1,2
        IF (DX(I).LT.0.9) GOTO 20
        DO 30 CTT=0.,DX(I)
          WRITE (CNT,'(F10.2)') CTT*MATXS+XSTART
          CALL JHSTRG (CNT)
          CALL JRMOVE (MATXS,0.0)

```



```
30      CONTINUE
20      CONTINUE
        CALL JMOVE(XSTART,YMIN)
        CALL JREMOVE(0.,-YZ*.1)
        MATSX=-MATSX
10      CONTINUE
        CALL JBASE (1.,0.,0.)
        RETURN
        END
```

```

      SUBROUTINE PLLABEL (XMAX,XMIN,YMAX,YMIN,XLABEL,YLABEL,GTITLE)
C *****
C *
C * PROGRAM NAME: SUBROUTINE PLLABEL
C *
C * WRITTEN BY: DONALD A. SMITH
C *
C *
C *
C *****
C
      CHARACTER XLABEL*30,YLABEL*30,GTITLE*35
      CALL JPEDGE (2)
      CALL JCOLOR(0)
C      CALL JRECT (XMIN,YMIN,XMAX,YMAX)
      C=ABS(( 6./100.)*(XMAX-XMIN))
      D=ABS((5./100.)*(YMAX-YMIN))
C      CALL NOBLNK (GTITLE)
      CALL JSIZE (C,D)
      CALL JJUST (1,2)
C      CALL JMOVE (XMIN,YMAX)
      CALL JMOVE (-9.*(XMAX-XMIN)/19.,YMAX)
      CALL JRMOVE (0.,D*1.25)
      CALL JHSTRG (GTITLE)
      CALL JSIZE (0.66*C,0.66*D)
      CALL JMOVE (XMIN,YMIN)
      CALL JRMOVE (0.,-D*.4/.05)
      CALL JHSTRG (XLABEL)
      CALL JMOVE (XMIN,YMIN)
      CALL JRMOVE (-C*.4/.06,0.)
      CALL JBASE (0.,1.,0.)
      CALL JSIZE (0.66*D,0.66*C)
C      CALL JBASE (0.,1.,0.)
      CALL JHSTRG (YLABEL)
      CALL JBASE (1.,0.,0.)
      RETURN
      END

```

```

      SUBROUTINE PLOTT (XMAX,XMIN,YMAX,YMIN,MATL,MITL,
      @                XLABEL,YLABEL,GTITLE,XX,YY,IDATN,
      @                IDATNY,IDEV,AUTOSC)
C *****
C *
C * PROGRAM NAME: SUBROUTINE PLOTT
C *
C * WRITTEN BY: DONALD A. SMITH
C *
C *
C *
C *****
C
      REAL MATL,MITL,MATXS,MITSX,MATSY,MITSY,DX(4),DY(4),
      @      XX(6,10000),YY(6,10000),MALX,MALY,MILX,MILY
      LOGICAL AUTOSC
      CHARACTER XLABEL*30,YLABEL*30,GTITLE*35
D      IF (AUTOSC) WRITE (3,*) 'AUTOSC TRUE PLOTT SUB'
      CALL AUTOSCL(XX(1,1),YY(1,1),IDATN,XMAX,XMIN,YMAX,YMIN,MATXS,
      @      MITSX,MATSY,MITSY,AUTOSC,XSTART,YSTART,DX,DY)
D      IF (AUTOSC) THEN
D        WRITE (3,*) 'JUST DID AUTO SCALE'
D      ELSE
D        WRITE (3,*) 'NO AUTO SCALE'
D      ENDIF
      CALL JBEGIN
      CALL JDINIT (IDEV)
      CALL JDEVON (IDEV)
      CALL JVPORT (-0.3,0.7,-0.2,0.8)
C      IF (ABS(XMAX-XMIN).LT..1) XMIN=XMAX-.1
C      IF (ABS(YMAX-YMIN).LT..1) YMIN=YMAX-.1
      CALL JWINDO (XMIN,XMAX,YMIN,YMAX)
      CALL JDEVWN (IDEV,-1.0,1.0,-0.7,1.0)
      CALL JOPEN
C
D      WRITE (3,*) 'OPENED WINDOW'
      IW=10000
      MALY=MATL*ABS(XMAX-XMIN)
      MALX=MATL*ABS(YMAX-YMIN)
      MILY=MITL*ABS(XMAX-XMIN)
      MILX=MITL*ABS(YMAX-YMIN)
D      WRITE (3,*) '1ST TICKS CALL'
      CALL TICKS (XMAX,XMIN,YMAX,YMIN,MATXS,MATSY,MALX,MALY,DX(1),
      &      DY(1),IW,XSTART,YSTART)
D      WRITE (3,*) '1ST LABEL CALL'
      CALL LABEL (XMAX,XMIN,YMAX,YMIN,DX(1),DY(1),MATXS,MATSY,
      &      XSTART,YSTART)
      IW=8000

```

```

D    WRITE (3,*) '2ND TICKS CALL'
    CALL TICKS (XMAX,XMIN,YMAX,YMIN,MITSX,MITSY,MILX,MILY,
&      DX(3),DY(3),IW,XSTART,YSTART)
D    WRITE (3,*) 'PLLABEL CALL'
    CALL PLLABEL (XMAX,XMIN,YMAX,YMIN,XLABEL,YLABEL,GTITLE)
    CALL JJUST (2,2)
    RR=(6./100.)*ABS(XMAX-XMIN)
    TT=(6./100.)*ABS(YMAX-YMIN)
    CALL JSIZE (RR,TT)
    CALL JJUST (2,2)
    CALL JSIZE (RR*2./6.,TT*2./6.)
    DO 500 J=1,IDATNY
C    CALL JCOLOR(J)
    IF (J.GT.1) THEN
        CALL JLSTYL(3)
    END IF
    CALL JMOVE (XX(IDATNY,1),YY(IDATNY,1))
D    WRITE (3,*) XX(IDATNY,1),YY(IDATNY,1)
    DO 100 I=2,IDATN
    IF ((XX(J,I).LE.XMAX) .AND.
&      (XX(J,I).GE.XMIN) .AND.
&      (YY(J,I).LE.YMAX) .AND.
&      (YY(J,I).GE.YMIN)) THEN
        CALL JDRAW (XX(J,I),YY(J,I))
D    WRITE(3,*) XX(J,I),YY(J,I)
D    WRITE(3,*) J,'FIELD',I,'COUNT'
    ENDIF
100  CONTINUE
500  CONTINUE
C
    CALL JCLOSE
    CALL JPAUSE (IDEV)
    CALL JDEVOF (IDEV)
    CALL JDEND (IDEV)
    CALL JEND
    RETURN
END

```

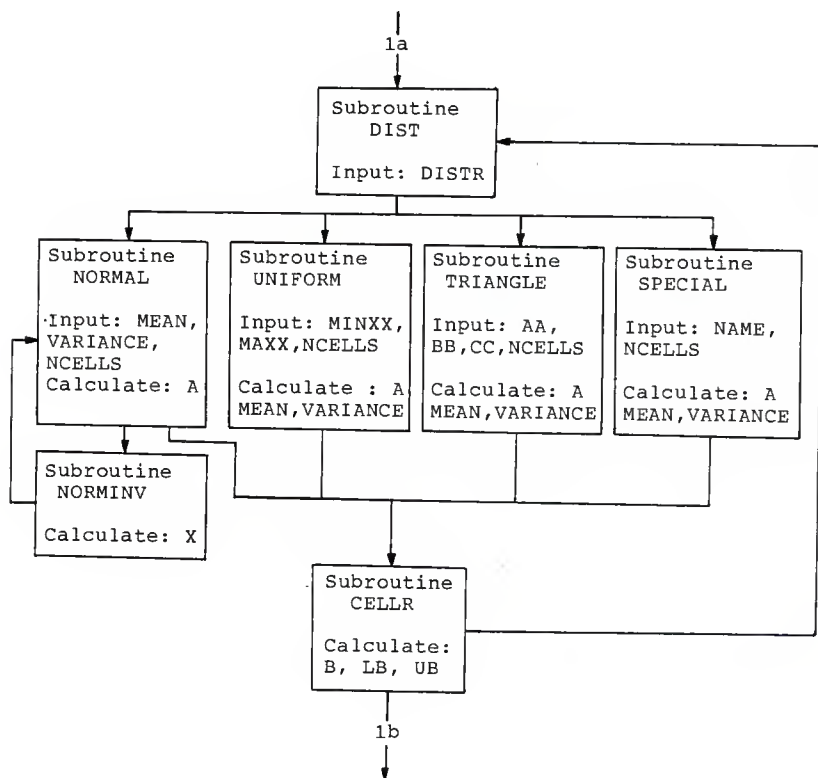
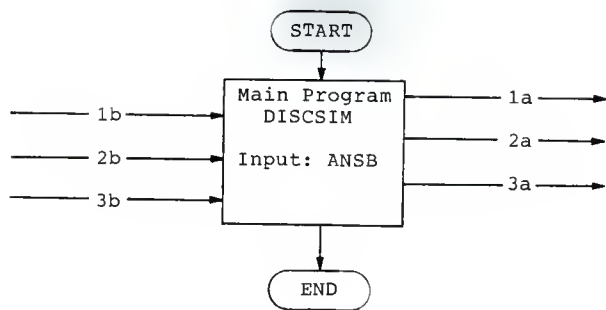
```

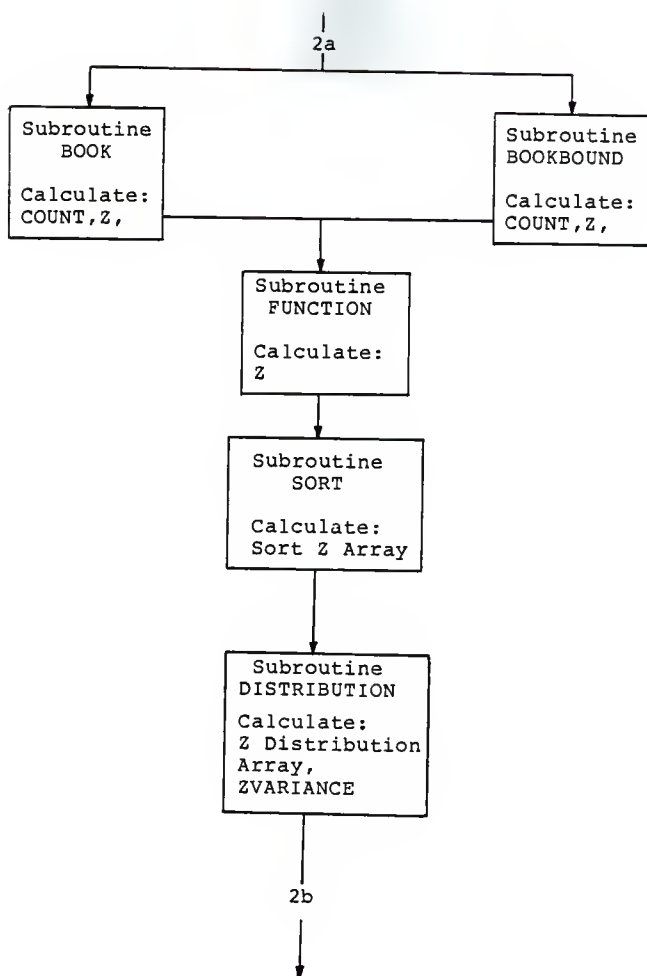
      SUBROUTINE TICKS (XMAX,XMIN,YMAX,YMIN,MAT SX,MAT SY,MLX,MLY,
      @                DX,DY,IW,XSTART,YSTART)
C *****
C *
C * PROGRAM NAME: SUBROUTINE TICKS
C *
C * WRITTEN BY: DONALD A. SMITH
C *
C *
C *
C *****
C
      REAL MAT SX,MAT SY,MLX,MLY,DX(20),DY(20)
      CALL JLSTYL (0)
      CALL JCOLOR (7)
      CALL JLWIDE (20000)
      CALL JMOVE (XMIN,YSTART)
      CALL JDRAW (XMAX,YSTART)
      CALL JMOVE (XSTART,YMAX)
      CALL JDRAW (XSTART,YMIN)
      CALL JMOVE (XSTART,YSTART)
D  WRITE (3,*) XSTART,XSTART ',YSTART,YSTART'
      CALL JLWIDE (IW)
      DO 20 I=1,2
      DO 10 SP=1.,DX(I)
          CALL JRMV (MAT SX,-MLX/2)
          CALL JRDRAW (0.,MLX)
          CALL JRMV (0.,-MLX/2)
10  CONTINUE
15  CONTINUE
      CALL JMOVE (XSTART,YSTART)
      MAT SX=-MAT SX
20  CONTINUE
      CALL JMOVE (XSTART,YSTART)
D  WRITE (3,*) 'TICKS CALC DY'
      DO 30 I=1,2
      DO 40 SP=1.,DY(I)
          CALL JRMV (-MLY/2.,MAT SY)
          CALL JRDRAW (MLY,0.)
          CALL JRMV (-MLY/2.,0.)
40  CONTINUE
45  CONTINUE
      CALL JMOVE (XSTART,YSTART)
      MAT SY=-MAT SY
30  CONTINUE
      RETURN
      END

```

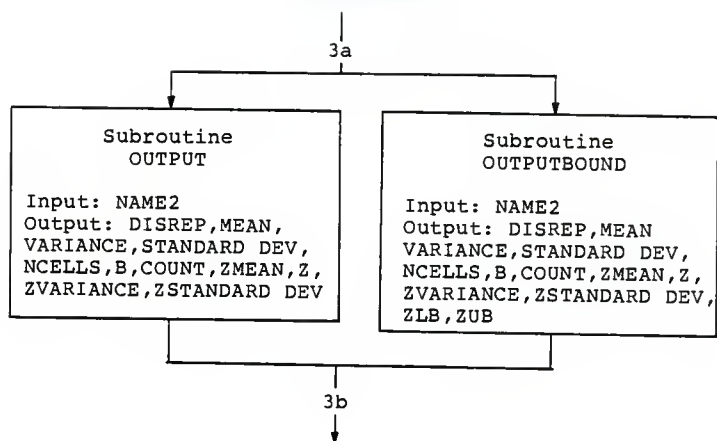
## APPENDIX E

### FLOWCHART FOR THE DISCSIM PROGRAM









## APPENDIX F

### SAMPLE OUTPUT OF THE DISCRETE SIMULATION PROGRAM

RANDOM VARIABLE: 1  
DISTRIBUTION: NORMAL  
MEAN: 170.000000  
VARIANCE: 525.070000  
STANDARD DEVIATION:  
NUMBER OF CELLS: 10

CELL NO.	X1 MINIMUM VALUE	CENTER OF GRAVITY	X1 MAXIMUM VALUE
1	-338.7965316	129.867140	137.9567811
2	137.9567811	144.082129	148.9635820
3	148.9635820	153.162715	156.8999533
4	156.8999533	160.3780729	163.6766633
5	163.6766633	168.8534752	176.3233317
6	176.3233317	173.1362248	179.6219271
7	179.6219271	183.1000467	183.8547285
8	183.8547285	191.0364180	195.9107871
9	195.9107871	202.0432199	211.1302860

25.000000

RANDOM VARIABLE: 2  
DISTRIBUTION: NORMAL  
MEAN: 29.400000  
VARIANCE: 9.000000  
STANDARD DEVIATION:  
NUMBER OF CELLS: 10

CELL NO.	X2 MINIMUM VALUE	CENTER OF GRAVITY	X2 MAXIMUM VALUE
1	-49.6556234	24.4643657	25.5548137
2	25.5548137	26.2907055	26.8756298
3	26.8756298	27.3774326	27.8279944
4	27.8279944	28.2453687	28.6412002
5	28.6412002	29.0238570	29.4000003
6	29.4000003	29.7761430	30.1587998
7	30.1587998	30.5446313	30.9720056
8	30.9720056	31.4225674	31.9243702
9	31.9243702	32.5092945	33.2651863
10	33.2651863	34.3356343	49.6556234

\*\*\*\*\* Z FUNCTION \*\*\*\*\*  
NUMBER OF UNIQUE Z LOCATIONS: 100  
MEAN: 94.25149  
VARIANCE: 771.93855  
STANDARD DEVIATION: 27.78198

CELL NO.	DISCRETE APPROXIMATION Z	LOWER BOUNDS DISTRIBUTION	Z	UPPER BOUNDS DISTRIBUTION
1	22.5014176	0.0100000	-49.8188562	-445.1651580
2	36.7463186	0.0200000	55.7483597	-436.2811629
3	37.7209165	0.0300000	60.3503526	-426.9347864
4	43.9331503	0.0400000	60.8256572	-421.0052830
5	46.7769751	0.0500000	64.3046270	-416.4032901
6	49.0728565	0.0600000	65.7531607	-412.4490157
7	51.9358175	0.0700000	57.941992	-408.7894435
8	53.2361557	0.0800000	58.7620284	-405.1622231
9	54.0097765	0.0900000	71.3371536	-401.2616751
10				0.0100000
11				0.0200000
12				0.0300000
13				0.0400000
14				0.0500000
15				0.0600000
16				0.0700000
17				0.0800000
18				0.0900000

10	57-0655046	0-1030000	71-5914195	0-1000000	-396-3968746	0-1000000
11	59-1526492	0-1100000	74-6915319	0-1100000	31-5884847	0-1100000
12	50-4972754	0-1200000	75-3114279	0-1200000	40-4724797	0-1200000
13	60-6733757	0-1300000	75-4919676	0-1300000	42-5952856	0-1300000
14	61-0413751	0-1400000	75-5387434	0-1400000	49-8188562	0-1400000
15	64-2923854	0-1500000	78-9710002	0-1500000	50-5316569	0-1500000
16	64-3939065	0-1600000	78-9735248	0-1600000	51-4792807	0-1600000
17	66-7643284	0-1700000	80-3367680	0-1700000	55-7483597	0-1700000
18	63-2937075	0-1800000	81-8620777	0-1800000	57-3083719	0-1800000
19	58-2746775	0-1900000	82-5982205	0-1900000	59-4156519	0-1900000
20	68-5154546	0-2000000	83-2477992	0-2000000	60-3503526	0-2000000
21	68-5124782	0-2100000	86-0702398	0-2100000	60-8256572	0-2100000
22	72-2861053	0-2200000	86-4987686	0-2200000	63-6317061	0-2200000
23	73-2536307	0-2300000	86-9073714	0-2300000	64-3046270	0-2300000
24	73-3484440	0-2400000	87-7915812	0-2400000	66-1923669	0-2400000
25	74-7620798	0-2500000	88-1854059	0-2500000	66-7551607	0-2500000
26	74-8701849	0-2600000	90-0245142	0-2600000	67-9641992	0-2600000
27	75-4415092	0-2700000	90-5345917	0-2700000	68-7620284	0-2700000
28	75-8983746	0-2800000	91-3635890	0-2800000	69-9550353	0-2800000
29	77-5717132	0-2900000	92-3935741	0-2900000	71-3571536	0-2900000
30	79-6136054	0-3000000	93-6940864	0-3000000	71-5914196	0-3000000
31	80-4864321	0-3100000	94-1149104	0-3100000	72-5157012	0-3100000
32	80-5812454	0-3200000	94-4351398	0-3200000	74-6915319	0-3200000
33	81-0311294	0-3300000	94-9621219	0-3300000	75-3114279	0-3300000
34	81-3421639	0-3400000	96-3678484	0-3400000	75-6919676	0-3400000
35	81-9289115	0-3500000	97-3113068	0-3500000	75-5387434	0-3500000
36	83-8519271	0-3600000	98-7169033	0-3600000	76-7317504	0-3600000
37	84-5045166	0-3700000	99-2999402	0-3700000	78-8390304	0-3700000
38	84-5934632	0-3800000	100-0074207	0-3800000	79-2935248	0-3800000
39	87-0586477	0-3900000	100-8916254	0-3900000	80-3567680	0-3900000
40	87-5185317	0-4000000	101-2119548	0-4000000	81-4682469	0-4000000
41	88-1979611	0-4100000	102-6711776	0-4100000	81-8620777	0-4100000
42	89-5769653	0-4200000	102-8984931	0-4200000	82-5982205	0-4200000
43	89-6694641	0-4300000	103-6345410	0-4300000	83-2477992	0-4300000
44	89-5424907	0-4400000	105-4934183	0-4400000	84-6681216	0-4400000
45	91-2919169	0-4500000	106-0766552	0-4500000	85-6157454	0-4500000
46	92-1872345	0-4600000	106-3307499	0-4600000	85-0702398	0-4600000
47	92-1872345	0-4700000	107-3351890	0-4700000	86-4987686	0-4700000
48	92-9040357	0-4800000	108-3279965	0-4800000	86-9073714	0-4800000
49	93-3376774	0-4900000		0-4900000		0-4900000

50	94-5353634	3-5300000	109-4474926	0-5000000	87-7915812	0-5000000
51	94-7511331	0-5100000	107-9573702	0-5100000	88-1854069	0-5100000
52	95-0523676	0-5200000	112-1375535	0-5300000	90-0245142	0-5200000
53	95-9022655	0-5300000	112-3998895	0-5400000	90-5545917	0-5300000
54	97-5509655	0-5400000	113-1074649	0-5500000	91-3635690	0-5400000
55	98-6745364	0-5500000	113-4299895	0-5600000	92-3935741	0-5500000
56	99-1457424	0-5600000	113-8595193	0-5700000	93-5521166	0-5600000
57	99-2253936	0-5700000	113-952941	0-5800000	93-6840864	0-5700000
58	100-1405371	0-5800000	116-7346852	0-5900000	94-1149104	0-5800000
59	101-3314172	0-5900000	117-3342639	0-6000000	94-4351398	0-5900000
60	101-9131648	0-6000000	118-7231887	0-6100000	94-9621219	0-6000000
61	102-3326577	0-6100000	118-8347876	0-6200000	95-6749225	0-6100000
62	103-2737317	0-6200000	120-6352333	0-6300000	96-3478484	0-6200000
63	104-0433688	0-6300000	121-0433361	0-6400000	97-3113048	0-6300000
64	104-7619396	0-6400000	123-1443344	0-6500000	98-7169033	0-6400000
65	104-9436365	0-6500000	124-4367905	0-6600000	98-2999402	0-6500000
66	104-9436365	0-6600000	124-6710565	0-6700000	100-0074207	0-6600000
67	106-6232394	0-6700000	125-5000337	0-6800000	100-8916254	0-6700000
68	107-0579010	0-6800000	125-5716048	0-6900000	101-2118548	0-6800000
69	107-8148195	0-6900000	125-5716045	0-7000000	102-6711776	0-6900000
70	108-6587174	0-7000000	129-3910648	0-7100000	102-8984931	0-7000000
71	110-9742234	0-7100000	131-0807257	0-7200000	103-6346410	0-7100000
72	111-2811702	0-7200000	132-0506371	0-7300000	104-5589176	0-7200000
73	111-4310334	0-7300000	133-6346049	0-7400000	104-9356183	0-7300000
74	112-3059461	0-7400000	135-6778574	0-7500000	106-0766552	0-7400000
75	112-9972390	0-7500000	137-8574407	0-7600000	106-3307499	0-7500000
76	115-0516209	0-7600000	139-5794055	0-7700000	107-5351890	0-7600000
77	115-1461197	0-7700000	144-1907749	0-7800000	108-8279966	0-7700000
78	116-1139597	0-7800000	144-4432059	0-7900000	109-4478926	0-7800000
79	118-638902	0-7900000	150-5041041	0-8000000	109-9579702	0-7900000
80	119-0269006	0-8000000	157-2908191	0-8100000	112-3999895	0-8000000
81	119-1349957	0-8100000	165-2171904	0-8200000	113-1074649	0-8100000
82	119-1349957	0-8200000	176-2239913	0-8300000	113-4299895	0-8200000
83	120-3372283	0-8300000	250-6599363	0-8400000	113-8585183	0-8300000
84	122-3789210	0-8400000	256-5394402	0-8500000	113-9052941	0-8400000
85	124-1076735	0-8500000	261-1904331	0-8600000	116-7346852	0-8500000
86	125-3232323	0-8600000	265-1467075	0-8700000	117-3842639	0-8600000
87	125-3232323	0-8700000	268-8042797	0-8800000	118-7233187	0-8700000
88	125-3232323	0-8800000	272-4315001	0-8900000	119-6347976	0-8800000
89	125-3232323	0-8900000	276-3320441	0-9000000	120-6352333	0-8900000
90	125-3232323	0-9000000	281-1908486	0-9100000		

90	131-3334345	0.9000000	312.9776349	121.0438361	0.9000000
91	131-4349727	0.9100000		124.4347905	0.9100000
92	132-5351773	0.9200000		124.6710565	0.9200000
93	135-5567277	0.9300000		125.5000337	0.9300000
94	135-5735513	0.9400000		125.5910649	0.9400000
95	139-1271784	0.9500000		128.5716045	0.9500000
96	141-9112580	0.9600000		132.0506371	0.9600000
97	142-9394477	0.9700000		133.4364049	0.9700000
98	166-5544736	0.9800000		135.6778574	0.9800000
99	150-6930502	0.9900000		139.5744055	0.9900000
100	157-1307569	1.0000000		144.4432059	1.0000000

## APPENDIX G

NOTES ABOUT THE COMPARISON OF THE MONTE CARLO SIMULATION  
VS. THE DISCRETE SIMULATION

## METHODS OF COMPARISON

The Monte Carlo program was written using some of the same subroutines that are used in the discrete simulation program. These subroutines are: DISTRIBUTION, FUNCTION, OUTPUT, and SORT. Subroutine BOOK is also used, however, a few lines of code had to be modified for the generation of the random numbers. The random numbers are generated using a normal random number generator provided by Harris Computer Systems.

Because the two methods are slightly different, a method of comparison had to be devised. Three methods were possible:

1. Use the same number of points to represent the random variables' density functions.
2. Use the same number of possible combinations of outcomes to represent the joint random variable's distribution function.
3. Use the same number of unique outcomes to represent the joint random variable's distribution function.



Method 1 was thought to give a distinct advantage to the discrete approximation program, while method 2 would give an advantage to the Monte Carlo simulation. Method 3 was chosen to be used as the method of comparison because it is somewhat of a compromise between methods 1 and 2. However, this method does give the discrete simulation a slight advantage over the Monte Carlo method for the example discussed in Chapter 5.

The discrete simulation program was run first each time the two methods were compared. After each discrete approximation, the number of unique outcomes  $m$ , was noted. The joint distribution function was then approximated using the Monte Carlo program by generating  $m$  pairs of random variables  $X$  and  $Y$ .

#### CPU TIME COMPARISON

Figures 5.18 and 5.19 were plotted to give an overall comparison of the computer processing unit time for both programs. A binary sorting routine was used for both programs. Since the outcomes generated by the discrete method are somewhat ordered, the discrete values are sorted quicker than the data that is created randomly from the Monte Carlo program. It should be noted that

these times are largely dependent on the efficiency of the sorting routine. Quicker times could be achieved by using a faster method of sorting the outcomes. For example, a quick sort subroutine would decrease the run time of the Monte Carlo program since this method sorts random numbers faster than ordered numbers. Also, a heap sort could be used to decrease the CPU time for the ordered discrete values.

## ACKNOWLEDGEMENTS

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I am also grateful to Dr. Hugh S. Walker and Dr. Mark S. McNulty for serving as graduate committee members.

Appreciation is also extended to the Department of Mechanical Engineering for their financial support throughout my graduate study.

My thanks are also due to Donald A. Smith for the graphics routines he wrote for the Harris computer system.

And finally, I would like to thank my parents for their moral and financial support during all of my years of study.

VITA

Bruce Eugene Swanson

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THE APPROXIMATION OF JOINT DISTRIBUTION FUNCTIONS FOR  
APPLICATION IN PROBABILISTIC MECHANICAL DESIGN

by

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## ABSTRACT

In this study, a discrete simulation technique is developed to approximate the joint probability distribution functions of complex algebraic expressions containing several independent random variables. The discrete simulation package is written in Fortran on a Harris H-800 Super Mini-Computer.

The discrete approximation is compared to the traditional Monte Carlo technique for a joint probability distribution function that can be solved exactly. The accuracy of the discrete simulation is found to be better than the Monte Carlo simulation.

A cantilever I-beam is analyzed to illustrate how the discrete simulation package can be used by design engineers in a variety of engineering applications.

A bounding program is also discussed that bounds the exact joint probability distribution function of an algebraic expression. Although a formal proof has not been developed, the results of several different examples show considerable promise for further study.